

SPT-3G D1: CMB temperature and polarization power spectra and cosmology from 2019 and 2020 observations of the SPT-3G Main field

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We present measurements of the temperature and E-mode polarization angular power spectra of the cosmic microwave background (CMB) from observations of 4% of the sky with SPT-3G, the current camera on the South Pole Telescope (SPT). The maps used in this analysis are the deepest used in a CMB TT/TE/EE analysis to date. The maps and resulting power spectra have been validated through blind and unblind tests. The measurements of the lensed EE and TE spectra are the most precise to date at $\ell = 1800\text{--}4000$ and $\ell = 2200\text{--}4000$, respectively. Combining our TT/TE/EE spectra with previously published SPT-3G CMB lensing results, we find parameters for the standard Λ CDM model consistent with *Planck* and ACT DR6 with comparable constraining power. We report a Hubble constant of $H_0 = 66.66 \pm 0.60 \text{ km s}^{-1} \text{ Mpc}^{-1}$ from SPT-3G alone, 6.2σ away from local measurements from SH0ES. For the first time, combined ground-based (SPT+ACT) CMB primary and lensing data have reached *Planck*'s constraining power on some parameters, a milestone for CMB cosmology. The combination of these three CMB experiments yields the tightest CMB constraints to date, with $H_0 = 67.24 \pm 0.35 \text{ km s}^{-1} \text{ Mpc}^{-1}$, and the amplitude of clustering $\sigma_8 = 0.8137 \pm 0.0038$. CMB data alone show no evidence for physics beyond Λ CDM; however, we observe a 2.8σ difference in Λ CDM between CMB and baryon acoustic oscillation (BAO) results from DESI-DR2, which is relaxed in extended models. The combination of CMB and BAO yields $2\text{--}3\sigma$ shifts from Λ CDM in the curvature of the universe, the amplitude of CMB lensing, or the dark energy equation of state. It also drives mild preferences for models that address the Hubble tension through modified recombination or variations in the electron mass in a non-flat universe. This work highlights the growing power of ground-based CMB experiments and lays a foundation for further cosmological analyses with SPT-3G.

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I. INTRODUCTION

One of the main driving forces behind the phenomenal progress of cosmology in the past thirty years has been observations of the cosmic microwave background (CMB). CMB measurements from the *Planck* satellite have confirmed the Λ Cold Dark Matter (Λ CDM) model as the standard model of cosmology, constraining Λ CDM parameters at the percent or sub-percent level [1]. At the same time, results from *Planck* show a number of inconsistencies when compared with other probes, which could hint at cracks developing in this heretofore very successful framework. The most notable of these is the Hubble tension, i.e., the discrepancy between the expansion rate of the universe inferred from early-universe probes, such as the CMB and baryon acoustic oscillations (BAO), and the value measured directly using distance-ladder measurements, such as the Cepheid-calibrated Type Ia supernova measurements from the SH0ES collaboration [2]. Moreover, recent BAO results from the second data release from the Dark Energy Spectroscopic Instrument (DESI DR2, [3]) suggest an evolving dark energy equation of state, a suggestion that is reinforced by a discrepancy in Λ CDM parameter values preferred by BAO, CMB, and uncalibrated SNe Ia observations. Furthermore, classical questions in cosmology, such as the nature of dark matter and dark energy, remain unanswered.

While the *Planck* constraints have dominated the CMB field since the first *Planck* data release in 2013, improved measurements of CMB anisotropies, particularly at small scales and in polarization, promise to bring additional and complementary information [4]. CMB anisotropies in polarization are less contaminated by extragalactic foregrounds than in temperature, allowing the extraction of cosmological information from smaller angular scales. Furthermore, polarization measurements are less contaminated by atmospheric fluctuations, making them easier to achieve with ground-based experiments. Recently, the Atacama Cosmology Telescope (ACT) collaboration published its newest data release (DR6), showing constraints on cosmology which are consistent with *Planck* with an almost equivalent constraining power on many cosmological parameters [5–7].

In this paper, we present results from 2019 and 2020 observations with SPT-3G, the current camera mounted on the 10-meter South Pole Telescope (SPT). The SPT [8] is located at the Amundsen-Scott South Pole Station, one of the premier sites on Earth for CMB research, and was designed specifically for low-noise, high-resolution observations of the CMB. SPT observations began in 2007 with the SPT-SZ camera, which was sensitive to total intensity in three bands [9], and SPT capabilities were expanded to polarization with the SPTpol camera in 2012 [10]. SPT-3G is the third-generation SPT camera, installed in 2016-2017. It features $\sim 16\,000$ polarization-sensitive detectors

(over 10 times more detectors than SPTpol or SPT-SZ) configured to observe at 95, 150, and 220 GHz [11].

The first cosmological results from SPT-3G [12–15] were derived from observations taken in 2018 of the 1500 deg² SPT-3G Main field over four months (half a normal SPT “winter” season) and using half of the focal plane. We refer to these results as SPT-3G 2018. Results were found to be consistent with the Λ CDM model and in agreement with other CMB experiments, with error bars on cosmological parameters such as the Hubble constant a factor of three larger than those from *Planck*.

Starting from austral summer 2018-2019, the whole focal plane was operational, and SPT-3G has been used to observe the Main field every winter season since (except 2024, during which the SPT-3G Wide Survey was conducted) and with consistently high observing efficiency (typically 60% of all time spent scanning the CMB field). We refer to the observations of the SPT-3G Main field taken during the austral winters in 2019 and 2020 as SPT-3G D1 observations, and first cosmological results based on them were published in [16]. There, a Bayesian map-based approach was used to infer the unlensed EE power spectra and the CMB lensing power spectrum using polarization alone. We refer to these results as MUSE. Despite being based on polarization alone, this work provided competitive constraints on cosmological parameters—for example, the Hubble constant was constrained with precision within a factor of 1.5 of *Planck*. Recently, [17] used the same observations to present a measurement of the B-mode polarization power spectrum on large angular scales ($\ell \in [32, 502]$). Enabled by new techniques for mitigating polarized atmospheric emission [18], the analysis of the resulting data places a 95% confidence upper limit on the tensor-to-scalar ratio, $r < 0.25$, which is the second-best ground-based B-mode constraint after the one from BICEP/Keck [19]. Finally, [20] combined these same observations with SPTpol and *Herschel*-SPIRE [21] data on 100 deg² of the sky. The analysis set constraints on the duration of the epoch of reionization using the non-Gaussianity of the kinematic Sunyaev-Zel’dovich effect.

In this paper, we present TT/TE/EE (or T&E) power spectrum measurements and cosmological results from the same observations of the SPT-3G Main field. We call our measurement of CMB power spectra SPT-3G D1 T&E. The inclusion of TT and TE data significantly increases the constraining power of these observations. In particular, when combining with the SPT-3G measurements of the CMB lensing power spectrum from [16], results from SPT-3G alone have comparable (within $\sim 25\%$) constraining power to that of *Planck* on some cosmological parameters, such as H_0 and the amplitude of the matter density perturbations today, σ_8 .

The maps used in this analysis are the deepest high-resolution CMB data for a T&E analysis, with coadded white noise levels of $3.3 \mu\text{K}$ –arcmin in temperature

and $5.1 \mu\text{K}$ –arcmin in polarization. The production and validation of the maps are described in detail in an upcoming companion paper (W. Quan et al., in preparation, hereafter Q25).

Producing these CMB power spectrum measurements motivated a series of advances with respect to previous SPT-3G analyses to address the sensitivity of our data. These include:

- A curved-sky pipeline that uses HEALPix [22] and fast spherical-harmonic tools [23, 24], an improvement from the flat-sky approximation used in previous SPT-3G T&E analyses.
- A new code to produce realistic and fast end-to-end simulations called *Quickmock*, which will be described in a future publication (E. Hivon et al., in preparation).
- Modeling of systematic effects, including map-making filtering artifacts, quadrupolar temperature-to-polarization leakage, and polarized beams.
- A new algorithm to inpaint point source masks with Gaussian constrained realizations, called *cork*.
- An accurate semi-analytical covariance matrix, based on the work of Camphuis et al. [25].
- A differentiable parameter estimation pipeline, using a JAX-based likelihood, called *candl* [26, 27], and machine learning emulators, such as OLE [28] and *CosmoPower* [29, 30].
- A CMB-only foreground-marginalized likelihood, *SPT-lite* [31].

The analysis pipeline is kept blind to the final results until all cosmology-independent consistency tests pass a pre-determined threshold. We did discover effects that were not caught in the blind validation process, and we chose to implement non-trivial post-unblinding changes to address these.

Our analysis pipeline allows us to produce a robust and validated set of band powers from the SPT data. These are the most precise determinations of the lensed EE spectrum at $\ell = 1800$ –4000, and of the lensed TE spectrum at $\ell = 2200$ –4000, with comparable uncertainties to ACT DR6 at $\ell = 1800$ –2200 in TE. The ACT DR6 data set is more sensitive at large angular scales, owing to its broader sky coverage. In contrast, SPT-3G achieves better performance at small angular scales because of its lower noise levels. We use the SPT-3G D1 T&E binned power spectrum measurements, or band powers, described in this paper, together with the results of the measurement of CMB lensing from MUSE, to set constraints on cosmology. The results presented in this paper, as well as the combination of data products associated with them, is referred to as SPT-3G D1.

This paper represents a milestone for CMB cosmology in many respects, and we summarize our key findings in Section II. For the first time, a combination of ground-based experiments, SPT-3G D1 and ACT DR6, reaches *Planck*'s precision on some cosmological parameters, such as the Hubble constant. This is the beginning of a new era for CMB cosmology, in which our knowledge of the universe will increasingly be driven by experiments from the ground. Furthermore, while we find that the results from SPT-3G, *Planck*, and ACT DR6 are in excellent agreement and consistent with Λ CDM, we report a growing inconsistency with the BAO results from DESI DR2. The results from the combination of SPT-3G, *Planck*, and ACT DR6 are discrepant with those of DESI DR2 at the 2.8σ level in the Λ CDM model. This is alleviated in extended models of cosmology, and in some cases we find deviations from Λ CDM at the level of $2\text{--}3\sigma$ when CMB and BAO data are combined.

The rest of this work is structured as follows. In Section III we provide a summary of the processing of raw telescope data into CMB maps. In Section IV we present our power spectrum pipeline and the measured band powers. In Section V we present our model for the data and detail our cosmological inference procedure (the likelihood). In Section VI we present the validation of our data products through an extensive suite of consistency and robustness tests. We highlight that our model accurately describes the data to a high degree of precision. In Section VII we present the cosmological analysis before closing with concluding remarks in Section VIII. The data and likelihood code used in this paper are publicly available.¹

II. SUMMARY OF RESULTS

In this section, we summarize the key results of the paper. Specifically, we present the band powers estimated from SPT-3G measurements of the temperature and polarization anisotropies of the CMB and the cosmological parameter constraints that these band power measurements enable. The results are based on observations of the SPT-3G Main field, a region that covers roughly 4% of the sky, taken during the austral winter seasons of 2019 and 2020. The maps used in this analysis are the deepest CMB data used for a T&E analysis to date. The minimum-variance band powers obtained from these maps (by combining band powers from all combinations of SPT-3G frequency maps) are shown in Fig. 1. These data provide the tightest measurements of the lensed EE and TE power spectra at angular multipoles $\ell = 1800\text{--}4000$ and $\ell = 2200\text{--}4000$, respectively, while being comparable to ACT DR6 at $\ell = 1800\text{--}2200$ in TE. We use these band powers, in

combination with the lensing power spectrum from [16], to set constraints on cosmology. Here we highlight our key findings.

1. The SPT-3G data align well with the Λ CDM model when considered independently of other datasets, providing strong confirmation on angular scales not accessible to *Planck* [33]. The TT, TE, and EE channels independently are well fit by Λ CDM and yield consistent results. Our strongest cosmological constraints come from the TE channel. The Λ CDM parameter results are reported in Table I. They are consistent with the results of *Planck*² and ACT DR6³, as shown in Fig. 2.
2. The constraining power of SPT-3G is comparable to that of *Planck* on some Λ CDM cosmological parameters. For the Hubble constant, we find $H_0 = 66.66 \pm 0.60 \text{ km s}^{-1} \text{ Mpc}^{-1}$, in excellent agreement with *Planck* and ACT DR6 and 6.2σ away from the local measurements of SH0ES [2]. The SPT-3G data alone is thus able to confirm the Hubble tension with a high level of statistical significance.
3. For the first time, a combination of CMB ground-based experiments, SPT-3G D1 and ACT DR6, which we refer to as SPT+ACT, reaches *Planck*'s constraining power on some Λ CDM and extended model parameters. This is a milestone for CMB cosmology. With SPT+ACT, we obtain $H_0 = 66.59 \pm 0.46 \text{ km s}^{-1} \text{ Mpc}^{-1}$. We report the strongest CMB constraints to date by combining SPT, ACT, and *Planck* together into CMB-SPA, resulting in a constraint of $H_0 = 67.24 \pm 0.35 \text{ km s}^{-1} \text{ Mpc}^{-1}$, a 6.4σ discrepancy from SH0ES. We also highlight that CMB-SPA constrains the scalar spectral index to $n_s = 0.9684 \pm 0.0030$, a 10.5σ difference from a scale-invariant primordial power spectrum with $n_s = 1$.
4. The amplitude of matter density perturbations today, σ_8 , and of the matter density, Ω_m , from SPT-3G are in excellent agreement with the findings of other CMB experiments. Many large scale structure probes now provide results consistent with those of the CMB on these parameters, including the latest cosmic shear analysis of the Kilo-Degree Survey (KiDS) [37], the 3×2 point analysis of the Dark Energy Survey (DES) [38], the CMB lensing analysis of [39], and the galaxy cluster statistics analysis of [40].

² We use the combination of *Planck* T&E from Public Release 3 (PR3) [34] with the *Planck* $\phi\phi$ CMB lensing [35] estimated from PR4 maps [32]. To constrain reionization, instead of using the PR3 large scale EE polarization, we use a prior on τ_{reio} from the analysis of PR4 maps [32].

³ We use the combination of ACT DR6 T&E band powers [6] and CMB lensing $\phi\phi$ [36] with a prior on τ_{reio} from the analysis of PR4 maps [32].

¹ <https://pole.uchicago.edu/public/data/camphuis25/>

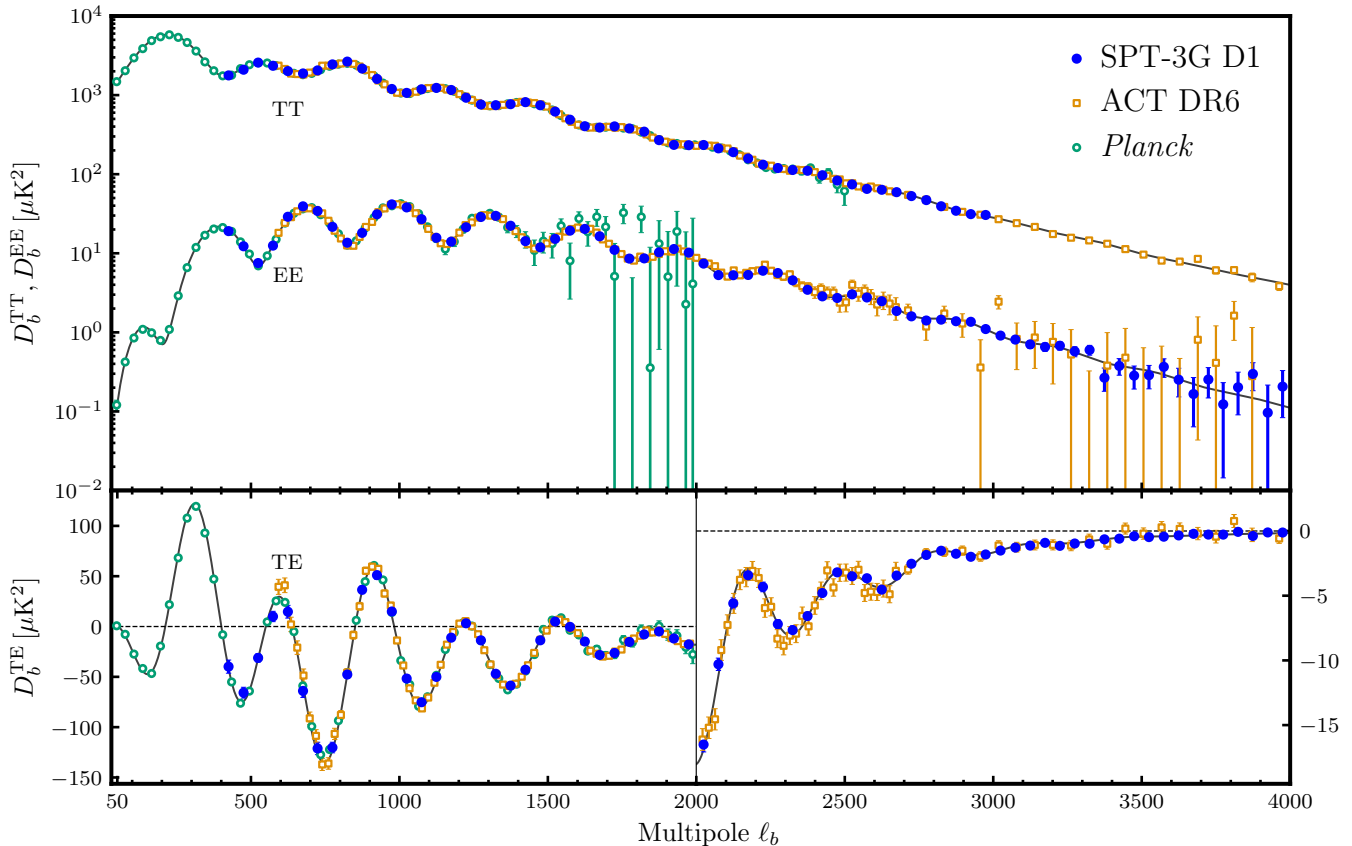


FIG. 1. TT, TE, and EE band powers from SPT-3G D1 (blue dots), ACTDR6 (orange empty squares) and *Planck* PR3 (green empty dots). Band powers from each experiment are foreground- and nuisance-parameter cleaned combinations of all auto- and cross-frequency spectra. We also show the best-fit Λ CDM model to SPT-3G D1 T&E (solid line). *Top*: TT and EE band powers on a logarithmic scale. SPT TT band powers are estimated in the multipole range $\ell = 400$ to 3000, while the range for TE and EE band powers is $\ell = 400$ to 4000, see §IV D for details. *Bottom*: TE band powers in linear scale, with a zoomed-in view of the $\ell > 2000$ region where ground-based experiments dominate the measurement. These data sets demonstrate excellent agreement with each other, and the SPT-3G D1 T&E data provide the tightest measurement of the lensed EE and TE band powers at $\ell = 1800$ -4000 and $\ell = 2200$ -4000, respectively.

5. We investigate the amplitude of CMB lensing implied from its effect on the primary CMB power spectra and find a value consistent with the Λ CDM prediction, $A_{\text{lens}} = 1.016^{+0.048}_{-0.054}$ from SPT+ACT T&E data, a result that differs at $\sim 2\sigma$ from the mild anomaly in the *Planck* data [1].
6. We report a growing discrepancy between CMB data and BAO data from DESI DR2 in Λ CDM, at the level of 2.8σ in the Ω_m - hr_d plane⁴ when SPT, ACT DR6, and *Planck* are combined.
7. While the CMB data alone do not prefer any extended model over Λ CDM, the discrepancy

between the CMB experiments and DESI is alleviated in some extended models of cosmology. The combination of CMB and BAO yields 2 - 3σ deviations from the standard model of cosmology.

III. SKY MAPS

There are several steps between observations of the microwave sky and cosmological analysis of the CMB. In this section, we discuss the steps to produce CMB maps from the raw observations. The observations themselves and the processing of the data are described in great detail in Q25; in this work we highlight characteristics of the data that we need to take into account in our power spectrum modeling.

⁴ r_d is the sound horizon at the drag epoch and $h \equiv H_0/100 \text{ km s}^{-1} \text{ Mpc}^{-1}$.

Parameter	<i>Planck</i>	SPT-3G D1	ACT DR6	SPT+ACT	SPT+ <i>Planck</i>	CMB-SPA
<i>Sampled</i>						
$10^4 \theta_s^*$	104.184 ± 0.029	104.171 ± 0.060	104.157 ± 0.030	104.158 ± 0.025	104.176 ± 0.026	104.162 ± 0.023
$100 \Omega_b h^2$	2.238 ± 0.014	2.221 ± 0.020	2.257 ± 0.016	2.247 ± 0.013	2.230 ± 0.011	2.2381 ± 0.0093
$100 \Omega_c h^2$	11.98 ± 0.11	12.14 ± 0.16	12.26 ± 0.17	12.22 ± 0.12	12.050 ± 0.089	12.009 ± 0.086
n_s	0.9657 ± 0.0040	0.951 ± 0.011	0.9682 ± 0.0069	0.9671 ± 0.0058	0.9636 ± 0.0035	0.9684 ± 0.0030
$\log(10^{10} A_s)$	3.042 ± 0.011	3.054 ± 0.015	3.038 ± 0.012	3.042 ± 0.011	3.046 ± 0.010	3.0479 ± 0.0099
τ_{reio}	0.0535 ± 0.0056	0.0506 ± 0.0059	0.0513 ± 0.0060	0.0514 ± 0.0059	0.0538 ± 0.0054	0.0559 ± 0.0055
<i>Derived</i>						
H_0 [km/s/Mpc]	67.41 ± 0.49	66.66 ± 0.60	66.51 ± 0.64	66.59 ± 0.46	67.07 ± 0.38	67.24 ± 0.35
Age [Gyr]	13.797 ± 0.022	13.826 ± 0.027	13.797 ± 0.021	13.805 ± 0.016	13.812 ± 0.017	13.805 ± 0.014
$10^9 A_s e^{-2\tau_{\text{reio}}}$	1.883 ± 0.010	1.915 ± 0.021	1.884 ± 0.013	1.889 ± 0.011	1.8890 ± 0.0092	1.8843 ± 0.0060
Ω_Λ	0.6854 ± 0.0067	0.6753 ± 0.0091	0.670 ± 0.010	0.6722 ± 0.0072	0.6810 ± 0.0054	0.6833 ± 0.0051
Ω_m	0.3145 ± 0.0067	0.3246 ± 0.0091	0.330 ± 0.010	0.3277 ± 0.0072	0.3189 ± 0.0054	0.3166 ± 0.0051
r_d [Mpc]	147.13 ± 0.25	146.92 ± 0.47	146.20 ± 0.46	146.43 ± 0.34	147.06 ± 0.23	147.07 ± 0.22
σ_8	0.8099 ± 0.0051	0.8158 ± 0.0058	0.8171 ± 0.0055	0.8169 ± 0.0042	0.8132 ± 0.0042	0.8137 ± 0.0038

TABLE I. Λ CDM parameter constraints from different CMB experiments. We report mean values and 68% confidence intervals. Data sets are described in Table III. All data sets include T&E measurements, lensing reconstruction and a prior on τ_{reio} from *Planck* PR4 [32].

A. Data and map-making approach

The maps used in this work were constructed using SPT-3G data from the austral winter observing seasons of 2019 and 2020. The 1500 deg^2 SPT-3G Main field footprint spans declinations from -42 to -70 degrees and right ascensions from 20h40m0s to 3h20m0s, see Fig. 3. In the same figure, we also display the footprints of the SPT-3G Summer and SPT-3G Wide fields. The combination of these SPT-3G fields will constitute the total Ext-10k field, which will probe 25% of the sky with low noise and high resolution (see [41]). Additionally, Fig. 3 includes survey masks from other CMB experiments, such as ACT and *Planck*, which are used in the cosmological analysis of this work in combination with the SPT-3G Main field data.

The SPT-3G Main field that this work is based on is split into four subfields in declination (equivalent to elevation at the South Pole) to avoid detector nonlinearity from large changes in airmass without retuning the detectors. The data are acquired in the form of timestreams for each detector following an observing strategy that raster scans each subfield many times. We call one two-hour raster scan of each subfield an observation, and there were roughly 3000 total observations of the four subfields over the two years.

Following previous SPT analyses (see for example [12]), we use the filter-and-bin approach of map-making [42]. The timestreams are first low-pass filtered (to remove high frequency noise and reduce aliasing) and high-pass filtered (to remove low frequency sources of noise). The low-pass filtering is performed in Fourier space, while the high-pass filtering is achieved through fitting the individual detector timestreams to a set of low-

order polynomials and low-frequency sinusoids. The filtered timestreams are combined and projected into sky maps of the temperature (T) and linear polarization Stokes parameters (Q and U) in each frequency band for each individual observation of a SPT-3G Main subfield. Alongside the sky maps, we also build a weight map, which is a measure of the inverse variance of the observation map in each pixel. Some details of the data set and timestream processing are discussed in Section II of [16], and additional details are discussed in Q25. We note that [16] and this work are based on the same data and timestream processing but use maps with different pixelization schemes (see below).

In the following paragraphs, we highlight two important differences between the maps used here and those used in previous SPT analyses.

In previous SPT T&E power spectrum analyses, we pixelized the maps on a two-dimensional grid using a projection of the sphere, and we used two-dimensional Fourier transforms to substitute for spherical harmonic transforms (SHTs) under the flat-sky approximation. In particular, [16] used a map in the Lambert azimuthal equal-area projection (ZEA) with $0.56'$ pixels. In this new analysis (and the BB analysis in [17]), we pixelize the maps using HEALPix⁵ [22]. We choose an N_{side} parameter of 8192, which is equivalent to $0.4'$ resolution. With the size of the SPT-3G Main field and the sensitivity levels achieved in this work, the non-idealities caused by the flat-sky approximation are no longer negligible or trivial to account for. In particular, we found that projection effects were causing excess correlation between

⁵ <https://healpix.sourceforge.io/>

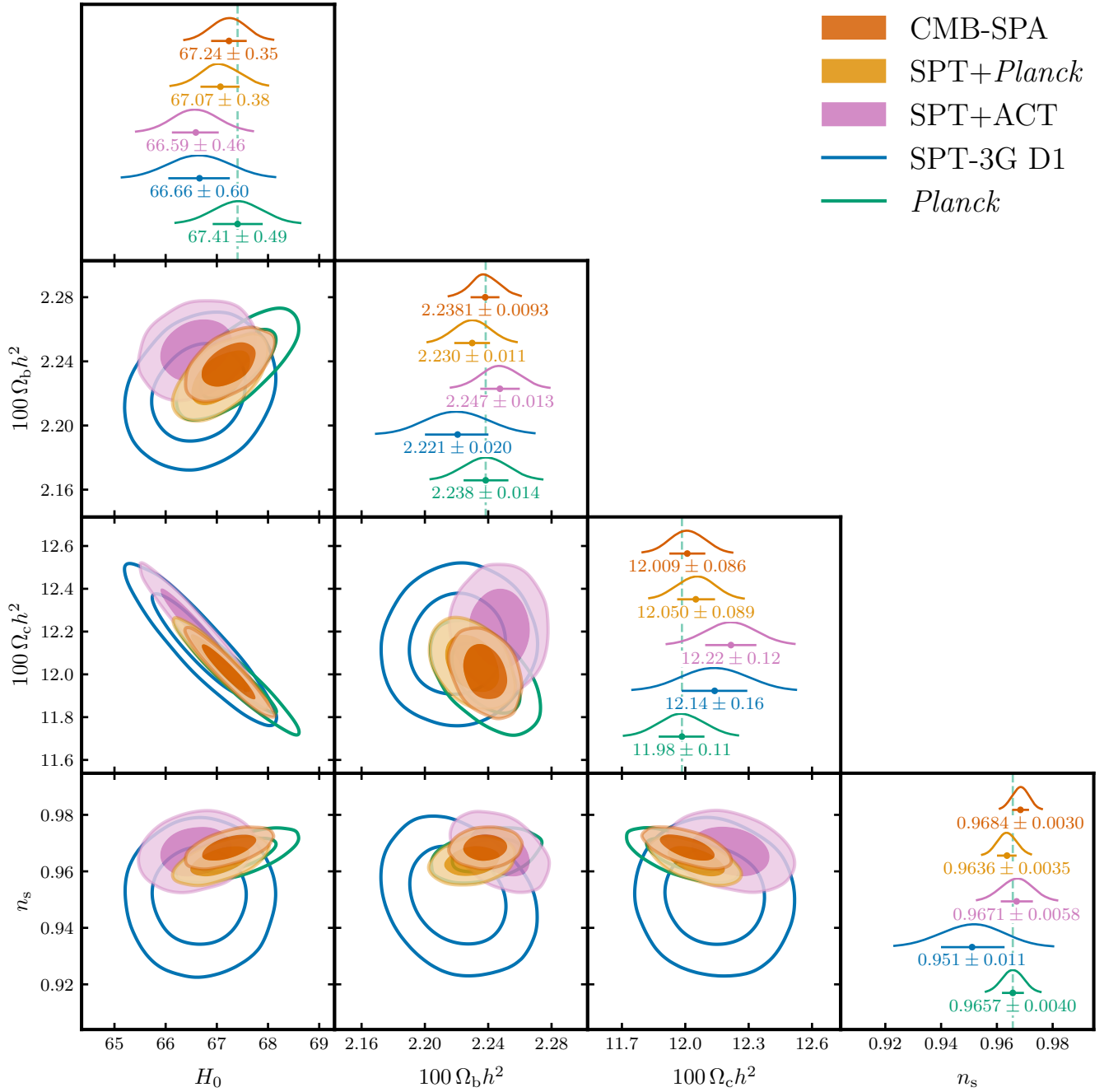


FIG. 2. Summary of CMB (primary and lensing) Λ CDM cosmological constraints. The data sets used are described in Table III, with SPT+ACT being the combination of (ground-based) SPT and ACT DR6, data, and CMB-SPA being the combination of SPT+ACT with *Planck*. The diagonal panels display the 1D posterior distributions of the parameters, with corresponding error bars. The off-diagonal panels show the 2D 68% and 95% confidence intervals.

power spectrum bins at high ℓ , in a manner that is difficult to compute analytically. For this reason, the current analysis is performed with full SHTs, and we take advantage of the many tools that have been developed for fast estimation of SHTs and resulting power spectra from HEALPix maps.

The filtering scheme is broadly similar to previous

work. In particular, as a result of the high-pass filtering, information at multipoles below 300 along the scan direction is absent from our maps, which results in a power spectrum bias that we evaluate and discuss in §IV A 2. An important difference involves a newly identified systematic effect from timestream filtering, one that only appears at the level of sensitivity of the data

in this work. As in most previous SPT analyses, we do not include timestream samples near bright emissive sources or strong Sunyaev-Zel’dovich-effect decrements from galaxy clusters when estimating the amplitudes of the polynomials and sinusoids to remove. This avoids creating extended features along the scan direction near strong sources, which are often referred to as “filtering wings.”⁶ We refer to this procedure of avoiding strong sources in the timestream filtering as “map-making masking” in the rest of this work. We mask all sources above 6 mJy at 150 GHz and any galaxy cluster detected at more than 10σ in a separate set of maps optimized for cluster detection, for a total of 2655 objects. While eliminating filtering wings is advantageous, in this work we detect small and previously undiagnosed side effects of this masking process. We refer to these as “filtering artifacts” which we describe in detail—along with our method for mitigating and accounting for them in our power spectrum estimation pipeline—in §IV A 1.

B. Coadds

We produce different weighted averages of single observation maps, in order to produce the different data products that are the baseline of the power spectrum pipeline. These are: the full depth coadd, the bundles, and the noise realizations or sign-flips. For each of these, we use the observation weight maps to weight the observation maps in the coadd.

We call the maps produced by the inverse-variance weighting of all of the observation maps at each frequency the full-depth coadds. Those maps are shown in Q25. We use them in the construction of the noise maps described below, as well as for inpainting emissive sources and galaxy clusters (§IV C).

The main product that is used in the power spectrum estimation are the bundles. We randomly group all of our observation maps at each frequency into 30 bundles, ensuring that the maps formed by the inverse-variance-weighted coaddition of all the maps of each bundle have a similar signal-to-noise. All bundles have similar noise levels but no noise correlation between them. We use the cross-spectra of these bundle maps to form the noise-unbiased spectra that are used in the rest of this analysis. This strategy is similar to that used in other CMB power spectrum analyses, which use various data split strategies to produce noise-unbiased power spectra ([6, 12, 34]). In contrast to most other CMB experiments, the highly redundant SPT observing strategy allows us, as in previous SPT analyses, to use a higher number of splits in order to allow for a more precise estimation of the noise covariance (see [12, 44]). While the depth of the 2019-2020 Main field data would allow us to produce

more than 30 bundles to improve the noise covariance estimate, we find that this number is a good compromise between accuracy and efficiency as the computing cost of increasing the number of bundles grows quadratically.

Finally, we call noise realizations, or sign-flips, the noise estimates formed by the difference between random selections of the observations:

$$N^i \equiv w_C^{-1} \sum_o \epsilon_o^i w_o (M_o - C), \quad (1)$$

where $\epsilon_o^i \in \{-1, 1\}$, M_o and w_o are the map and weight map for observation o , C is the full-depth coadd, w_C is the weight map of the full-depth coadd, and N^i is the i th noise realization. We form 500 of these noise realizations which are used in our simulations described in §III F. More details on the noise realization procedure are given in Q25.

C. Masks and point sources

For power spectrum analysis of the SPT-3G Main field, we use two different masks: a sky mask and a point source mask. The SPT-3G Main field binary sky mask is obtained by applying a threshold to the weight map, excluding regions where the weights are below 10% of the median across the map. The resulting footprint is shown in Fig. 3.

For power spectrum estimation, we further apodize the SPT-3G Main field mask using a 0.4° Gaussian taper. To produce the binary point source mask, we mask the 2655 emissive sources and clusters which have already been masked during map-making, using a radius where the signal-to-noise of beam-convolved sources falls to 1 (the same radius used in masking these objects during map-making), ranging from 2 to 15 arcmin. The point source mask removes 0.06% of the sky. We apodize this point source mask using a Gaussian taper with $\sigma = 0.03^\circ$. An important distinction from previous SPT analyses (and CMB analyses more generally) is that, when estimating signal power spectra from the maps, we do not use this point source mask; instead we replace the pixels where the sources are localized with a Gaussian constrained realization of the CMB informed by the rest of the map, in a process generally known as inpainting. The reason for this choice is the effect of the large number of point-source holes on the power spectrum covariance (§IV C).

The total apodized mask is the product of the apodized border mask and the apodized point source mask. This mask is used for computing the noise band powers, as our analysis and map inpainting rely on accurate knowledge of the noise within this region. While the sum of the weights in the apodized border mask corresponds to 4.0% of the sky, the sum of the weights in the total apodized mask is reduced to 3.8% as a result of point source masking and apodization.

⁶ An illustration of those can be found in Fig. 3 of [43].

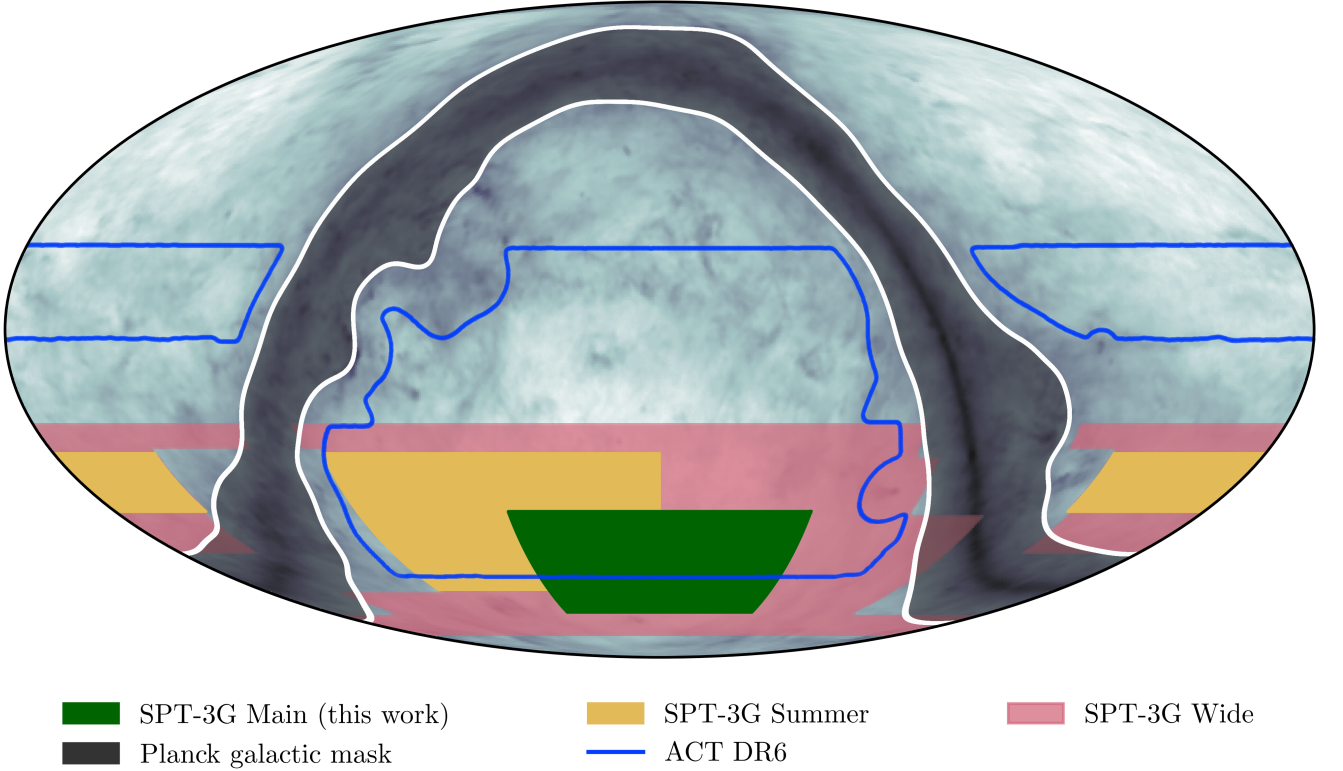


FIG. 3. Footprints of the SPT-3G surveys—Main (green filled area), Summer (orange filled area), and Wide (red filled area)—shown alongside the ACT DR6 survey mask (blue solid line). We show Galactic dust as measured by *Planck* in the background [45] and indicate the most contaminated region with the *Planck* GAL080 mask^a shown in gray shading with a white border. This work is based on observations of the SPT-3G Main field. The combination of the SPT-3G Main, SPT-3G Summer, and SPT-3G Wide fields will constitute the Ext-10k field, which will probe 25% of the sky with low noise and high resolution.

^a <https://pla.esac.esa.int/>

D. Calibration and cleaning

As discussed in Q25, the signal and noise maps at each frequency need to be recalibrated and cleaned to account for inaccuracies in the gain, polarization efficiency, and polarization angle of each detector. This process consists of four different operations, applied in the order listed below.

1. *Gain calibration*: We multiply all the T, Q, and U coadds by a common scaling factor to correct for the overall miscalibration of detector gains.
2. *Leakage from differential gain calibration*: We subtract small copies of a T coadd from the corresponding Q and U coadds to correct for the monopole temperature-to-polarization leakage due to differential gain miscalibration between detectors with different polarization angles.
3. *Polarization angle calibration*: We apply small rotations to the Q and U coadds to correct for the overall miscalibration of the detector polarization

angles. The calibration angle is determined by nulling the correlation between E and B modes in each frequency channel.

4. *Polarization efficiency calibration*: We multiply the Q and U coadds by a common scaling factor to correct for the overall miscalibration of the detector polarization efficiencies.

Note that in the case of gain and polarization efficiency calibration, the correction is done in two steps. First, we perform an external recalibration of the SPT-3G 150 GHz maps by cross-correlating them with the *Planck* PR3 143 GHz maps within the SPT-3G Main field mask, using the multipole range from 800 to 1200 for gain recalibration and from 400 to 1300 for polarization efficiency. To account for the filtering of the SPT map, the *Planck* map is mock observed through our data processing pipeline, in the same manner as the sky simulations used for transfer function modeling (see §III F). We note that this comparison uses only a small portion ($\sim 4\%$) of the total *Planck* PR3 143 GHz map and thus does not compromise our blindness to the *Planck*

data in the cosmological analysis. After the external recalibration of the SPT-3G 150 GHz map, we perform an internal recalibration of the SPT-3G 95 and 220 GHz maps by comparing them to the recalibrated 150 GHz map. The four steps summarized above are what [16] refers to as the alternative systematics estimates and discusses in some detail in its Appendix B. Additional information on these steps is presented in Q25.

We propagate the gain calibration and polarization efficiency calibration uncertainties in the likelihood analysis as described in §V B. Note that in most cases (except when using the polarization data alone), we ignore the external polarization efficiency priors and let this calibration be determined from the data within a given cosmological model. Similarly, we only use the internal gain calibration priors and internal polarization efficiency priors in specific validation tests and only explore the residual uncertainty of the internal recalibration under flat priors.

The polarization angle calibration and differential gain calibration do not correct for all of the sources of polarization leakage. We discuss in §IV B 2 how we model and correct for quadrupolar leakage induced by beam ellipticity correlated with detector polarization angle.

E. Map-level null tests

In addition to mitigating biases from known systematic effects, we search for potential unmodeled systematic errors in our maps through null tests. In a null test, we split the full set of individual-observation maps into two halves in a way that maximizes sensitivity to the suspected systematic error. We then subtract one half of the data from the other and check whether the difference is consistent with expectation spectra calculated from simulations to within noise fluctuations. The tests performed for this analysis include splits of the data based on the following: scan direction, detector wafer, date of observation, moon position, sun position, and observation azimuth.

The initial scan direction and wafer null tests failed our predetermined threshold for consistency with noise, and we traced both failures to excess power localized in spherical harmonic space. We choose to remove this excess power by applying a harmonic-space mask, defined as

$$W_{\ell m}^{\text{notch}} = \begin{cases} 0 & \text{if } \ell \in [500, 680] \text{ and } m \in [350, 425], \\ 1 & \text{otherwise.} \end{cases} \quad (2)$$

This masking is applied to the $a_{\ell m}$ coefficients of the data maps, as described in §IV A 1, and we account for it in the transfer function and covariance matrix estimations.

After applying this correction, all the TE and EE null spectra are consistent with expectations (see Section IV.B.1 of [16] for those spectra and more information on the null tests). Several TT null spectra formally fail our

pre-determined threshold, but they have amplitudes that are a tiny fraction of the sample variance error bars on the signal band powers, such that if we add 10^{-4} times the sample variance to the null spectrum covariance, all tests pass (see Q25). We consider this acceptable, as the potential systematic errors are too small to affect the cosmological analysis.

F. Simulations

In this analysis, we rely on simulations to propagate the effect of the data processing pipeline to our power-spectrum estimation and to validate our pipeline. When possible, we use analytical models to compute data products, which we validate with simulations.

All simulations are generated based on the *Planck*- Λ CDM cosmology, with added Gaussian realizations of extragalactic foreground components based on AGORA simulations [46] which have been calibrated to match the mid to high- ℓ portion of the measured spectra from data. Note that AGORA does not contain the reionization kSZ signal. We do not include Galactic foregrounds. We discuss in §IV A how we process these simulated skies to reproduce the effect of the data processing.

We also generate some simulations with an alternative underlying cosmological model to test our pipeline, which we describe in more detail in §VIC 3.

In most cases, we only simulate the signal part of the data. When needed (for inpainting in §IV C or for the validation of our pipeline in §VIC), we add a simulation of the noise contribution using the noise realizations (see §IIIB).

As an improvement over previous SPT analyses, we have developed a tool for fast modeling of the effects of the data processing on the power spectrum signal and covariance. This results in two types of simulations used in this analysis:

1. **Fullmock**: a set of 500 full-sky simulations that we use to calibrate the transfer function and validate the pipeline. This simulation set is similar to the one used in previous SPT analyses and described in Section IV.D.1 of [12] and in Q25. In this simulation set, the input skies are projected into individual-detector timestreams using the exact pointing and weights stored in the observation data files. These timestreams are then processed into maps with the same map-making pipeline used for the data, including the map-making masking procedure described in §III A even though there are no point sources in the simulations.
2. **Quickmock**: a set of 2000 fast simulations that we use to compute the covariance matrix and the transfer function. The **Quickmock** simulations use a lighter and faster simulation pipeline that implements several approximations to the scanning strategy to reduce computation while preserving

accuracy (E. Hivon et al., in preparation). We typically achieve percent-level accuracy or better with *Quickmock*.

Furthermore, inspired by the *CarPool* approach of [47], we use common CMB input skies for both *Fullmock* and *Quickmock* and can correct for any discrepancy between the two methods to obtain higher accuracy while limiting the impact of the Monte Carlo (MC) noise. We discuss in Appendix A a particular application of this method for the calibration of the power spectrum transfer function model and the propagation of the residual MC variance.

IV. POWER SPECTRUM

The data vector that we use in our likelihood analysis is formed by combining all of the available cross-bundle power spectra for each pair of frequency bands. To go from the maps described in Section III to the auto- and cross-frequency band powers requires a series of steps, which we summarize here and then describe in detail in the following sections.

First, the bundle maps are inpainted at the location

$$\hat{C}_\ell^{XY;\mu\nu} = \sum_{\ell'} K_{\ell\ell'}^{XY} \left[F_{\ell'}^{XY;\mu\nu} P_{\ell'}^2 B_{\ell'}^\mu B_{\ell'}^\nu C_{\ell'}^{XY;\mu\nu;\text{signal}} + A_{\ell'}^{XY;\mu\nu} + I_{\ell'}^{XY;\mu\nu} \right], \quad (3)$$

where $C_{\ell'}^{XY;\mu\nu;\text{signal}}$ is the underlying signal power spectrum, $B_{\ell'}^\mu$ is the instrument beam, $P_{\ell'}$ is the HEALPix pixel window function, and $K_{\ell\ell'}^{XY}$ is the residual kernel from the *Polspice* regularization. We describe in §IV A how the timestream filtering and the filtering artifacts around the masked point sources can be modeled by a multiplicative ($F_{\ell'}^{XY;\mu\nu}$) and an additive ($A_{\ell'}^{XY;\mu\nu}$) bias, respectively. We show in §IV C that the inpainting residual $I_{\ell'}^{XY;\mu\nu}$ can be treated as an additive correction.

In the following sections, we describe how we estimate each part of this data model necessary to debias our final band powers (see §IV D). At the end of this section, after accounting for all processing steps and the estimation of the noise power spectra (see §IV E), we describe the computation of the covariance matrix for the final band powers (see §IV F). The final band powers and covariance matrix are used in the next section to form our primary CMB likelihood (Section V).

of bright point sources and massive galaxy clusters with a process described in detail in §IV C. The bundle maps are then multiplied by the apodized sky mask described in §III C, and bundle cross-power spectra are estimated. For this operation, as well as for all power spectrum estimations throughout this work, we use *Polspice*⁷ [23, 24], a pseudo-power-spectrum framework similar to the *MASTER* algorithm [42] which corrects for mode coupling induced by the mask. As discussed in [25], this method regularizes the inversion of the *MASTER* matrix by apodizing the correlation function at the scales poorly explored in the mask.⁸ This regularization needs to be accounted for in the data model.

Taking into account the weights designed to mitigate excess power causing null test failures (see §III E) and filtering artifacts, the timestream filtering (see §IV A), the suppression or reweighting of spherical harmonic modes from the beam (see §IV B), the effect of pixelization, the residuals from the inpainting procedure (see §IV C), and the *Polspice* regularization, we form the following data model, which states that, on average, the measured estimation of the XY ($X, Y \in \{T, E\}$) power spectrum between two frequencies μ, ν ($\mu, \nu \in \{95, 150, 220\}$ GHz) can be related to the underlying signal by:

A. Filter modeling

During map-making, we apply a high-pass filter to the timestreams to remove large-scale noise and a low-pass filter to prevent aliasing, as described in §III A. These operations are performed at the level of the individual-detector time-ordered data and are by nature anisotropic and inhomogeneous. For the specific case of the SPT, operating at the geographical South Pole, the filtering affects spherical harmonic modes primarily as a function of m (assuming the map is in equatorial coordinates). However, as the filter cutoffs are defined in terms of absolute angle on the sky, not angle in right ascension, the effective cutoff in m varies with declination. This makes any attempt at a fully analytical estimation of the filtering challenging.

For this reason, in this work (as in past SPT analyses) we use MC simulations to estimate the effect of the filtering at the power spectrum level. We expand significantly on the approach from previous analyses to reach the level of precision required by the sensitivity of the new data set.

As before, we assume that a transfer function description is sufficient to capture the effect of the filtering on the signal power spectrum. In an extension of previous

⁷ <http://www2.iap.fr/users/hivon/software/PolSpice/>

⁸ We use a Gaussian taper with $\sigma_{\text{apo}}^{\text{PolSpice}} = 30^\circ$.

work, we allow for a different filtering correction at the covariance level (beyond the effect of mode loss), under the assumption that the effect is faithfully captured by a diagonal rescaling of the matrix. In estimating the effects on both the power spectrum and the covariance, we use the combination of **Quickmock** and **Fullmock** simulations to achieve the required precision on the calibration of the filter model.

We assume that, on average, the effects of our filtering (including artifacts) on an input signal at the power spectrum level can be approximated by the combination of a multiplicative and an additive bias

$$\left\langle C_\ell^{\text{fW};\text{XY};\mu\nu} \right\rangle_{\text{sims}} = F_\ell^{\text{XY};\mu\nu} \left\langle C_\ell^{\text{u};\text{XY};\mu\nu} \right\rangle_{\text{sims}} + A_\ell^{\text{XY};\mu\nu}, \quad (4)$$

where $C_\ell^{\text{u};\text{XY};\mu\nu}$ is the power spectrum of a simulation in the absence of filtering, $C_\ell^{\text{fW};\text{XY};\mu\nu}$ is that same spectrum in the presence of filtering, $F_\ell^{\text{XY};\mu\nu}$ is the transfer function, $A_\ell^{\text{XY};\mu\nu}$ represents the residual filtering artifacts after the ℓ, m weighting described in the next section, and all averages $\langle \dots \rangle_{\text{sims}}$ are over a set of MC simulations. All spectra are computed using the point source mask and corrected for mask effects with **Polspice**. We further added a W marker to $C_\ell^{\text{fW};\text{XY};\mu\nu}$ to indicate that we have applied an $a_{\ell m}$ filter when computing the power spectrum and need to take its effect into account along with the filter. We discuss these two biases in the following sections.

1. Filtering artifacts

One of the key features of our maps is the presence of a large number of point sources due to the high resolution of the instrument and the low noise level in the maps. As discussed in §III A, if the brightest point sources are included in the parametric fit used to high-pass filter the data, the resulting map has large scan-direction features near the locations of those sources, which we referred to as filtering wings in Section III. These features make it difficult to mask or inpaint over the sources in subsequent analyses. Removing the point source region from the parametric fit, a procedure we refer to as map-making masking, solves this issue, but at a price.

By coupling with the effective holes in the maps at the location of point sources, the filtering now introduces smaller spurious features in the map near the location of these sources; we refer to these as filtering artifacts. This effect leaks a fraction of the large-scale modes along the scan direction of the map that are targeted by the high-pass filter to small scales (also along the scan direction). The amplitude of these new features is on the order of the CMB signal that we are missing in the fit. This is much smaller than the filtering wings that have been avoided, the amplitude of which is on the order of 10% of the peak point-source amplitude; thus, the choice to mask point sources is a good trade-off for our data analysis.

Further details are provided in Appendix A 1. A map-level illustration is shown in Fig. 31.

The filtering artifacts, which depend on the location and size of the masked regions, can be described by an additive bias to the measured power spectrum. We perform an initial estimate of the bias from filtering artifacts using the difference of simulation pairs with and without map-making masking. This first estimate is performed to identify the angular scales at which the filtering artifacts are significant and to determine the range of multipoles that we need to target with our filtering and weighting, but is not used in the final analysis.

The artifacts are found to be well-localized in pseudo- $a_{\ell m}$ space, specifically at $m \lesssim 200$ (which is not surprising, as this is the range of modes targeted by the filter). We use this fact to mitigate the artifacts through a weighting of the pseudo- $a_{\ell m}$ coefficients. Again, assuming that the effect is on average well described by an additive effect, we decompose the signal $a_{\ell m}$ as a sum of filtered CMB (corresponding to our **Fullmock** simulations without map-making masking) and artifacts. Then, we can write the Wiener filter

$$W_{\ell m}^{\text{wiener}} \equiv \left\langle \frac{|a_{\ell m}^{\text{f;no MMmask}}|^2}{|a_{\ell m}^{\text{f}}|^2} \right\rangle_{\text{sims}}, \quad (5)$$

$$= \left\langle \frac{|a_{\ell m}^{\text{f;no MMmask}}|^2}{|a_{\ell m}^{\text{f;no MMmask}} + a_{\ell m}^{\text{artifacts}}|^2} \right\rangle_{\text{sims}}, \quad (6)$$

where $a_{\ell m}^{\text{f;no MMmask}}$ is the pseudo- $a_{\ell m}$ of a **Fullmock** simulation without map-making masking and $a_{\ell m}^{\text{f}}$ is the pseudo- $a_{\ell m}$ of the same simulation with the standard map-making masking procedure. We perform the MC estimation on 110 pairs of **Fullmock** simulations and estimate the pseudo- $a_{\ell m}$ using the point source mask.

We combine this new filter with the notch filter, see Eq. (2), such that

$$W_{\ell m}^{\text{total}} \equiv W_{\ell m}^{\text{notch}} W_{\ell m}^{\text{wiener}}. \quad (7)$$

We apply this weighting, $W_{\ell m}^{\text{total}}$, to mitigate both the excess power contributing to initial null test failures and the filtering artifacts in the power spectrum estimation. This procedure reduces the amplitude of the filtering artifacts by a factor of ~ 10 . Of course, this weighting also affects the measured power spectrum and covariance, and we include this contribution when accounting for the filtering multiplicative effect in the next section.

After application of the Wiener filter, despite the strong reduction in the filtering artifact power, there is still a detectable residual effect on the measured power spectra. We account for this residual bias with a template subtraction approach at the power spectrum level. The shape of the template for the contribution of the residual artifacts to the measured power spectra, which we denote as A_ℓ in Eq. (3), is determined by the large-scale power in the map, the position and size of holes in the maps, and

the timestream filtering strategy and the Wiener filter. This final A_ℓ is estimated using

$$A_\ell = \left\langle C_\ell^{\text{fWt};i} - C_\ell^{\text{fWt};i;\text{no MMmask}} \right\rangle_{\text{sims}}, \quad (8)$$

with similar notation to Eq. (5), and using the marker “Wt” to note that we are now using the total filter in Eq. (7) and the marker “i” to note that the maps have been inpainted, as detailed later in §IV C. We show the templates for the various spectra and band combinations, along with the inpainting bias templates discussed in §IV C, in Fig. 4. The template for residual filtering artifacts in TE and EE is $< 0.1\sigma$ at all scales; for TT, it is negligible on large angular scales but reaches 0.3σ at $\ell = 3000$. While we use these templates to clean simulations and data, we also verify that leaving them uncleaned would result in negligible ($< 0.1\sigma$) biases on cosmological parameters; we thus conclude that our power spectrum pipeline is robust to our modeling of the filtering artifacts.

2. Transfer functions

As discussed at the beginning of this section, the combined effects of the filtering, the $a_{\ell m}$ weighting, inpainting, and the mask can result in multiplicative and additive biases and mix power between multipoles. We also noted that in our modeling of the power spectra we assume that effects from the mask can be dealt with separately from the rest. We further assume that the multiplicative bias from the filtering and $a_{\ell m}$ weighting can be treated as a one-dimensional function of ℓ , i.e. a transfer function.

Similarly to the filtering artifact estimation performed above, we rely on simulations to estimate the transfer function. We define the transfer function in terms of an average ratio of the (weighted) power spectrum of filtered simulations without map-making masking $C_\ell^{\text{fWt},XX;\mu\nu;\text{no MMmask}}$ and the (unweighted) power spectrum of unfiltered simulations $C_\ell^{\text{u},XX;\mu\nu}$:

$$F_\ell^{XX;\mu\nu} \equiv \frac{\left\langle C_\ell^{\text{fWt},XX;\mu\nu;\text{no MMmask}} \right\rangle_{\text{sims}}}{\left\langle C_\ell^{\text{u},XX;\mu\nu} \right\rangle_{\text{sims}}}. \quad (9)$$

In the case of TE, the denominator of Eq. (9) can vanish and, therefore, we define the corresponding transfer function as the geometric mean of the auto-spectrum transfer functions

$$F_\ell^{\text{TE};\mu\nu} \equiv \left(F_\ell^{\text{TT};\mu\mu} \times F_\ell^{\text{EE};\nu\nu} \right)^{1/2}. \quad (10)$$

We show the 150×150 GHz transfer functions in Fig. 5. (The transfer functions at other frequency combinations look similar.) The effect of our filters is important: a strong suppression at the lower range of multipoles

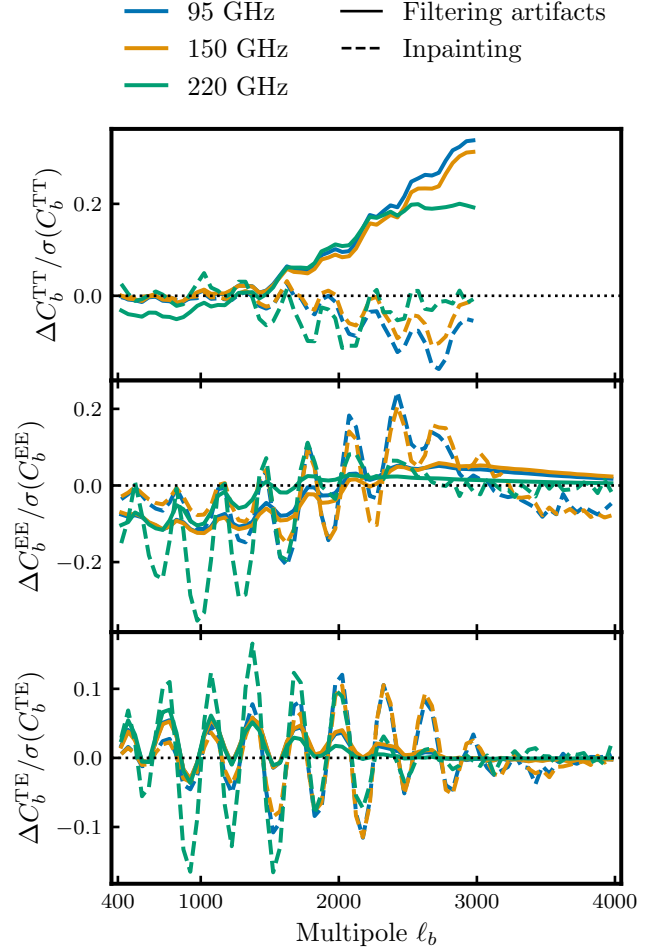


FIG. 4. Filtering artifacts (solid lines) and inpainting biases (dashed lines) at 95 GHz (blue), 150 GHz (orange), and 220 GHz (green). We remove both effects from the power spectrum estimates using a template subtraction approach. The filtering artifacts are the result of masking bright sources in the high-pass filter applied to the maps and the inpainting biases are due to the inpainting of the masked point sources. The filtering artifacts are negligible on large angular scales but reach up to 0.3σ at $\ell = 3000$. The inpainting biases are the largest in EE but always remain below 0.3σ .

owing to the high-pass filter is evident, as is a sharp cut induced by the notch filter, see Eq. (2). While the high-pass is a sharp cut in the Fourier conjugate to scan-direction angle, the dependence of the effective m cutoff on declination in the map translates this into a softer cutoff as a function of m . Effectively averaging over m at each multipole ℓ results in the high-pass filter transfer function being quite broad in one-dimensional ℓ space. We find that, up to the effect of the notch filter, the transfer function is close to a simple model $F_\ell^{\text{mod}} = 1 - 200/\ell$. This can be understood by looking at Eq. (B11) of [42] and considering that the SPT scanning strategy consists of parallel scans. At small angular scales, the transfer function approaches unity.

The EE transfer function differs from the TT transfer function because of residual leakage from large to small scales. This leakage is due to the cut-off, it is unrelated with filtering artifacts, and it is relatively more significant in EE than in TT, owing to the presence of small-scale foregrounds in TT and their absence in EE.

We achieve excellent precision for this MC estimation, as can be seen in the lower panel of Fig. 5, thanks to our joint use of the 2000 **Quickmock** simulations and the use of the 500 **Fullmock** simulations to calibrate them. We discuss in Appendix A how we use both simulations to achieve the best possible precision and how we propagate residual errors from the MC estimate.

The filtering, masking, and reweighting of the $a_{\ell m}$ also affect the covariance of the power spectra in a non-trivial way, potentially modifying the off-diagonal structure of the covariance. Similarly to the case of the power spectrum, we assume that the mask and filtering effects decouple and we discuss how we deal with the former in §IV F. For filtering and $a_{\ell m}$ reweighting, see Eq. (7), we again assume that the multiplicative effect is dominant compared to the coupling of multipoles and approximate the effect of the filters by a diagonal rescaling (see Eq. (35) in the covariance section). We estimate this rescaling from our simulations using the same method we used for the transfer function:

$$H_{\ell}^{XY\mu\nu;X'Y'\alpha\beta} \equiv \frac{\left\langle \delta C_{\ell}^{\text{fWt};\text{no MMmask};XY\mu\nu} \delta C_{\ell}^{\text{fWt};\text{no MMmask};X'Y'\alpha\beta} \right\rangle_{\text{sims}}}{\left\langle \delta C_{\ell}^{\text{u};XY\mu\nu} \delta C_{\ell}^{\text{u};X'Y'\alpha\beta} \right\rangle_{\text{sims}}}, \quad (11)$$

where

$$\delta C_{\ell}^{XY;\mu\nu} \equiv C_{\ell}^{XY;\mu\nu} - \left\langle C_{\ell}^{XY;\mu\nu} \right\rangle_{\text{sims}} \quad (12)$$

is the discrepancy between a random realization of the power spectrum and its MC average. A similar method was used in [48] in the case of inhomogeneous survey depth.

The power spectrum transfer function $F_{\ell}^{XY;\mu\nu}$ and the multiplicative correction $H_{\ell}^{XY,X'Y';\mu\nu}$ are, respectively, quadratic and quartic in mode filtering, and one would therefore expect $H_{\ell}^{XY,X'Y';\mu\nu} \sim (F_{\ell}^{XY;\mu\nu})^2$. However, since the high pass filtering implemented here is by nature mostly binary, meaning that the frequency modes along the scan direction are either left unchanged or totally removed, we expect the effects of the filtering on the mean and variance of the power spectrum to be similar. Indeed, using the large number of **Quickmock** simulations, we found that, at the $\sim 10\%$ level,

$$H_{\ell}^{XY,XY} \simeq F_{\ell}^{XY}. \quad (13)$$

This was confirmed in **Fullmock** simulations, with a larger scatter. We show how we integrate this correction in our estimation of the covariance matrix in §IV F.

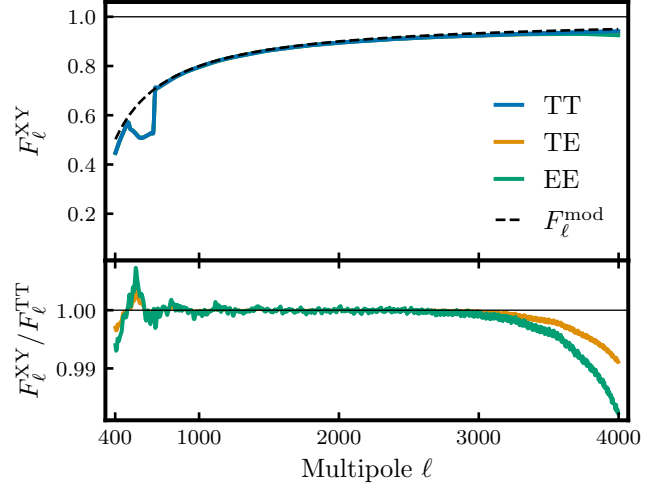


FIG. 5. One-dimensional transfer function for the SPT-3G D1 T&E data at 150×150 GHz. The transfer function is computed using **Quickmock** simulations not masking point sources during map-making and corrected for the measured discrepancy between **Quickmock** and **Fullmock** simulations. *Upper panel:* The transfer function for the TT (blue), TE (orange), and EE (green) spectra. The lines overlap for most of the multipole range and are just distinguishable above $\ell = 3000$. Except for masking due to the notch filter in Eq. (2), the transfer function is close to a simple model $F_{\ell}^{\text{mod}} = 1 - 200/\ell$. *Lower panel:* Ratio of the TE and EE transfer functions to the TT transfer function.

B. Beams

The beam, or point-spread function, describes the angular dependence of the instrument response to a point source. For microwave instruments (for which atmospheric seeing is not a significant contribution), the beam is primarily determined by diffraction in the optical system but can also include effects such as finite detector response time. The effect of the beam on sky signal can be represented as a convolution in real space. The effective beam is well described by a single transfer function in harmonic space, represented by the B_{ℓ} terms in Eq. (3). The SPT-3G beams are described in Appendix 4 of [16], and a more complete characterization will be given in an upcoming work (N. Huang et al., in preparation, hereafter H25). Here, we only review the aspects of the beam modeling that are relevant to our analysis.

1. Temperature beams

We measure the angular response in total intensity to an unpolarized source, or the temperature beam, using the combination of bright active galactic nuclei (AGN) located in the Main field and dedicated observations of Saturn. While bright AGN allow us to characterize

the main beam close to the peak of the response, they are too faint to map the telescope beam far from the center, i.e. the beam sidelobes. Meanwhile, planet observations are useful for mapping the sidelobes, but planets are sufficiently bright to saturate some detectors and thus cannot be used for the main beam. We stitch together observations of Saturn and individual AGN into a composite real-space beam map. A harmonic decomposition of this map gives us the temperature version of the B_ℓ^μ term in Eq. (3), which we notate as $B_\ell^{\text{T};\mu}$. We normalize the beam to unity at $\ell = 800$ to decorrelate the beam shape from the *Planck* based calibration, see §III D. One-dimensional harmonic-space temperature beams $B_\ell^{\text{T};\mu}$ are plotted in Fig. 6.

We quantify the uncertainty on $B_\ell^{\text{T};\mu}$ related to noise, systematic effects, and analysis choices (such as the radius at which the Saturn and AGN beams are stitched) in a beam covariance matrix, which covers all angular scales and all observational frequencies. We use it to propagate the beam uncertainty to the band powers in the signal part of our data model as shown in Eq. (40). The uncertainty in the temperature beam represents at most 0.2% of the beam at $\ell = 3000$ for all frequencies. We display the uncertainty on the temperature beams in Fig. 6. More details on the uncertainty estimation are given in H25.

2. Quadrupolar beam leakage

The “leakage beam” describes the response in polarization to an unpolarized source. The monopole leakage from temperature to polarization caused by gain differences between detectors has been removed from the maps in a previous analysis step, see §III D. However, the maps also contain a significant contribution from higher-order quadrupolar leakage sourced by the differential beam ellipticities of the detectors [49–51]. This effect is detected and measured by analyzing Q and U maps at the location of bright sources in temperature, which we present in Appendix B 1. We propagate the measured map-level contamination to band powers using an analytical model derived from Eq. (B1), which we confirm with simulations. We model quadrupolar leakage at the band power level as

$$\begin{aligned} C_\ell^{\text{TE};\mu\nu;\text{leak}} &= \epsilon_2^\nu \sigma_\nu^2 \ell^2 C_\ell^{\text{TT};\mu\nu}, \\ C_\ell^{\text{EE};\mu\nu;\text{leak}} &= \epsilon_2^\mu \sigma_\mu^2 \ell^2 C_\ell^{\text{TE};\mu\nu} + \epsilon_2^\nu \sigma_\nu^2 \ell^2 C_\ell^{\text{ET};\mu\nu} \\ &\quad + \epsilon_2^\mu \epsilon_2^\nu \sigma_\mu^2 \sigma_\nu^2 \ell^4 C_\ell^{\text{TT};\mu\nu}, \end{aligned} \quad (14)$$

where ϵ_2^μ is the amplitude of the quadrupolar leakage and σ_μ is the size of the leakage beam in the μ band. From the map-level study described in Appendix B 1 we deduce

the amplitudes of the quadrupolar leakage:

$$\epsilon_2^{95} = -(0.65 \pm 0.11)/100, \quad (15)$$

$$\epsilon_2^{150} = -(1.2 \pm 0.21)/100, \quad (16)$$

$$\epsilon_2^{220} = -(2.3 \pm 0.66)/100, \quad (17)$$

which are, respectively, 6, 6, and 3.5σ detections of the quadrupolar leakage. In Fig. 7, we display the leakage template, which corresponds to contamination amplitudes of $\sim 1\%$ of the signal on large scales in TE and, with respect to error bars, affects mostly the highest range of multipoles of this analysis, between $\ell = 3000$ and 4000. We find that the leakage is largest at 220 GHz due to the higher foreground power in TT leaking into TE and EE. The leakage is smaller, and less significant, in EE.

We note in Appendix B that the quadrupolar leakage correction was implemented only after unblinding (for details of the blinding procedure, please refer to §VI A). This is one of the reasons why the band powers are not debiased for this effect, which is instead treated at the likelihood level. The fit to the map-level measurement is used as a prior for the beam leakage parameters.

3. Polarized beams

We define the polarized beam as the angular response of the instrument in the linear-polarization Stokes parameters to a 100% linearly polarized source. In principle, this can be different from the temperature beam. The main beam, which is formed by optical paths that are well-controlled and pass through all optical elements as designed, is expected to be uniformly and highly polarized, but the polarization of the sidelobes, which include radiation scattered and reflected from non-ideal optical elements and the diffraction pattern formed by the gaps and surface imperfections in the telescope mirrors, can be more complex. We do not have sufficiently bright polarized sources in the survey region to directly measure the polarized beam on the relevant angular scales. For this reason, we allow for depolarization of the beam sidelobes relative to the main beam and marginalize over the fraction of sidelobe polarization at the likelihood level. This additional freedom slightly degrades the constraining power of the data set. In Appendix B 2 we show that this polarized beam model is preferred by the data. We emphasize that the evidence for sidelobe depolarization comes from requiring internal consistency between frequency bands in the EE data. Its detection does not require assuming any particular cosmological model.

We model the polarized beam as the sum of the main beam $B^{\text{main};\mu}$ and the sidelobe contribution $B_\ell^{\text{sidelobe};\mu}$ modulated by a scale-invariant polarization fraction β_{pol} :

$$B_\ell^{\text{P};\mu} = \frac{B_\ell^{\text{main};\mu} + \beta_{\text{pol}}^\mu B_\ell^{\text{sidelobe};\mu}}{B_{800}^{\text{main};\mu} + \beta_{\text{pol}}^\mu B_{800}^{\text{sidelobe};\mu}}. \quad (18)$$

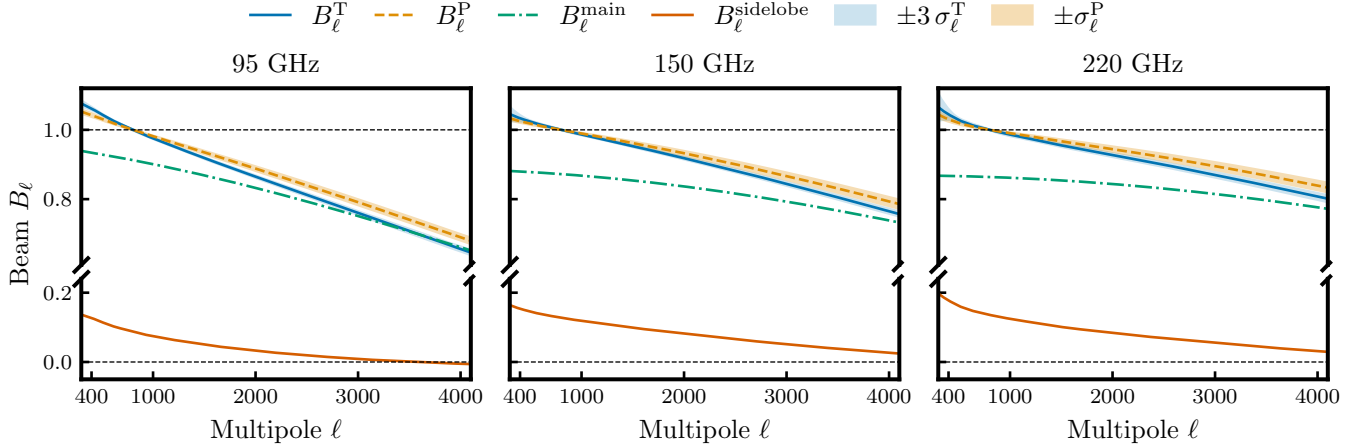


FIG. 6. Harmonic-space beam functions for each frequency band. The solid blue lines show the measured temperature beams, with shaded regions indicating three times the measurement uncertainty. Note that the beam uncertainty is correlated across both angular scales and frequency bands. The orange dashed lines represent the polarized beams, which are modeled as the sum of the main beam and a fractional sidelobe contribution, as described in Eq. (18). The sidelobe beams (full red line) are obtained by subtracting the main beam (green dash-dotted line) from the temperature beam, following Eq. (19). We display the best-fit polarized beams using the parameters from Table IX, as determined by fitting the SPT-3G D1 T&E data. The orange shaded regions indicate the uncertainty on the polarized beams, which is propagated from the uncertainty in the sidelobe polarization fraction. This uncertainty is correlated across both angular scales and frequency bands. Both the temperature and polarized beams are normalized to unity at $\ell = 800$.

When $\beta = 1$ we recover the temperature beam, and if the sidelobes are depolarized we expect to recover $\beta < 1$. The main beam $B_{\ell}^{\text{main};\mu}$ is calculated analytically from our knowledge of the optics, as described in Appendix 4 of [16], and shown in Fig. 6. The sidelobe contribution is taken as the difference between the measured temperature beam and the calculated main beam,

$$B_{\ell}^{\text{sidelobe};\mu} \equiv B_{\ell}^{\text{T};\mu} - B_{\ell}^{\text{main};\mu}. \quad (19)$$

We normalize the harmonic space polarized beam to unity at $\ell = 800$ to preserve the relative polarization efficiency priors obtained by comparison with *Planck* independent of changes to the assumed beam shape. As discussed in §III D, the external polarization efficiency calibration is only needed when using polarization data alone.

Similarly to §IV B 2, the baseline polarized beam model was implemented only after unblinding and we similarly fit for it at the likelihood level, see Eq. (40). We find strong support for this model with cosmology-independent methods, see [16] and the SPT-*lite* discussion in Appendix B 2. In the context of Λ CDM, the data strongly support the polarized beam model with a 5σ preference for $(\beta_{\text{pol}}^{95}, \beta_{\text{pol}}^{150}, \beta_{\text{pol}}^{220}) \neq (1, 1, 1)$ from SPT data alone. Posterior values for the β_{pol} parameters are given in Table IX. In Fig. 6, we show the best-fit polarized beams and their propagated uncertainties, which are derived from the uncertainty in the sidelobe polarization fraction β_{pol} . At small angular scales ($\ell > 2000$), these

beams are nearly identical to the temperature beam, except for the effect of normalization at $\ell = 800$. At large angular scales ($\ell < 2000$), however, the sidelobe contribution alters the polarized beam shape relative to the temperature beam.

C. Inpainting

We mentioned in §III A that the unprecedented depth of the SPT-3G Main field and the high resolution of the SPT results in many thousands of emissive sources and galaxy clusters detected at high signal-to-noise. We choose to mask over 2000 of them in map-making and to remove them from the map before estimating power spectra. Traditionally this has been accomplished by multiplying the map by a mask with (apodized) holes at the location of all sources and clusters we wish to remove. Under the assumption of statistical isotropy of the signal, such features in the mask do not bias the power spectrum estimation, as the mask effect can be properly taken into account in the MASTER framework. However, point-source masking strongly impacts the statistical properties of the estimator by inducing correlations in the estimated power spectrum across angular scales. In particular, we see correlation of large-scale power to smaller scales: bigger modes across the patch look like small-scale modes when interrupted by the point source holes and can be disambiguated from the real small-scale modes only on average. Using the Fullmock simulations, we show

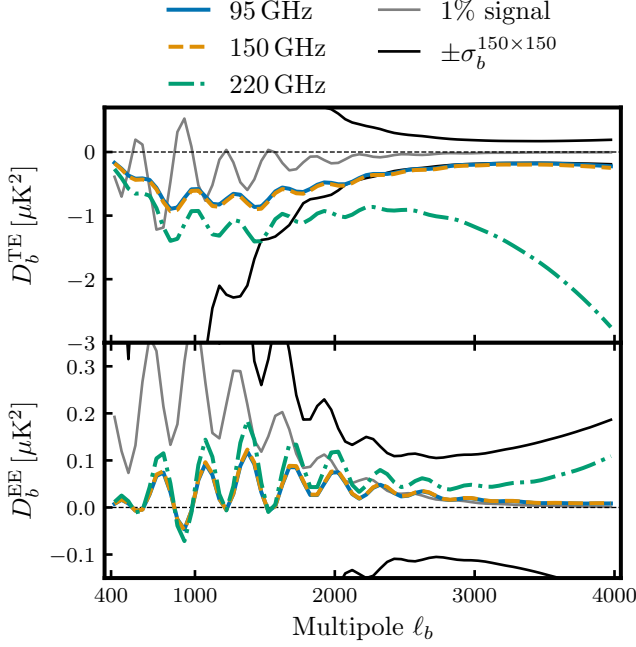


FIG. 7. Quadrupolar leakage templates for the TE (top) and EE (bottom) spectra. The colored lines indicate the leakage spectra, see Eq. (14), for the three frequency bands. The solid gray line represents 1% of the best-fit CMB power spectrum, while the solid black lines represent the 150×150 GHz error bars. The TE 150×150 GHz leakage reaches 1σ at small scales. The contamination is notably larger at 220 GHz, primarily due to the higher foreground power in TT leaking into TE and EE, however, the associated frequency channel error bars are also larger.

that we can expect point-source masking to increase the variance of the estimator on the three T&E channels at $\ell = 3000$ by 50%, see Fig. 8.

To avoid the consequences of point source masking on the statistical properties of our power spectrum estimator, we choose to fill the locations of point sources and galaxy clusters in our maps with Gaussian realizations constrained over the power in the rest of the map, a method also known as inpainting [52, 53]. This allows us to use only the border mask, which does not have any small-scale features, for power spectrum estimation, thereby reducing the mode-coupling of the estimator. However, we need to increase the covariance of our power spectra to account for the fact that we have added fake simulated signal to the data.

Gaussian-constrained inpainting is now a well-established technique in CMB lensing analysis [15, 36, 54–59]. This work is the first use of it for primary CMB analysis, which presents different challenges.

We follow here the method presented in [15, 54, 55] and only use a small region around each point source to predict the CMB signal inside the region we wish to inpaint. Even with this simplification, the inpainting of the SPT-3G maps is challenging due to the large number of point sources and the high resolution of the maps.

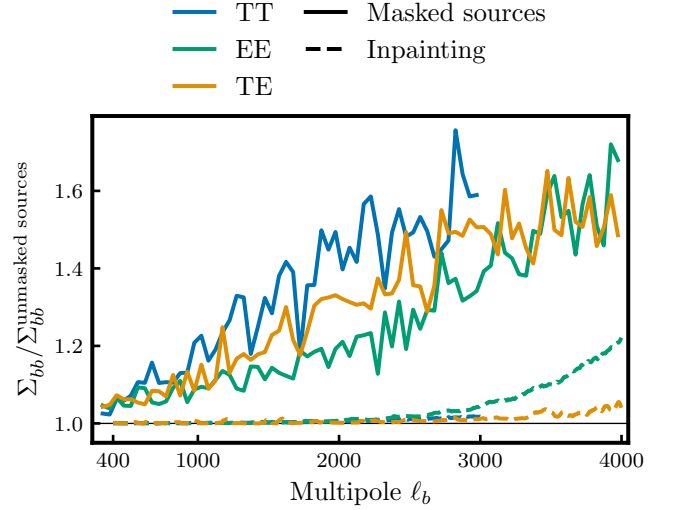


FIG. 8. Ratio of the covariance matrix diagonal elements with and without point source masking. The ratio is computed from 500 Fullmock simulations, for which we either mask (or not) the point sources during power spectrum estimation. The source masking results in 50% more variance at small scales than without. In comparison, we show the increase in variance due to inpainting, see Eqs. (26) and (36), which is much smaller, justifying our choice to use the inpainting method.

For this work, we improved on the inpainting code called *cork* used in the *Planck* lensing analysis [54], extending it to apply to multiple frequencies and polarization data and perform efficiently in the high-resolution regime ($N_{\text{side}} = 8192$).

Schematically, the method proposes to create a new set of T, Q, and U bundle maps M^i from the M observed bundle maps such that

$$\begin{bmatrix} M_{\text{msk}}^i \\ M_{\text{brd}}^i \\ M_{\text{rst}}^i \end{bmatrix} = \begin{bmatrix} 0 & W & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} M_{\text{msk}} \\ M_{\text{brd}} \\ M_{\text{rst}} \end{bmatrix} + \begin{bmatrix} 1 & -W & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} M_{\text{msk}}^{\text{sim}} \\ M_{\text{brd}}^{\text{sim}} \\ M_{\text{rst}}^{\text{sim}} \end{bmatrix}, \quad (20)$$

where M_{msk} is the set of pixels of the observed bundle maps inside the point source masks, M_{brd} is the set of pixels in the constraining zone, and M_{rst} is the set of pixels in the rest of the bundle maps outside of the mask and the constraining zone. M^{sim} is a set of simulated bundle maps that have the same power spectrum and noise as the observed one. Finally, following [54], W is the Wiener filter that predicts the data inside the masked region from the border region and is given by

$$W = \Sigma_{\text{msk}, \text{brd}} \Sigma_{\text{brd}, \text{brd}}^{-1}, \quad (21)$$

the two covariances here being the joint pixel covariances of the bundle of maps. When all of the remaining data is used as a constraining region and the size of the masked

region is much smaller than the total map, the power spectra of the inpainted map bundle are unbiased.

In full generality, the pixel covariance Σ must couple pixels across bundles (each bundle sees the same sky up to noise), frequencies (each frequency band sees the same CMB up to foregrounds), and temperature and polarization (because of the cosmological TE correlation). The resulting W matrix is large, dense, and based on the inversion of the also large and full $\Sigma_{\text{brd,brd}}$. To reduce the computational complexity of the W matrix and enable parallelization, we adopt several simplifying assumptions that decouple its components. These approximations introduce a residual inpainting-induced correlated noise, which we calibrate using simulations and correct for in the power spectrum analysis.

First, following the lensing analysis procedure, we reduce the constraining region to a small border around each point source hole of the mask. This simplification allows us to decouple the problem by inpainting region and parallelize over the list of regions. While our tests show that constraining regions of approximately degree scale are needed to leave a vanishingly small residual, a smaller constraining region corresponding to a few tens of arcminutes, which we adopt, is a sufficient trade-off between computation time at $N_{\text{side}} = 8192$ and amplitude of the residual.

Second, instead of inpainting all of the bundles together taking into account their correlations, we instead inpainted only the full-depth coadd (see §III B) and used it to fill the masked regions in each bundle. This correctly takes into account the correlation between bundles. This simplification does not increase the inpainting residual.

Third, we assume that the different frequencies are uncorrelated and inpaint each frequency separately. The assumption of uncorrelated frequencies induces biases in the cross-frequency spectra that we estimated.

After those assumptions, the problem can be reduced to a suite of parallelizable Gaussian constrained realizations for each inpainting region on a triplet of T, Q, and U maps using small constraining regions.

We checked that a perfect knowledge of the exact power spectrum of the map was not a strong requirement for the building of the Wiener filters. Eventually, we used the mean of the **Fullmock** power spectra to compute the filter. This directly includes the transfer function correction. The noise contribution needed for the filter is obtained from the data noise spectra (discussed in §IV E).

The quality of the simulation used for the inpainting (M^{sim} in Eq. (20)) is paramount. In particular, since we only account for the timestream filtering in the pixel covariances at the 1D transfer function level, we miss the anisotropic effects of the filtering. For this reason, to inpaint the maps, we use **Fullmock** simulations plus noise realizations to include the full filtering effect and our best noise model.

Finally, we estimate the residual inpainting bias $I_{\ell'}^{XY;\mu\nu}$ arising from all of those approximations by inpainting the set of **Fullmock** plus sign-flip simulations with the same

settings, which we define as

$$I_{\ell'}^{XY;\mu\nu} = \left\langle C_{\ell}^{\text{fWt};i} - C_{\ell}^{\text{fWt}} \right\rangle_{\text{sims}}. \quad (22)$$

This bias is displayed in Fig. 4. It is found to be of similar order of magnitude to the filtering artifact biases and we remove it from the computed power spectra.

The inpainting procedure must be taken into account in the covariance of the power spectra, as a small part of the maps has been replaced with fake simulated signal. We must marginalize over it in the estimation of the covariance matrix of the band powers. To do so, following Eq. (20), we first note that the power spectra of the inpainted maps can be split into a data component and a random component, both Wiener-filtered,

$$C_{\ell}^i = C_{\ell}^{W;\text{data}} + C_{\ell}^{1-W;\text{sim}}. \quad (23)$$

Up to the inpainting bias I_{ℓ} (that we neglect in this discussion), C_{ℓ}^i is by construction unbiased and its covariance Σ is given by the calculation that we detail in §IV F. If we fix the data part and explore the covariance of C_{ℓ}^i under the variation of simulations, we expect that

$$\langle \delta C_{\ell}^i \delta C_{\ell'}^i \rangle_{\text{sims}} = \left\langle \delta C_{\ell}^{1-W;\text{sim}} \delta C_{\ell'}^{1-W;\text{sim}} \right\rangle_{\text{sims}}. \quad (24)$$

The marginalization over the random fake data can be approximated by adding to Σ an extra term corresponding to the covariance of $C_{\ell}^{1-W;\text{sim}}$. Our problem is reduced to the computation of the covariance of the Wiener-filtered fake data spectra $C_{\ell}^{1-W;\text{sim}}$. We assume that a 1D transfer function sufficiently captures the effect at the power spectrum level, so that

$$\langle \delta C_{\ell}^{1-W} \delta C_{\ell'}^{1-W} \rangle_{\text{sims}} \approx \rho_{\ell} \rho_{\ell'} \langle \delta C_{\ell} \delta C_{\ell'} \rangle_{\text{sims}}, \quad (25)$$

$$\text{with } \rho_{\ell} \equiv \frac{\langle C_{\ell}^{1-W} \rangle_{\text{sims}}}{\langle C_{\ell}^i \rangle_{\text{sims}}}, \quad (26)$$

which we estimate from the inpainted **Fullmock** simulations. We show the impact of inpainting on the diagonal of the covariance matrix in Fig. 8. The inpainting covariance is significantly smaller than the masked point source covariance, as intended by the inpainting procedure. We employ this methodology to increase the band power covariance matrix, thereby marginalizing over the artificial simulated signal introduced into the maps, see §IV F.

D. Band powers

After inpainting the maps, we estimate the auto- and cross-frequency pseudo-power spectra from the 95, 150, and 220 GHz maps on the curved sky using **Polspice** and the harmonic-space filter defined in Eq. (7). Each spectrum is calculated as the average of the cross-power

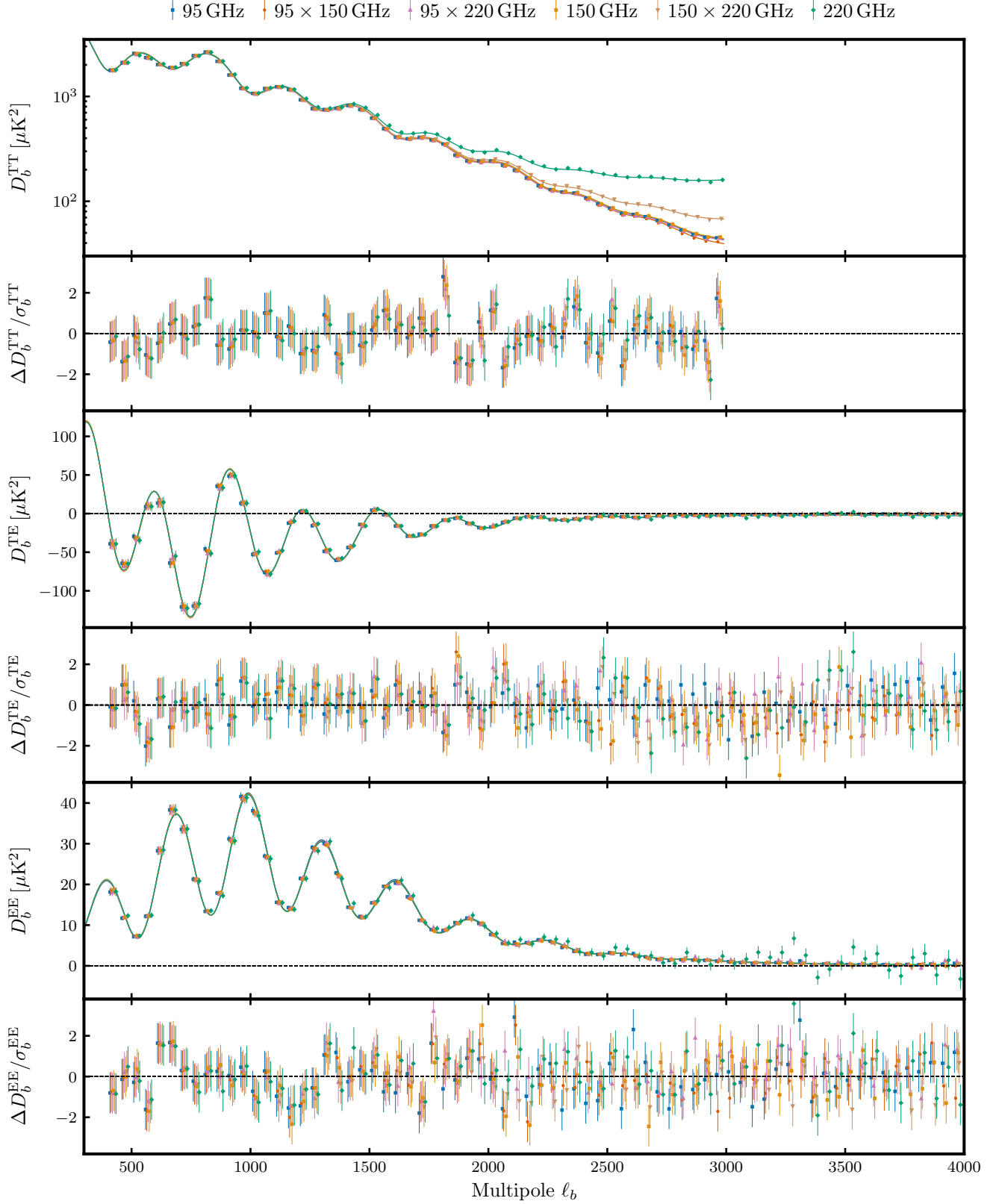


FIG. 9. *Large panels:* Measured SPT-3G D1 T&E band powers for each cross-frequency spectrum, together with the best-fit theoretical prediction under the Λ CDM cosmological model, incorporating the complete data model as described in Eq. (40). For clarity, the 150×95 GHz, 220×150 GHz, and 220×95 GHz TE band powers are omitted. The error bars represent the square roots of the diagonal elements of the covariance matrix as computed in §IV F. *Small panels:* Corresponding residuals with respect to the best-fit prediction. The Λ CDM model, combined with the full data model, provides an excellent fit to the measurements, with all residuals consistent with statistical expectations. The residuals further illustrate the transition between signal-dominated (correlated residuals) and noise-dominated (uncorrelated residuals) multipole ranges.

spectra calculated from different bundles (i, j) , defined in §III B, as

$$\bar{C}_\ell^{XY;\mu\nu} = \frac{1}{N_{\text{bundles}}(N_{\text{bundles}} - 1)} \sum_{i \neq j} C_\ell^{XY;\mu\nu;ij}. \quad (27)$$

Note that for the case where $\mu = \nu$, the normalization of the sum is still correct as the right-hand side sum double counts the spectra. We do not use auto-frequency spectra and cross-frequency spectra calculated from the same bundle to avoid noise bias and reduce co-temporal systematics, i.e. we never use $C_\ell^{XY;\mu\nu;ij}$ with $i = j$.

Due to our filtering strategy, we limit ourselves to multipoles above $\ell_{\min} \equiv 400$, where the transfer function is above 0.5 (see Fig. 5). This threshold is above the hard limit of our filtering at $\ell = 300$. This conservative choice avoids the multipoles where the filtering effect is large and where we would be limited by the MC precision of our transfer function estimate, see Appendix A.

Given that the main goal of this work is measuring the CMB primary anisotropies, we cut the temperature power spectrum at $\ell_{\max}^T \equiv 3000$ where the signal starts to be dominated by foreground contamination in all six cross-frequency spectra. The foreground contamination is much weaker in polarization, and we include multipoles up to $\ell_{\max}^E \equiv 4000$ in TE and EE.

We debias the band powers for all the effects included in Eq. (3), i.e. the multiplicative HEALPix pixel window function, transfer function, and temperature beam, and the additive filtering artifact and inpainting residuals. As described in §IV B, the quadrupolar beam leakage as well as the polarized beams are treated as parametrized systematics in the signal data model at the likelihood level. We bin the power spectrum estimates into $\Delta\ell = 50$ band powers using uniform weighting in D_ℓ with the binning operator

$$Q_{b\ell} = \begin{cases} 0 & \text{if } \ell \notin [50b, 50(b+1)], \\ \frac{\ell(\ell+1)}{\sum_{50b \leq \ell' < 50(b+1)} \ell'(\ell'+1)} & \text{otherwise.} \end{cases} \quad (28)$$

Binning reduces the impact of possible errors on the off-diagonal terms of the covariance matrix and reduces the residual mode mixing due to the Polspice regularization. Camphuis et al. [25] show that with our mask, the covariance precision is of order 1% at this binning size. Binning also reduces the size of the data vector, reducing the computational cost of likelihood estimation. Binning can potentially hide features in the power spectrum and reduce constraining power, but Fig. 1 of [60] shows that the binning choice made here leads to a negligible increase in cosmological parameter errors. This yields the final expression for the debiased binned band powers,

$$\hat{C}_b^{XY;\mu\nu} \equiv \sum_\ell Q_{b\ell} \frac{\bar{C}_\ell^{XY;\mu\nu} - A_\ell^{XY;\mu\nu} - I_\ell^{XY;\mu\nu}}{F_\ell^{XY;\mu\nu} B_\ell^{T;\mu} B_\ell^{T;\nu} P_\ell^2}, \quad (29)$$

where $\bar{C}_\ell^{XY;\mu\nu}$ is the average of the cross-power spectra calculated from different bundles from Eq. (27) and other terms have been defined above.

We also calculate $Q_{b\ell}^{XY;\mu\nu}$, the band power window functions that summarize all the effects of debiasing and binning and relate the theoretical sky signal power spectra to the data vector:

$$Q_{b\ell}^{XY;\mu\nu} = F_\ell^{XY;\mu\nu} B_\ell^{T;\mu} B_\ell^{T;\nu} P_\ell^2 \sum_{\ell'} \frac{Q_{b\ell'} K_{\ell'\ell}^{XY}}{F_{\ell'}^{XY;\mu\nu} B_{\ell'}^{T;\mu} B_{\ell'}^{T;\nu} P_{\ell'}^2}. \quad (30)$$

We use the band power window functions to build our likelihood in Section V and Eq. (40).

We show the final multi-frequency band powers in Fig. 9, along with the best-fit predictions assuming Λ CDM (see §VII B 2). Contributions from foregrounds are clearly seen at high ℓ in TT. At this stage, we can already appreciate the excellent agreement between the different cross-spectra. We discuss the internal consistency of the data in §VIB.

E. Noise power spectra

We use the noise realizations described in §III B to estimate noise power spectra. To estimate it, we take the average of the auto-power spectrum of sign-flip maps over the $N_{\text{signflip}} = 500$ realizations, masked using the combination of the point source and border masks. We show TT and EE noise power spectra in Fig. 10.

The temperature noise is dominated on large scales by contributions from the atmosphere, resulting in a characteristic red spectrum [61]. At high multipoles, white instrumental noise dominates. We report a white noise floor of 5.4, 4.4, and 16.2 μK –arcmin at 95, 150, and 220 GHz, respectively, and a coadded noise level of 3.3 μK –arcmin. As noted in [14], the atmospheric noise is significantly correlated in temperature between different frequencies. At high multipoles, we expect the cross-frequency noise to vanish since the white noise of the detectors and readout is uncorrelated, and we confirm that the cross-frequency noise vanishes at sufficiently high ℓ up to the uncertainty of our estimation. We include the cross-frequency noise in the covariance estimation. We recall here that the cross-frequency noise only affects co-temporal data, i.e., bundles at different frequencies that see the same atmosphere at the same time. The cross-bundle based auto-frequency or cross-frequency band powers are not noise-biased.

In polarization, the noise is dominated by white detector and readout noise and is hence nearly scale-independent. We report a white noise level of 8.4, 6.6, and 25.8 μK –arcmin at 95, 150, and 220 GHz, respectively, and a coadded noise level of 5.1 μK –arcmin. These values are larger than the expected noise from temperature, $n^P = \sqrt{2}n^T$, as they are affected by the correction for the polarization efficiencies. Atmospheric noise in polarization is subdominant on the angular scales probed by this analysis and hence cross-frequency noise is negligible (see [18] for detailed discussion of polarized

atmospheric noise). We find no evidence of significantly correlated noise between temperature and polarization.

F. Covariance matrix

In contrast with previous SPT-3G analyses which relied mostly on simulations for the power spectrum covariance estimation [12, 14], here we use a semi-analytic procedure akin to what was used in *Planck* analyses (e.g., [34]). We highlight in this section the different steps of the covariance estimation and where we differ from the *Planck* analysis.

Since the SPT-3G Main field covers 4% of the sky, an analytical computation of the covariance matrix requires proper modeling of the mask coupling effect. The problem is also present in the *Planck* analysis, but the near-unity sky fraction observed by *Planck* enabled the use of a simpler approximation. In this work, we compute the band power covariance matrix based on the framework developed in [25]. Following the notation in that work, we compute the covariance matrix of the biased pseudo-power spectrum $\tilde{C}_\ell^{XY;\mu\nu}$ (i.e. before the *Polspice* debiasing) with

$$\begin{aligned} \tilde{\Sigma}_{\ell\ell'}^{XY\mu\nu;ZV\alpha\beta} &= \Xi_{\ell\ell'}^{XY;ZV}[W^2] \sum_{L_1, L_2} \quad (31) \\ &\left[\bar{C}_{L_1}^{XZ\mu\alpha} \bar{\Theta}_{\ell\ell'}^{XZ;YV;L_1 L_2}[W] \bar{C}_{L_2}^{YV\nu\beta} + (Z \leftrightarrow V) \right], \end{aligned}$$

where $\Xi[W^2]$ and $\bar{\Theta}[W]$ are coupling matrices describing the mask effect and are entirely computed from the apodized border-only mask W described in §III C.

Following Eq. (3), the model signal is

$$\bar{C}_\ell^{XZ\mu\nu} \equiv F_\ell^{XY\mu\nu} B_\ell^\mu B_\ell^\nu P_\ell^2 \left(C_\ell^{XY\mu\nu;\text{signal}} + N_\ell^{XY\mu\nu} \right). \quad (32)$$

This model neglects the contribution of the inpainting and filtering artifact residuals to the covariance but does include the isotropic part of the filtering and $a_{\ell m}$ weighting, beams, and pixelization effects. At the end of the analysis, the signal part is evaluated at the best-fit cosmology and foreground model. For previous iterations of the analysis we used a best-fit *Planck* model, along with the foreground components used to generate simulated skies (§III F). The noise contribution is described in §IV E and includes the cross-frequency TT noise. The noise maps are found to be statistically isotropic, and, contrary to some previous CMB studies [34, 48], we do not include any anisotropic noise correction.

Similarly to the power spectrum, the covariance matrix is debiased from the mask effects. The debiasing is done using the G^{XY} kernel from [25] that includes the mask debiasing and *Polspice* residual kernel. We further correct for the instrumental beam, transfer function, and

factors of the pixel window function, such that

$$\hat{\Sigma}_{\ell\ell'}^{XY\mu\nu;ZV\alpha\beta} = \frac{\sum_{L_1 L_2} G_{\ell L_1}^{XY} \tilde{\Sigma}_{L_1 L_2}^{XY\mu\nu;ZV\alpha\beta} G_{L_2 \ell'}^{ZV}}{F_\ell^{XY\mu\nu} B_\ell^\mu B_\ell^\nu P_\ell^2 F_{\ell'}^{ZV\alpha\beta} B_{\ell'}^\alpha B_{\ell'}^\beta P_{\ell'}^2}. \quad (33)$$

The data vector is formed by the averaged cross spectra of the 30 bundles. Instead of summing the covariance of each of the cross spectra, we follow [44] by computing the covariance of the coadded map and correcting it following the methodology described in Appendix C 1, so that

$$\hat{\Sigma} \rightarrow \hat{\Sigma} + \frac{1}{N_{\text{bundles}} - 1} \hat{\Sigma}^{\text{noise-only}}. \quad (34)$$

We discussed in §IV A 2 that at the level of the covariance matrix, the anisotropic contribution from the filtering and weighting can be modeled by a rescaling of the covariance, $H_\ell^{XY;\mu\nu}$ estimated from simulations, so that

$$\hat{\Sigma}_{\ell\ell'}^{XY\mu\nu;ZV\alpha\beta} \rightarrow \frac{1}{\sqrt{H_\ell^{XY\mu\nu} H_{\ell'}^{ZV\alpha\beta}}} \hat{\Sigma}_{\ell\ell'}^{XY\mu\nu;ZV\alpha\beta}. \quad (35)$$

In order to avoid large-to-small scale couplings in the covariance, we have filled the map at the location of bright point sources and massive galaxy clusters with Gaussian constrained realizations of the CMB. Following the discussion in §IV C, we correct the covariance with

$$\hat{\Sigma}_{\ell\ell'}^{XY\mu\nu;ZV\alpha\beta} \rightarrow \hat{\Sigma}_{\ell\ell'}^{XY\mu\nu;ZV\alpha\beta} (1 + \rho_\ell^{XY;\mu\nu} \rho_{\ell'}^{ZV;\alpha\beta}) \quad (36)$$

where ρ_ℓ is the inpainting ratio defined in Eq. (26). We compare the impact of the inpainting correction on the covariance matrix to the effect of point source masking in Fig. 8, and we show that the increase in variance from inpainting is much smaller than what would have been incurred from masking.

At this stage, we do not include other data processing effects in the covariance. Beam uncertainties, quadrupolar beam leakages, and polarized beam corrections are explored along with the CMB and foreground signal in the likelihood and, in contrast to the approach in ACT DR6 [6], we do not marginalize over them in the covariance. We note that we marginalize over some foreground contributions in the covariance, as described in §V B 2, in order to speed up our likelihood evaluation.

We account for the lensing contribution to the covariance by adding to the final covariance the lensing checkerboard term described in [62], computed using *FisherLens*⁹ [63]. As in previous SPT analyses [12, 14] and, in contrast with [6], we do not include the contribution of super-sample lensing modes in the covariance, but marginalize over its effect in the likelihood. Also, since we limit the TT analysis to $\ell \leq 3000$, we ignore

⁹ <https://github.com/ctrendafilova/FisherLens>

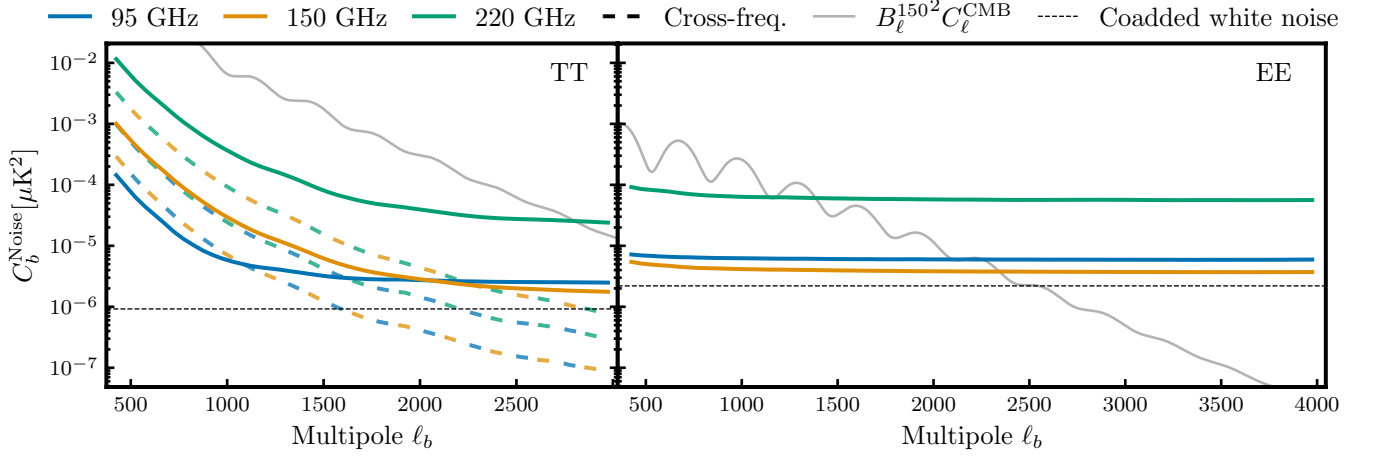


FIG. 10. Noise power spectra of the maps based on SPT-3G D1 observations. Solid lines correspond to auto-frequency spectra (95 GHz: blue, 150 GHz: orange, 220 GHz: green), dashed lines to cross-frequency spectra (alternating colors according to the frequency combination: 95 \times 150 GHz: blue-orange, 95 \times 220 GHz: blue-green, 150 \times 220 GHz: orange-green). The gray line shows the expected beam-convolved CMB signal for the Λ CDM best-fit of the SPT-3G D1 T&E data. The horizontal dashed lines indicate the coadded white noise levels in temperature (3.3 μ K-arcmin) and E-mode polarization (5.1 μ K-arcmin). We do not show the cross-spectra for EE channels, as they are consistent with zero. The noise spectra have been corrected for the transfer function (see §IV A 2) and the pixel window function. The white noise levels in temperature and E-mode polarization are 5.4, 4.4, and 16.2 μ K-arcmin and 8.4, 6.6, and 25.8 μ K-arcmin at 95, 150, and 220 GHz, respectively. The SPT-3G D1 EE data is signal-dominated up to $\ell \sim 2550$.

the non-Gaussian contributions to the covariance from foregrounds. Finally, the resulting matrix is binned as we did for the band powers with the $Q_{b\ell}$ matrix defined in Eq. (28)

$$\Sigma_{bb'}^{XYZV} = \sum_{\ell, \ell'} Q_{b\ell} \hat{\Sigma}_{\ell\ell'}^{XYZV} Q_{b'\ell'}. \quad (37)$$

The resulting covariance matrix is highly conditioned because of the low noise levels, the high resolution, and the large number of band powers. Slight inconsistencies between the different frequencies in the data vector that are not modeled in the approximated covariance, and can be much smaller than the cosmological constraining power, have a large effect on the goodness-of-fit of any cosmological model solution. While this is not an issue when performing band-power difference tests discussed in the validation section of this article (§VIB 1), we fail our more stringent conditional tests (§VIB 2) because the associated error bars are much smaller than the sample variance error bars. We acknowledge this limitation of our covariance approximation and regularize the matrix by minimally increasing its diagonal. This slightly decorrelates the different cross-frequency spectra. Specifically, we add a small fraction of the sample variance to the diagonal

$$\Sigma^{XYXY} \rightarrow \Sigma^{XYXY} + \alpha_{XY}^2 \text{diag}(\Sigma^{XYXY}) \quad (38)$$

where diag is an operator extracting only the diagonal of the matrix. The regularization factor is chosen to be $\alpha_{TT} = 0.1\%$, $\alpha_{TE} = \alpha_{EE} = 1\%$, so that the precision

of our tests (consistency tests or goodness-of-fit) cannot be sensitive to effects which are below 1% of the sample variance error bars. This factor is similar to that used in the map null tests, see Q25. We measured on **Fullmock** simulations that the regularization decreases the average χ^2 by 60 points. We checked that the χ^2 nevertheless follows a χ^2 distribution with the expected number of degrees of freedom, allowing us to easily interpret it. Finally, we verified that, as expected, the regularization step has no measurable impact on cosmological results.

We describe the validation of the covariance matrix in §VIC 1 and show that it is consistent with the **Fullmock** covariance matrix.

The resulting covariance matrix can be used to investigate the contribution of each of the frequency channels to the minimum-variance combination of the band powers, as discussed in Appendix C 2. We build the minimum-variance band powers using the covariance matrix, as described in Eq. (C2), and the associated minimum-variance covariance matrix. Applying the same steps to *Planck* and ACT DR6 data, we show the resulting signal-to-noise ratio in Fig. 11. We note that the bin width factor is divided out in this plot, so that the SNR for all experiments is shown with an effective bin width of $\Delta\ell = 1$. The SPT-3G D1 T&E data are the most sensitive in EE and TE at $\ell = 1800$ -4000 and $\ell = 2200$ -4000¹⁰, respectively. In the multipole range

¹⁰ At higher multipoles, the tightest constraints were published in [64].

$\ell = 1800$ -2200 in TE, the constraining power of SPT-3G is comparable to that of ACT DR6. The signal-to-noise per multipole is greater than one at $\ell \lesssim 3300$ in both TE and EE.

Uncertainties in the SPT-3G D1 T&E band powers are dominated over much of the ℓ range by the large sample variance resulting from the small sky fraction used in this analysis. For example, the signal contribution to the covariance is larger than the noise contribution up to $\ell \simeq 2600$ in EE, and over the full range reported here in TT. This makes the signal-to-noise per bin at large scales lower than that for *Planck* and ACT DR6, which used larger sky areas. This limitation will be largely overcome by the addition of the SPT-3G Summer and Wide fields, with observations obtained on a total of $\sim 10\,000\text{ deg}^2$, 25% of the sky. The forecasts for the full SPT-3G Ext-10k survey shown in Fig. 3, including the effect of filtering as described in Eq. (13), are shown as blue dotted lines in the plot.

V. LIKELIHOOD

The temperature and polarization band powers described in §IV D and shown in Fig. 1 and Fig. 9 are one of the two major results of this paper. The other is the constraints on cosmological parameters that these unprecedentedly sensitive power spectrum measurements enable. To go from one to the other, we need a framework in which to compare the measured band powers to cosmological models; we describe that framework in this section. The four primary components to this framework are: (1) the likelihood function quantifying the probability of obtaining the measured band powers given an underlying model; (2) the cosmological model and the procedure to produce model band powers from it; (3) the model of the instrument, data analysis, and foreground contamination that transforms the theory band powers into a data vector we can compare directly to the measured data; and (4) the method for exploring posterior distributions of the cosmological and data-model (nuisance) parameters. We discuss all of these in detail below.

A. Likelihood function

On the angular scales probed in this analysis, even given our limited sky coverage, each band power bin averages over a sufficiently large number of independent modes that the distribution of band powers can be well approximated as Gaussian [65]. Under this assumption, the formal likelihood \mathcal{L} , i.e. the probability of the data given the model, is

$$-\ln \mathcal{L}(\hat{C}|C^{\text{model}}(\theta)) \propto \frac{1}{2} \left[\hat{C}_b - C_b^{\text{model}}(\theta) \right] \Sigma_{bb'}^{-1} \left[\hat{C}_{b'} - C_{b'}^{\text{model}}(\theta) \right], \quad (39)$$

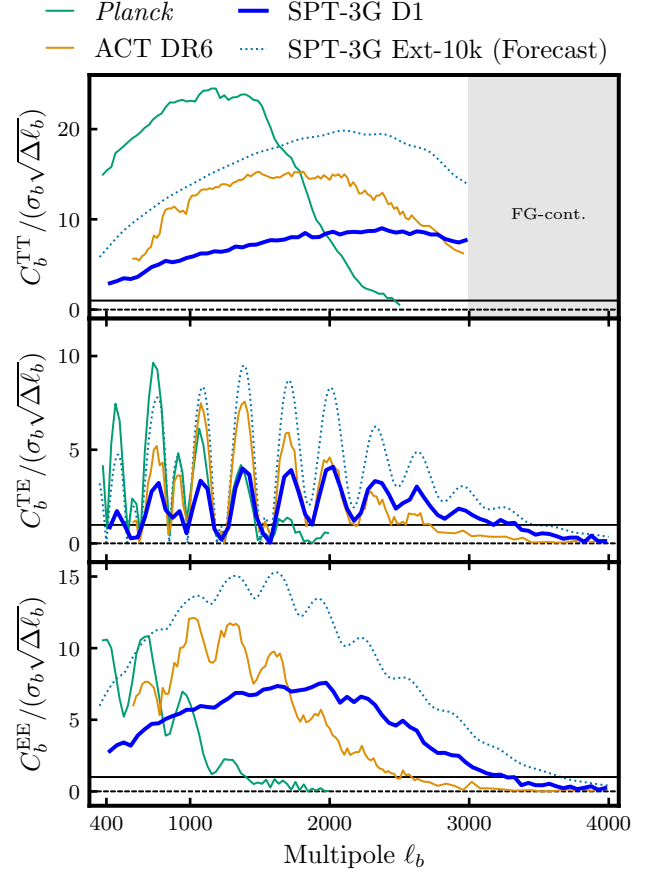


FIG. 11. Signal-to-noise ratio (SNR) of the minimum-variance band powers for SPT-3G D1 T&E (blue), *Planck* (green), and ACT DR6 (orange). At multipoles $\ell > 1800$, SPT-3G D1 T&E achieves the highest sensitivity in EE. For TE, the signal-to-noise ratio is similar to that of ACT DR6 between $\ell = 1800$ and $\ell = 2200$ and exceeds it at higher multipoles. In dotted blue, we show the forecasted signal-to-noise for the complete SPT-3G Ext-10k data, including all fields (see Fig. 3 and [41]). In that line, we included the effect of filtering assuming the simple model of the transfer function described in Fig. 5 and the relation of Eq. (13). We do not display the TT SNR above $\ell_{\text{max}}^T = 3000$, as the minimum-variance band powers become correlated due to foreground contamination and as these band powers are not measured in this work. The black dashed line indicates SNR=0, while the solid black line indicates SNR=1.

where \hat{C}_b represents the binned power spectrum estimates (band powers) after all debiasing steps, defined in Eq. (3) and Eq. (29); C_b^{model} are the model band powers, including CMB and foreground contributions, and transformed into a quantity comparable to the data vector as described below; Σ is the band power covariance matrix described in §IV F; and summation over repeated indices is implied. The full data vector \hat{C}_b consists of 21 auto- and cross-frequency TT, TE, and EE spectra (including TE and ET separately for cross-frequency pairs) for a total of 1392 elements.

The full set of model parameters θ consist of cosmo-

logical parameters ψ (six Λ CDM parameters plus any extension parameters), which are handed to Boltzmann solver routines or emulators to create theory CMB power spectra (see §V E), and 43 nuisance parameters ϕ , which are used along with the band power window functions (§IV D) to transform the theory CMB spectra into C_b^{model} , the model band powers appropriate for comparing to our data band powers. This transformation, which we refer to as the data model, includes both additive and linear corrections to the theory spectra, accounting for binning and residual mask effects, astrophysical foregrounds, and any residual systematic biases not fully removed in the processes described in Section IV. The data model is performed within the likelihood implementation described in §V C.

In the following sections, we discuss the data model

in detail, including implementation details and efforts to speed up the computation. We also introduce the different codes we relied on in this work to evaluate the CMB power spectra and briefly describe their relative merits.

B. Data model

The theory CMB spectra are computed externally to our likelihood implementation, while the transformation of these spectra into band powers that can be compared with the data (including the addition of foreground spectra) occurs within our likelihood code. With this in mind, we write the transformation, which we refer to as the data model for our likelihood, as follows:

$$C^{\text{model}}(C^{\text{CMB}}(\psi), \phi) = A_{\text{cal}} \cdot E_{\text{cal}} \cdot Q \cdot B^{\text{P}}(\beta_{\text{pol}}) \cdot B^{\Delta}(\beta_i) \cdot L(\epsilon_2) \cdot [A \cdot S(\kappa) \cdot C^{\text{CMB}}(\psi) + C^{\text{fg}}(\phi^{\text{fg}})], \quad (40)$$

where the nuisance parameters ϕ have been split into foreground parameters ϕ^{fg} and individual systematic parameters; double line symbols (such as Q) denote linear operators that can depend on nuisance parameters; and indices indicating spectrum types (i.e. TT, EE, TE), frequency pairs, or band power bins have been suppressed for clarity. The different ingredients of this long equation are, from left to right:

- A_{cal} : calibration factors; one external calibration $A_{\text{cal}}^{\text{ext}}$, and the relative calibration factors compared to the 150 GHz channel, $A_{\text{cal}}^{\text{rel};95}$ and $A_{\text{cal}}^{\text{rel};220}$
- E_{cal} : polarization efficiencies, affecting only the TE, ET, and EE spectra. They are similarly separated into an external parameter E_{cal} , and relative polarization efficiencies compared to the 150 GHz channel, $E_{\text{cal}}^{\text{rel};95}$ and $E_{\text{cal}}^{\text{rel};220}$.
- Q : the band power window functions defined in Eq. (30), accounting for the binning and residual Polspice mixing.
- B^{P} : the polarized beam correction (§IV B 3) defined in Eq. (18).
- B^{Δ} : the beam error modes operator which we use to propagate the error on the effective beam measurement (§IV B), defined in Eq. (43).
- L : the quadrupolar beam leakage (§IV B 2) defined in Eq. (14).
- A : the aberration defined in Eq. (42).
- $S(\kappa)$: the super-sample lensing correction as defined in Eq. (41).
- $C^{\text{fg}}(\phi^{\text{fg}})$: the additive foreground contamination model. It does not depend on the CMB signal,

and we describe its content in §V B 2.

We detail the different linear and additive corrections in the following two sections: §V B 1 for the linear operators and §V B 2 for the additive foreground model. Table IX and Table X list all of the nuisance parameters and, when relevant, the priors we use. We note that we do not include the uncertainty associated with filtering artifacts (see §IV A 1) and inpainting (see §IV C) in the likelihood. We have tested that the impact of these effects on the cosmological parameters we constrain is less than 0.1σ .

Finally, we note that after unblinding, the foreground model was updated and the T-to-P leakage corrections and polarized beam modeling beam were added. We only report the final model here and discuss the changes in detail in Appendix B.

1. Linear corrections

Most of the linear corrections arise from the non-idealities of the instrument and data processing and depend on parameters summarized in Table IX. Two of them, S and A , are of astrophysical origin.

The super-sample lensing S is the distortion of the CMB on our field from weak gravitational lensing due to modes larger than the survey area. Following [14] and [66], we treat the overall mean convergence in the survey field κ as a free parameter and include its effect on power spectra as:

$$[S(\kappa) \cdot C]_{\ell} \equiv C_{\ell} - \frac{\partial \ell^2 C_{\ell}}{\partial \ln \ell} \frac{\kappa}{\ell^2}. \quad (41)$$

We use a prior on κ of $\mathcal{N}(0, 0.00045)$ based on [12].¹¹

The aberration \mathbb{A} accounts for the distortion due to the motion of the Earth with respect to the rest frame of the CMB [67]. The SPT-3G Main field covers a small fraction of the sky and we do not explore the forward/backward symmetry of the local dipole, in which case a first-order approximation to the aberration is sufficient:

$$[\mathbb{A} \cdot C]_\ell \equiv C_\ell - \beta_{\text{ab}} \ell \frac{\partial C_\ell}{\partial \ell}. \quad (42)$$

The correction depends on the dipole and mean survey angle relative to the direction of the dipole, the product of which [14] computed for the SPT-3G Main survey as $\beta_{\text{ab}} = -0.0004826$.¹²

The group of nuisance operators, \mathbb{L} , \mathbb{B}^Δ , and \mathbb{B}^P all describe the limitations of our beam modeling and debiasing. The first of the operators, \mathbb{L} , encodes the quadrupolar temperature-to-polarization leakage described in §IV B 2 and Eq. (14). We describe in Appendix B 1 how we propagate the leakage measured on point source maps to the priors on the parameters of \mathbb{L} .

Like the analyses of [68–70], we treat uncertainty on the temperature beam at the parameter level (as opposed to including it in the power spectrum covariance, as in [6, 14, 34]). We first estimate a beam covariance matrix by varying the AGN and Saturn observations included in the stitched beam, as well as parameters that explore our analysis choices and instrument systematics (see H25 for details). We then extract from that matrix nine eigenmodes, discarding eigenvalues smaller than 1% of the largest eigenvalue, and include the amplitudes of these eigenmodes as free parameters in the likelihood, in order to fit for potential residual systematic errors and marginalize over the associated uncertainty. We introduce the beam error modes operator \mathbb{B}^Δ , which modifies power spectra as:

$$\mathbb{B}^\Delta \cdot C \equiv \left(1 + \sum_{i=1}^9 \beta_i B_i^{\Delta, \mu} \right) \left(1 + \sum_{i=1}^9 \beta_i B_i^{\Delta, \nu} \right) C^{\mu\nu}, \quad (43)$$

where B_i^Δ is the i th beam error mode and β_i is the amplitude of that mode. We include nine error modes and thus add nine free β_i parameters to the model. The diagonalized beam covariance matrix spans the full frequency and spectrum (TT, TE, EE) space, so that the B_i^Δ error modes span this full space and include any correlations of beam error between frequency pairs and spectra. The modes are normalized by their eigenvalues, so we use uncorrelated standard normal priors for each

of them. We show in Table IX that in our standard cosmological fits all β_i are compatible with zero. Because some of the constraints in the standard fits appear prior-dominated, we also checked that after widening the priors by a factor of 10 the constraints on all β_i remain compatible with zero.

Finally, the \mathbb{B}^P operator accounts for the difference between the beams in temperature and polarization. Using the notation in §IV B 3, this operator is defined through the relations:

$$\mathbb{B}^P \cdot C^{\text{EE}; \mu\nu} \equiv \frac{B^{\text{T}; \mu} B^{\text{T}; \nu}}{B^{\text{P}; \mu} \left(\beta_{\text{pol}}^\mu \right) B^{\text{P}; \nu} \left(\beta_{\text{pol}}^\nu \right)} C^{\text{EE}; \mu\nu}, \quad (44)$$

$$\mathbb{B}^P \cdot C^{\text{TE}; \mu\nu} \equiv \frac{B^{\text{T}; \mu}}{B^{\text{P}; \mu} \left(\beta_{\text{pol}}^\mu \right)} C^{\text{TE}; \mu\nu}. \quad (45)$$

As discussed in §IV B 3, when $\beta_{\text{pol}}^\mu = 1$ the polarized beam at frequency μ is identical to the temperature beam, and when $\beta_{\text{pol}}^\mu = 0$ the polarized beam at frequency μ is equal to the model of the main temperature beam only (no sidelobes). In the likelihood we use uniform priors between 0 and 1 on all β_{pol}^μ , and, as shown in Appendix D and Table IX, $\beta_{\text{pol}}^\mu < 1$ is clearly preferred by our data in all frequency bands.

The last two operators \mathbb{A}_{cal} and \mathbb{E}_{cal} recalibrate the absolute gain and polarization efficiency, respectively, and propagate our uncertainty on the calibration of our maps (§III D) to the constraints on cosmology. In most cases, we only use an informative prior from our external calibration of the 150 GHz map on the *Planck* map, $A_{\text{cal}}^{\text{ext}}$ (§III D), and let the internal calibration parameters vary in flat priors as reported in Table IX. The same is true for the internal polarization efficiency estimates. The external polarization efficiency is jointly fitted with all other cosmological plus nuisance parameters assuming a cosmological model [71]. We adopt a flat prior for the external polarization efficiency parameter, except when reporting results from TE or EE spectra only. In this case, certain cosmological parameters are fully degenerate with polarization efficiency and we break the degeneracy with our *Planck*-based prior.

2. Foreground model

As has been documented in many results over the last 25 years, signals from the CMB are contaminated by many different astrophysical foregrounds and this contamination must be taken into account in cosmological modeling. Which foregrounds are necessary to model depends on many factors, including which area of sky and multipole range are targeted, what frequency bands are used, and whether the temperature or polarization power spectra (or both) are used (see, e.g., [45] and [72] for reviews).

In this analysis, we model foregrounds in a manner similar to [14]. We include Galactic (Milky Way)

¹¹ Note that as defined above $\kappa > 0$ corresponds to a demagnification effect.

¹² We note that we only apply the aberration correction to the CMB, while [6] applies it to both CMB and foregrounds.

and extragalactic contributions in the model. For the Galactic model we include dust emission in both temperature and polarization, but we neglect synchrotron emission as it is expected to be negligible compared to Galactic dust at the frequency bands used in this work [73]. Galactic dust emission affects the TT, TE, and EE spectra. Our extragalactic foreground modeling includes synchrotron and quasi-thermal dust emission from background galaxies and the thermal and kinematic Sunyaev-Zel’dovich effects (tSZ and kSZ), all in temperature only. Using improved estimates of the polarized point source contribution from [64] results in a prediction of negligible contribution from this component.

The computation of each of the foreground components follows many published works and we defer the details of it to Appendix E. We show the total foreground contribution to the band powers in Fig. 37. We discuss here some specific choices made in the foreground modeling that are different than [14]. The priors we place on foreground parameters, as well as the posterior constraints on those parameters in the context of the Λ CDM model, are listed in Table X.

We adopt the same Galactic dust model as in [14], featuring a modified black-body spectral dependence and a power law spatial dependence. The SPT-3G Main field was designed to be far from the Galactic plane and the contribution from Galactic dust is predicted to be small compared to the CMB anisotropy in both temperature and polarization. The priors on Galactic dust in Table X were derived from power spectra of *Planck* maps on the SPT-3G Main field, with a color correction to account for the SPT-3G bandpasses.

We model the contribution in temperature from Poisson-distributed unresolved radio sources and dusty star-forming galaxies with a fixed Poisson power law ($D_\ell \propto \ell^2$) and a free amplitude parameter for each cross-frequency combination. We set uninformative uniform priors $(0, 200) [\mu\text{K}^2]$ on the amplitude parameters, as they should be well determined by the data.

We model the clustering term of the dusty star-forming galaxy distribution (hereafter “CIB clustering”) with an angular dependence parametrized by a power law index (similar to, e.g., [74, 75]). We use the results of [76] to impose a prior on the power law index, though we widened the prior to $\alpha^{\text{CIB}} \sim \mathcal{N}(0.53 \pm 0.1)$. We do not attempt to model the spectral energy distribution (SED) of this term and instead adopt individual uncorrelated amplitudes for the 150×150 , 150×220 , and 220×220 GHz cross-spectra. We do not detect any contamination from CIB clustering in cross-spectra involving 95 GHz data but we nevertheless add a suitable constant contribution to the band power covariance matrix by propagating the prior from the 2018 analysis to the updated data model, see Table X.

Finally, we model tSZ and kSZ signals using templates

derived from the AGORA [46] simulations.¹³ We allow the amplitudes of these templates to freely float in the likelihood, with priors adapted from the 2018 analysis and based on [77]. The tSZ amplitude parameter is defined at the *Planck* reference frequency of 143 GHz and the SED of this term is the standard tSZ frequency dependence relative to primary CMB fluctuations (see for example [78]), taking into account color corrections due to our bandpasses and ignoring any relativistic corrections. We marginalize over the tSZ-CIB cross-correlation by adding a constant contribution to the band power covariance matrix computed from the prior on tSZ-CIB cross-correlation used in [14] (see Appendix E and Table X). The kSZ SED is constant in CMB units.

C. Building a robust likelihood with `candl`

We implement the likelihood code in `candl` [60]¹⁴, a python-based CMB likelihood library with JAX support. Crucially, JAX exposes the code to an automatic differentiation algorithm, which allows for the easy and fast computation of accurate gradients. By combining `candl` with a differentiable theory code, we can then build a fully differentiable pipeline from cosmological and nuisance parameters, θ , through to the likelihood value, \mathcal{L} . This allows us to trivially evaluate the functions $\mathcal{L}(\theta)$, $\partial\mathcal{L}/\partial\theta|_\theta$, and $\partial^2\mathcal{L}/\partial\theta_i\partial\theta_j|_\theta$, which opens up a plethora of applications [see e.g. 60, 79].

In this analysis, we couple our `candl` likelihood to a differentiable model, *CosmoPower* (described in §V E), to greatly increase our ability to test the robustness of our analysis pipeline. Specifically, this unlocks two key tests:

1. Shortcutting MCMC analyses. We perform a gradient-descent minimization using the truncated Newton-Raphson algorithm implemented in `scipy` [80–82] and then approximate the parameter posterior distributions as Gaussian by evaluating the Hessian at the best-fit point to obtain the Fisher matrix [83].
2. Translating biases in band powers to biases in cosmological parameters. By performing a Taylor expansion of the likelihood around the best-fit point one can show that, to first order, parameter biases $\delta\theta$ from band power biases δD_ℓ are given by $\delta\theta = F^{-1} \frac{\partial D}{\partial \theta} \Sigma^{-1} \delta D_\ell$ [84]¹⁵, where F is the Fisher matrix and Σ is the band power covariance matrix (§IV F).

Both types of analyses can be performed in less than a minute, even for our high-dimensional multi-frequency

¹³ The templates are derived from the version of AGORA tSZ/kSZ maps with AGN heating temperature 10^8 K. When deriving these templates, a mask is applied with the same emissive source flux cut of 6 mJy at 150 GHz as used in this analysis.

¹⁴ <https://github.com/Lbalkenhol/candl>

¹⁵ We find higher-order contributions to be negligible.

likelihood. This allowed us to propagate any change to the likelihood—whether to the data vector (e.g., the residual bias from the inpainting procedure), the band power covariance matrix (e.g., the lensing checkerboard), or the data model (e.g., the instrumental beam models)—to parameter constraints with negligible computational cost. Still, since the above methods rely on certain approximations of the likelihood, they were used for testing purposes only; the final cosmological results presented in Section VII are calculated via traditional MCMC analyses.

D. CMB-only likelihood

We follow the procedure of [31] to construct a CMB-only, *lite*, likelihood. The underlying framework was first introduced by [75]. In this procedure we extract the best-fit CMB band powers and covariance from the combination of cross-frequency spectra of the same TT, TE, or EE channel, while marginalizing over nuisance parameters. These products are then used to construct a simple Gaussian likelihood following the functional form of Eq. (39) that can be used to explore cosmological models. The advantage of this approach is that it provides a fast and interpretable cosmological likelihood that compresses the information from different frequencies and marginalizes over foreground contamination and systematic effects. For details on the framework we refer the reader to [75] as well as to previous applications of this framework [6, 31, 85–88].

Using the new approach put forward by [31], we exploit the differentiability of our multi-frequency likelihood implementation to perform the data compression quickly and accurately. As in [31], we retain a global temperature and a global polarization calibration parameter (A_{cal} , E_{cal}) to minimize bin-to-bin correlations of the CMB-only band powers and account for the effect of aberration in the *lite* likelihood. The compression reduces the length of the data vector from 1392 to 196 and the number of operations in the data model from 18 to four (aberration, calibration of TT, TE, and EE spectra). This leads to a speed-up of a factor of 50 in the evaluation of the likelihood. Together with the reduction of the number of nuisance parameters from 43 to two, this greatly speeds up MCMC analyses. We refer to our compressed likelihood as SPT-*lite*.

Results from SPT-*lite* and the multi-frequency likelihood are consistent; in Λ CDM, the central values of cosmological parameter posteriors shift by $\lesssim 0.1\sigma$ and their widths match to within 10%; this is compatible with MC noise. The *lite* likelihood is made publicly available alongside the multi-frequency likelihood on the SPT website.¹⁶ Note that the compressed likelihood has

been constructed for the complete T&E data set; for constraints from individual spectra the multifrequency likelihood should be used. We provide further details on the construction of the *lite* likelihood and its performance in Appendix F.

E. Theory codes

A key component of the likelihood is the step of computing predictions of the CMB power spectra for a given set of cosmological parameters. The baseline cosmological model we use is the standard flat cold dark matter model with a constant dark energy component, Λ CDM. The six parameters we use to parameterize Λ CDM are: the physical density of baryons and dark matter, $\Omega_b h^2$ and $\Omega_c h^2$, the amplitude and spectral tilt of initial scalar perturbations, A_s and n_s , the optical depth to reionization τ_{reio} , and either the angular size of the sound horizon at recombination θ_s^* or the expansion speed of the universe today H_0 . In Section VII, we also explore several extended models and we define the extension parameters in that section. We assume one massive neutrino with $m_\nu = 0.06$ eV, unless otherwise stated. Definitions of all cosmological parameters appearing in this manuscript are provided in Appendix J and Table XI.

We generally use the Boltzmann solvers **CAMB** [89] and **CLASS** [90] to compute the CMB power spectrum expected for a given cosmological model. We use the same accuracy settings as the ACT DR6 results [6]. While these codes are accurate, they are slow; for some purposes, we use two different interpolation methods to replace the full Boltzmann computations by quicker, approximated ones: **CosmoPower** and **OLÉ**. We briefly introduce these two codes below and specify when we use them.

We use the models for the neural-network based **CosmoPower** emulator [29, 30, 91] developed in [14] at the following times: during the development of the likelihood (to enable the methods described in §VC), in the test of the CMB-only likelihood in Λ CDM (§VD and Appendix F), for the results presented in §VID, and when testing for scale coherence in §VIIB1. The emulator is trained on high-precision **CAMB** spectra and covers Λ CDM, as well as A_{lens} and N_{eff} extensions.¹⁷ The use of **CosmoPower** emulators over full Boltzmann solvers has two key advantages:

1. once trained, the emulator runs 10^5 times faster than **CAMB** or **CLASS** [29], and

¹⁶ <https://pole.uchicago.edu/public/data/camphuis25/>
https://github.com/SouthPoleTelescope/spt_candl_data

¹⁷ The metadata for the particular **CosmoPower** emulators we used is based on a training set with a higher accuracy than the one released with [29]. They are publicly available at https://github.com/alessiospurinomancini/cosmopower/tree/main/cosmopower/trained_models/SPT_high_accuracy

2. it is differentiable, which facilitates several robustness tests (as discussed in §VC).

However, our use of **CosmoPower** is limited by the availability of models that have been trained on high-precision reference spectra and cover the necessary multipole ranges for primary CMB and CMB lensing power spectra. One downside of the emulator compared to full Boltzmann solvers is also it does not return the full set of derived parameters of interest. Still, the list of publicly available models grows and derived parameters can also be emulated or may be calculated cheaply during post-processing using the emulator itself [91, 92].

It is difficult and expensive to train **CosmoPower** emulators on all possible cosmological extensions. We thus use a second code, **OLÉ**, to compute the CMB power spectra for cosmological models not covered by **CosmoPower**. **OLÉ** is an online learning emulator framework; integrated within the MCMC exploration codes **Cobaya** and **MontePython**, it trains a Gaussian-process emulator model while exploring the posterior distribution of the parameters [93]. After a short training period, the emulator is used during the MCMC and regularly validated and retrained to improve its accuracy and the efficiency of the parameter exploration. Inference with **OLÉ** is more flexible than **CosmoPower**, as the training takes place on the fly, exclusively where the posterior distribution has most of its mass. However, this also limits the speed gains compared to **CosmoPower** as the **OLÉ** emulator covers a smaller area of parameter space, focusing only on the region where high likelihood values are possible, relying on slower computation for the low probability region. During this work we encountered compatibility issues with the use of MUSE $\phi\phi$ and ACT DR6 $\phi\phi$ lensing likelihoods (see data set definitions in Table III) in **OLÉ** which limited our ability to exploit this promising tool at the time. Still, we were able to use **OLÉ** throughout Section VII for analyses without lensing data and to quickly obtain good proposal matrices for the other cases. We refer the reader to [93] for details on **OLÉ**.

Finally, when using **CLASS**, we use **HaloFit** [94, 95] to compute dark-matter only non-linear corrections, while we use **HMcode2020** [96, 97] when using **CAMB**. To assess the impact on our results, we first compared the primary and lensing CMB spectra for most of the models considered in §VII A using **HaloFit** or **HMcode2020**. Note that the current version of **HaloFit** included in **CLASS** incorporates corrections due to massive neutrinos or the CPL parametrization of dark energy [98, 99]. We found the relative differences between the two to be well below experimental uncertainties at all relevant scales. Moreover, as an additional test, we did runs using **HMcode2020** for Σm_ν with data sets SPT+ACT+WMAP+DESI [100] and N_{eff} with data sets SPT+ACT + DESI and compared them to equivalent runs with **HaloFit**. We find no deviations in parameter constraints beyond the 0.2σ level between the two codes. However, we should caution that true non-

linear corrections, especially for non-classical extensions of Λ CDM, might require performing N-body simulations, which is beyond the scope of this work (see [101] for the impact of non-linear corrections on cosmological analysis).

VI. VALIDATION OF THE ANALYSIS PIPELINE

In this section, we discuss the process for verifying that the model formulated in Section V accurately describes the measured data from Section IV.

A. Blinding

Building on [14], we constructed and validated our analysis pipeline “blind,” i.e. restricting ourselves from looking at certain results until a series of robustness tests have been passed. This methodology was designed to mitigate confirmation bias and, more concretely, to avoid stopping investigations early once results align with expectations or unconsciously modifying the analysis to achieve desired results.

During the blind stage of the analysis, we did not allow ourselves to compare the measured band powers to those from any other experiment, including previous SPT analyses¹⁸ or theoretical predictions. When deriving cosmological parameters for internal consistency checks or assessing the impact of analysis choices, the mean values were either hidden or systematically offset by an unknown amount. Finally, we did not perform any comparison that depends on the cosmological model, such as the cosmological parameter consistency test between TT, TE, and EE data, during the blind period.

These restrictions were lifted once the pre-defined tests detailed below were successfully passed. We committed to publishing the obtained results and documenting any post-unblinding modifications to the pipeline. For each test, a pre-defined passing criterion was established based on an associated probability-to-exceed (PTE), see Appendix G 1. The PTE threshold was set to $0.05/N$, where N represents the total number of independent tests performed and incorporates the correction for the look-elsewhere effect [102].

We performed the following consistency tests before unblinding: (1) null tests at the map level (see §III E and Q25), (2) differences between frequencies at the power spectrum level (§VI B 1), (3) conditional frequency

¹⁸ While the absolute calibration in temperature and polarization is obtained by comparing to *Planck* data (see §III D), these comparisons are performed only over the SPT-3G Main field mask; at no point before unblinding did we compare SPT-3G band powers to the full-sky *Planck* power spectra.

tests at the power spectrum level¹⁹ (§VIB 2), and (4) differences of Λ CDM parameters obtained from different frequencies (§VID). Finally, we ensured that our pipeline is unbiased and robust to differences in cosmological parameters by performing all our consistency tests on mock observations, both for our fiducial cosmology and alternate models.

While this analysis was still in the blind stage, the MUSE pipeline underlying [16] was mature enough to unblind those results. The MUSE analysis uses exclusively polarization data to estimate the unlensed EE power spectrum and lensing power spectrum $\phi\phi$, while the pipeline presented here also includes temperature information and estimates lensed T&E power spectra. For these reasons, we do not expect identical results from the two analyses. Initial results on Λ CDM parameters from MUSE were communicated across the collaboration on April 1st, 2024, after which key people working on this analysis were shut off from any further discussion of MUSE results. None of the MUSE findings and comparisons conducted after unblinding and changes to the MUSE pipeline were communicated to the analysis group for this work until they were ready to unblind as well.

We unblinded on September 9, 2024, after passing all pre-defined tests. Subsequently, we identified two previously unaccounted-for systematic effects in our data—quadrupolar temperature-to-polarization leakage (§IV B 2) and depolarization of beam sidelobes (§IV B 3)—and we made several other minor updates. The quadrupolar leakage is seen most clearly in stacked maps of bright point sources, while the beam sidelobe depolarization manifests in part as a subtle inconsistency between the EE band powers from different frequencies, just below our blinding threshold. Although our data passed the pre-defined consistency tests, these systematics resulted in failures of the consistency tests in the $\ell \in [3000, 4000]$ range for TE and EE. These multipole ranges had initially been excluded from the analysis and were only reintroduced after unblinding. While these corrections were made post-unblinding, and discovered at least in part through inconsistencies in cosmological fits, we emphasize that there is strong cosmology-independent evidence for them, as detailed in Appendix B. We are confident that the final pipeline provides a more accurate description of the data. Prior to these corrections, there were significant inconsistencies among the Λ CDM cosmological results derived from the individual TT,

TE, and EE spectra, resulting in unreliable combined T&E constraints and a poor overall fit, with the χ^2 statistic exceeding expectations. Consistent with our blinding protocol, we did not perform any tests strongly dependent on the cosmological model—such as cross-checks between TT, TE, and EE—before unblinding. Addressing these systematics substantially improved the agreement between temperature and polarization data within Λ CDM, as reflected in the improved fit quality. We compare cosmological constraints before and after the post-unblinding improvements in Fig. 35.

In the rest of this section we describe our suite of robustness tests, performed with the baseline likelihood, which includes post-unblinding corrections.

B. Power-spectrum level tests

We perform two types of consistency tests at the power spectrum level: (1) we assess the difference between power spectra estimated using different frequency pairs (hereafter “frequency spectra”) and (2) we compare a given frequency spectrum to the prediction for that spectrum conditioned on all other spectra. These tests have similarly been used in *Planck* [34] and SPT-3G 2018 analyses [14]. As discussed in §IV F, these tests are highly sensitive, with uncertainty margins significantly smaller than the band power error bars across most of the multipole range considered. This sensitivity is achieved by canceling common fluctuations, notably the sample variance, between spectra and enables stringent assessments of frequency-dependent systematic effects, such as residual foreground contamination, inter-frequency calibration, and beam effects. Notably, these tests are insensitive to systematic effects which impact all observation bands identically. While the covariance regularization described in §IV F make our tests insensitive to inconsistencies that are below 1% of the level of sample variance uncertainty, possible features of this size are irrelevant for our cosmological analysis.

As an example, in Fig. 12 we compare the error bars for the 150×150 GHz band powers with those associated with the difference test ($150 \times 150 - 150 \times 95$ GHz) and the conditional prediction test for the 150×150 GHz spectrum. Uncertainties are shown both with (‘regul’) and without (‘noregul’) covariance regularization. Notably, at large angular scales the conditional error bar with covariance regularization is two orders of magnitude smaller than the band-power error bars. Without covariance regularization, this reduction would reach three orders of magnitude for TT at large scales, highlighting the exceptional sensitivity of these tests. While the conditional test is more stringent than the difference test, the latter is easier to interpret, which is why we perform both. Both tests operate on CMB-only spectra and we remove the best-fit Λ CDM foreground and systematics contamination from our measured data. As such, there is a small model

¹⁹ Note that these two tests are slightly cosmology-dependent, since they require subtracting or correcting for any frequency-dependent contribution from each spectrum in order for them to be compared. From a practical point of view, this entails calculating a best-fit model assuming Λ CDM and using the inferred nuisance (foreground plus instrumental) parameters to correct the spectra before comparison. The relatively limited correlations between nuisance and cosmological parameters allow us to proceed in this way.

dependence to these tests.

1. Frequency difference

The first test we perform is the difference between pairs of frequency spectra. We correct each frequency spectrum for the contribution of foregrounds and systematic effects as described above, such that the spectra contain only the common CMB signal. We then calculate the difference of a pair of spectra as:

$$\Delta = \hat{C}^{\mu\nu} - \hat{C}^{\alpha\beta} \quad (46)$$

and the covariance of the difference Σ^Δ as:

$$\Sigma^\Delta = \Sigma^{\mu\nu;\mu\nu} + \Sigma^{\alpha\beta;\alpha\beta} - \Sigma^{\mu\nu;\alpha\beta} - \Sigma^{\alpha\beta;\mu\nu} \quad (47)$$

where $\Sigma^{\mu\nu;\mu\nu}$ and $\Sigma^{\alpha\beta;\alpha\beta}$ are the blocks of the covariance matrix for the two power spectra and $\Sigma^{\mu\nu;\alpha\beta}$ is the cross-covariance between the two. For each test, the number of degrees of freedom is equal to the number of bins in the difference spectrum, i.e. 52 for TT and 72 for TE and EE, as we cut TT data at $\ell_{\text{max}}^T = 3000$, see §IV D. In total, there are five independent tests for TT and EE and eight for TE spectra, such that the corresponding PTE threshold to pass is $0.05/18 = 0.0028$.

We list the PTEs of all tests and show the difference spectra in Appendix G 2 and Figs. 39 to 41. All tests pass, with the smallest PTE being 0.067, and no striking features that may indicate significant residual foreground contamination, calibration offsets, or beam mismodeling. In fact, all of the difference tests also pass even when not applying the covariance regularization (see Eq. (38)), signaling good consistency across frequencies.

2. Conditional spectra

The second way to test inter-frequency consistency is to calculate the difference between a frequency spectrum and its conditional prediction obtained from all the other frequency spectra [14, 34, 86]. The conditional prediction is obtained by decomposing, for each frequency combination $\mu\nu$, the data vector in two blocks

$$\hat{C} = \begin{bmatrix} \hat{C}^{\mu\nu} & \hat{C}^{\overline{\mu\nu}} \end{bmatrix} \quad (48)$$

where $\hat{C}^{\mu\nu}$ is the frequency spectrum being considered and $\hat{C}^{\overline{\mu\nu}}$ is the vector containing all the other frequency spectra. We similarly decompose the model vector, C , and the covariance matrix, Σ . For a Gaussian likelihood, the conditional prediction is

$$C^{\mu\nu|\text{cond}} = C^{\mu\nu} + \Sigma^{\mu\nu;\overline{\mu\nu}} \Sigma^{\overline{\mu\nu};\overline{\mu\nu}}^{-1} (\hat{C}^{\overline{\mu\nu}} - C^{\overline{\mu\nu}}), \quad (49)$$

with an associated covariance of

$$\Sigma^{\mu\nu;\mu\nu|\text{cond}} = \Sigma^{\mu\nu;\mu\nu} - \Sigma^{\mu\nu;\overline{\mu\nu}} (\Sigma^{\overline{\mu\nu};\overline{\mu\nu}})^{-1} \Sigma^{\overline{\mu\nu};\mu\nu}. \quad (50)$$

We then take the difference between the measured data and the conditional prediction, i.e. $\Delta = \hat{C}^{\mu\nu} - C^{\mu\nu|\text{cond}}$ and calculate a χ^2 test statistic using the covariance above and the relevant block of the band power covariance matrix. Similar to before, for each test the number of degrees of freedom is equal to the number of bins in the conditional spectrum and the total number of independent tests is the same as for the difference spectrum test.

We show the conditional test residuals and report their associated PTEs in Fig. 42. All PTEs lie above the predefined threshold. Remembering the small error budget of this test shown in Fig. 12, this is a strong sign for internal consistency. Given that this test, as well as the difference test, passes, we conclude that the differences between frequencies in our data are well-described by our data model and the band power covariance matrix.

C. Testing the pipeline on simulations

We validate the analysis pipeline on **Fullmock** simulations, treating them in the same manner as we do for the data. This validation is done to ensure that the covariance matrix is accurate and to confirm that the pipeline is unbiased and insensitive to the fiducial cosmology.

1. Covariance matrix validation

We begin by validating the analytic covariance matrix described in §IV F. We do so individually for the three constituents of the covariance—the sample variance, the chance correlation, and the noise variance terms—as well as for the complete matrix. First, we validate the sample variance term of our analytical covariance against 500 **Fullmock** simulations. When not masking point-sources in the simulations, we find that the covariance of the signal-only power spectra matches our analytical estimate. When point sources are masked, we recover the expected deviations discussed in §IV C.

Second, we validate the noise variance term. We feed only the noise term of Eq. (32) into Eq. (31) and compare this with the variance of the noise power spectra measured on the 500 noise realizations. We report agreement between the two estimates of the noise variance term. This test validates not only our covariance matrix framework, but also the computation of the four-point transfer function introduced in Eq. (13), as we compare our purely analytical prediction to an entirely data-based estimator.

Third, we validate the chance correlation term. We use our data bundle cross-spectra in the estimator presented in [44] to obtain the combined chance correlation and noise variance contribution to the covariance. We compare the resulting matrix to our analytical covariance after subtracting the sample variance contribution from

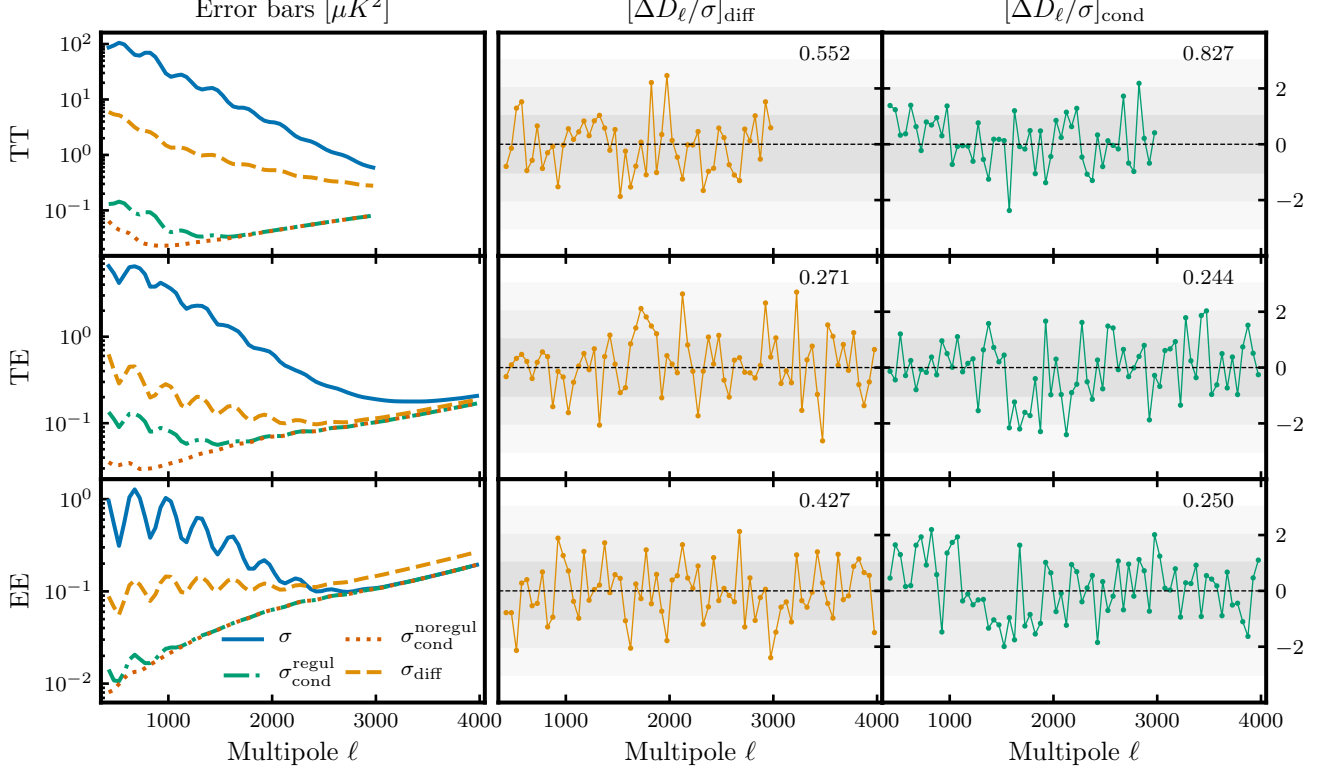


FIG. 12. Example of power spectrum consistency tests. The left column shows the error bars for the two 150×150 GHz band-power tests (difference test: dashed orange, conditional test: dash-dotted green) and compares them to the uncertainty of the power spectrum measurement (solid blue). We also show the error bars of the conditional test without covariance regularization, labeled ‘noregul’, for reference (dotted red). The central column displays the difference test for $(150 \times 150 - 150 \times 95)$ GHz, while the right column shows the conditional prediction test for the 150×150 GHz spectrum, both in relative units, i.e. deviations with respect to the expected uncertainty from the first column. PTEs are indicated in the top right corner of each panel. All tests across all spectra and frequencies pass, indicating the exceptional consistency of the measured data. The rest of the tests are shown in Appendix G Figs. 39 to 42.

the latter. We report good agreement given the expected precision for the number of available bundles.

Finally, to obtain an estimate of the complete covariance matrix, we use end-to-end simulations, created by combining simulated, inpainted **Fullmock** maps with noise realizations. From 100 effectively independent realizations constructed this way, we calculate the variance of the power spectra of these maps. The diagonal of the resulting matrix agrees well with our analytical calculation, given the number of realizations, as shown in Fig. 13. There is no apparent trend with multipole number; this is worth noting as the different constituents of the covariance dominate at different angular scales. This agreement indicates that our analytic framework provides a good description of the true covariance of our data vector. Because our analysis uses a curved-sky pipeline, we expect off-diagonal terms arising from sky masking to be negligible, as discussed in [25]. We have verified this expectation for the pure signal component using the full set of 500 **Fullmock** simulations. Off-

diagonal contributions from lensing and marginalized foreground components are included in our analysis. We conclude that the analytical framework presented here accurately describes the covariance of the data, in the case of the full matrix as well as for the sample variance, chance correlation, and noise variance terms individually.

2. Validation of the full pipeline

We run our full analysis pipeline on **Fullmock** simulations to check that our methodology is unbiased. This operation also validates that the scatter is accurately modeled by the covariance matrix. First, we take simulated signal-only maps, add sign-flip noise maps, and inpaint at the location of masked sources as we do for the data. Second, we run the band power estimation pipeline on these combined maps using the same procedure as for the data. Third, we use our cosmological likelihood from Section V to obtain Λ CDM parameter constraints

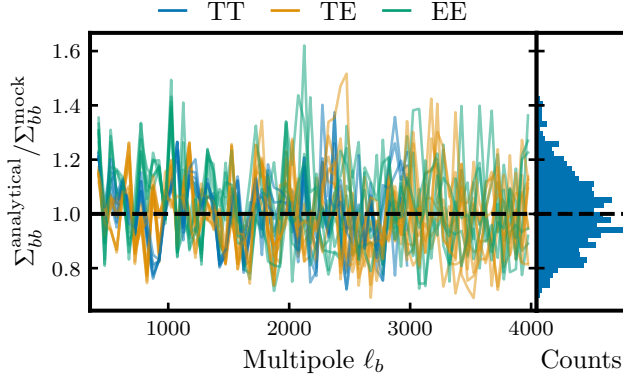


FIG. 13. *Left*: Ratio of the analytical covariance matrix diagonal to the covariance matrix diagonal based on end-to-end simulations for all frequency combinations. *Right*: Histogram of values of the ratio. We find excellent agreement between the two estimates of the covariance matrix.

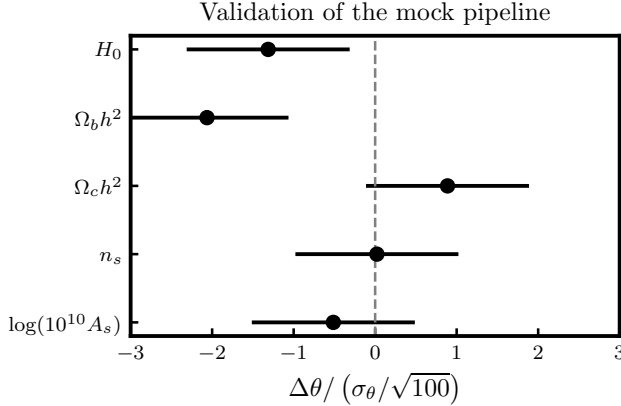


FIG. 14. Validation of the pipeline on **Fullmock** simulations. We show the offset between the input cosmological parameter values and the product of the posteriors obtained by analyzing the 100 **Fullmock** simulations in units of the expected uncertainty $\sigma = \sigma_\theta / \sqrt{100}$, where σ_θ is the expected data uncertainty on parameter θ . We report a PTE of 0.25 across the full parameter space.

for each set of band powers from step two. For those runs, we set polarized beam, quadrupolar beam leakage, and dust-related parameters to their default values (listed in Tables IX and X), since the simulations do not include these effects. We perform these steps for a total of 100 realizations and compare the parameter constraints from simulations to the input values. We find that the mean of the posterior distributions is consistent with the input values and that the scatter of the posteriors is well described by the expected covariance matrix. We show the results of this test, for which the PTE is 0.25, in Fig. 14.

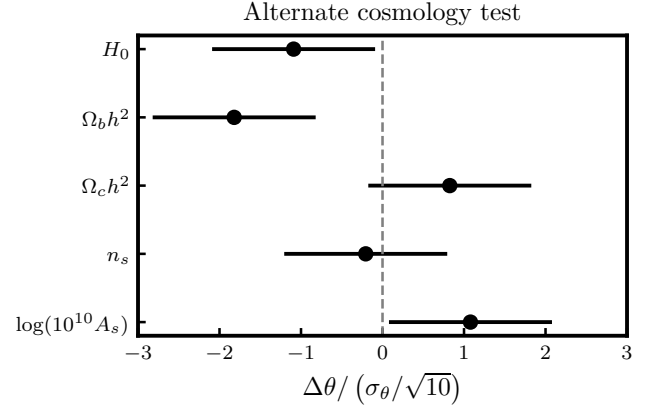


FIG. 15. Validation of the pipeline on 10 **Fullmock** simulations with an alternate cosmology. We show the offset between the input cosmological parameter values and the product of the posteriors obtained by analyzing the ten alternate cosmology simulations in units of the expected uncertainty $\sigma = \sigma_\theta / \sqrt{10}$, where σ_θ is the expected data uncertainty on parameter θ . All individual shifts are below 2σ and we report a PTE of 0.30 across the full parameter space, demonstrating that our analysis pipeline is largely insensitive to the chosen fiducial cosmological model.

3. Alternate cosmology test

The simulation pipeline, used to evaluate the transfer function and the additive biases and to validate the covariance matrix, is based on a fiducial cosmological model. To verify the robustness of our results with respect to this choice, we perform a test on ten mock observations based on a different cosmology. We do so by using the same simulation pipeline, but changing the input cosmological parameters, setting H_0 to a value consistent with the SH0ES measurement in [103], and changing other parameters to yield a 5σ discrepancy across the five constrained Λ CDM parameters compared to the fiducial values. We then run our regular pipeline on these alternate cosmology simulations and check that we recover the correct input cosmology within the expected uncertainties.

We show the averaged parameter constraints from the analyzed alternate cosmology simulations in Fig. 15. The shift we observe is consistent with zero at 1.0σ over the five dimensional parameter space and the overall PTE of this test is 0.30. We conclude that our analysis pipeline is largely insensitive to the fiducial cosmological model.

D. Parameter-level tests

We now assess the consistency of the Λ CDM cosmological parameters obtained from the spectra of data at different frequencies. This test is explicitly performed in Λ CDM and therefore has some model dependence.

However, this is limited to the assumption that Λ CDM provides an acceptable description of the common signal across frequencies. This allows us to further probe any potential inconsistencies across observational frequencies and short-comings of our data model. Since we are explicitly only interested in the differences between cosmological parameters, and not their absolute values, we were able to perform this test during the blind stage of the analysis. We emphasize that our consistency checks were performed only within the T&E constraints or within the individual TT, TE, and EE constraints, but not across these channels. Cross-channel consistency tests inherently depend on the assumed cosmological model and were therefore not conducted during the blind analysis stage.

We predict the expected correlation between parameters obtained from different frequency spectra using the framework developed in [84]. The covariance of the parameters derived from different spectra is given by

$$\text{cov}(\bar{\theta}^{\mu\nu}, \bar{\theta}^{\alpha\beta}) = (M^{\mu\nu})^T \Sigma^{\mu\nu;\alpha\beta} M^{\alpha\beta}, \quad (51)$$

where $M^{\mu\nu} \equiv (\Sigma^{\mu\nu;\mu\nu})^{-1} \frac{\partial D_{\ell}^{\mu\nu}}{\partial \theta} (F^{\mu\nu;\mu\nu})^{-1}$,

where F is the Fisher matrix [83]. Our differentiable likelihood gives us easy access to the derivatives $\partial D_{\ell}^{\mu\nu} / \partial \theta$ (as demonstrated by [60]), which we evaluate at the best-fit point in the full-frequency likelihood. We use parameter covariances from MCMC analyses (see Section VII for details) to obtain the relevant Fisher matrices. This allows us to compare parameter constraints from the individual frequency-spectrum likelihoods to the full-frequency likelihood and among themselves. There are five independent tests for each TT, TE, and EE channel, and the PTE threshold to pass is $0.05/15 = 0.0033$.

When analyzing subsets of the data, it is necessary to impose additional priors on the calibration and polarization efficiency parameters (see §III D), as these parameters are not sufficiently constrained by the data alone. The specific priors adopted in these cases are summarized in Table IX.

In Fig. 16 we show the results for the parameters H_0 , $\Omega_b h^2$, $\Omega_c h^2$, and n_s . The PTEs for inter-frequency comparisons are reported in Table II and pass the associated threshold, indicating once again the excellent internal consistency of our data. The ordering of the frequency spectra in the plot is designed to highlight potential foreground mismodeling. The rising (falling) trend for $\Omega_b h^2$ (n_s) on the TT channel might be indicative of a foreground contribution, but the PTEs are still consistent with the null hypothesis. Note that during the blind stage of the analysis, the absolute parameter values were obscured and no comparison between temperature and polarization data was carried out. We performed the latter after unblinding, as discussed in §VII B 1, but report it in the figure for completeness.

Overall, we find that the frequency channels are consistent across the full multipole range considered. The

pipeline is validated to be unbiased on simulations and the analytic covariance matrix is shown to provide an accurate description of the data. We also confirm that the analysis is largely insensitive to the choice of fiducial cosmological model. Finally, the Λ CDM constraints derived from individual frequency spectra are consistent with each other and with those from the full-frequency likelihood, as illustrated in Fig. 16. We conclude that the analysis pipeline is robust and suitable to be used for cosmological inference.

VII. COSMOLOGICAL ANALYSIS

Having demonstrated the internal consistency of our data and verified our ability to model it, we now use it for cosmological inference. This section is organized as follows. After introducing all the data sets that are used (§VII A), we first report constraints using only CMB data (§VII B). We assess the consistency of the SPT-3G D1 T&E data with Λ CDM (§VII B 1) and with other CMB data, and then report results for the standard model for SPT alone and in combination with other CMB data (§VII B 2). We test whether the signature of gravitational lensing in CMB data is consistent with the Λ CDM prediction (§VII B 3) and afterwards report constraints on new light particles (§VII B 4) and reconstruct the recombination history (§VII B 5). Following this, we consider joint constraints from CMB and BAO data. We first evaluate the consistency of the relevant data sets in Λ CDM (§VII C) before reporting constraints on extended cosmological models (§VII D). We revisit models from §VII B 3-§VII B 5 in §VII D 1-§VII D 3, now adding BAO data, and report results on additional models not considered before (§VII D 4-§VII D 7).

For the analyses presented here, we explore parameter posteriors via a Markov Chain Monte Carlo (MCMC) approach using *Cobaya* [104]. We consider chains with a Gelman-Rubin statistic of $R - 1 \sim 0.02$ to be converged unless otherwise specified. Although power spectrum emulators (§V E) were indispensable during this analysis, all of the final results reported below were generated using traditional Boltzmann solvers²⁰ unless otherwise stated. Almost all of the SPT results in this section are run with the *SPT-lite* likelihood (§V D). We use the full-frequency likelihood only when analyzing SPT TT, EE, or TE alone, or when showing SPT-3G D1 only results in Λ CDM.

A. Additional data sets

We present a list of data sets used in this work in Table III. The CMB power spectrum measurements

²⁰ We used CLASS v3.2.3 [90] and CAMB v1.5.8 [89]

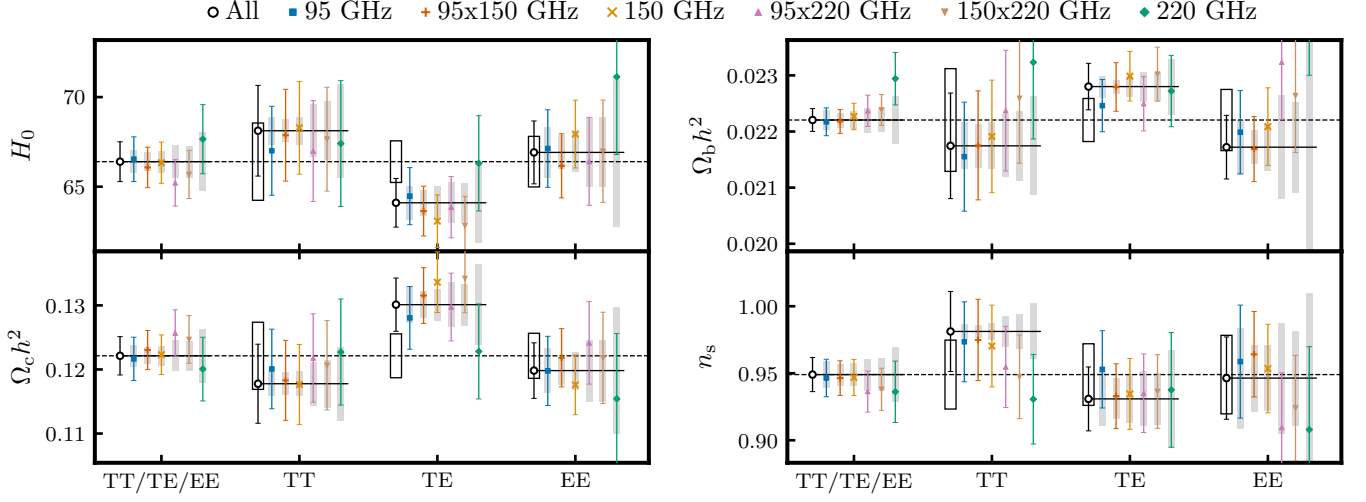


FIG. 16. Cosmological parameters obtained from single frequency-spectrum likelihoods (95×95 GHz: blue square, 95×150 GHz: orange cross, 150×150 GHz: yellow x, 95×220 GHz: pink triangle, 150×220 GHz: brown upside-down triangle, 220×220 GHz: green diamond) and the full-frequency likelihoods (open black circles) for combined T&E and individual TT, TE, and EE fits. The whiskers show the 1σ error bars obtained by the given subset, while the shaded area shows the size of the expected 1σ fluctuation between the TT, TE, and EE likelihoods and the subset, calculated via Eq. (51). The open boxes similarly show the size of expected fluctuation of the full-frequency TT, TE, and EE likelihoods from the complete likelihood. Although we display parameters on an absolute scale for readability, we only looked at the difference before unblinding. Similarly, the comparison between temperature and polarization results was carried out after unblinding (see §VII B 1). There appears to be a frequency-dependent trend in temperature constraints on $\Omega_b h^2$ and n_s , which is likely influenced by the foreground model. However this trend is not statistically significant and we report excellent consistency of our data across frequencies and spectra. Similarly, the excursion of the TE multi-frequency constraints is not statistically significant, as indicated in Table IV.

T&E \backslash TT	TT					
	95	95 × 150	150	95 × 220	150 × 220	220
95		0.93	1.00	0.48	0.79	0.49
95 × 150	0.87		0.97	0.83	0.91	0.40
150	0.86	0.82		0.82	0.84	0.55
95 × 220	0.62	0.36	0.51		0.99	0.29
150 × 220	0.36	0.19	0.08	0.74		0.33
220	0.25	0.16	0.17	0.39	0.60	

TE \backslash EE	EE					
	95	95 × 150	150	95 × 220	150 × 220	220
95		0.07	0.05	0.92	0.11	0.82
95 × 150	0.99		0.82	0.65	0.89	0.55
150	0.99	0.70		0.28	0.99	0.33
95 × 220	0.55	0.89	0.66		0.05	0.54
150 × 220	0.99	1.00	0.89	0.93		0.14
220	0.60	0.74	0.88	0.73	0.95	

TABLE II. Inter-frequency PTEs for the different frequency combinations. The PTEs are computed within each channel combination, and for comparison of single frequencies only. For example, the consistency between parameters derived from TT 95 GHz and all other TT frequencies is presented in the first row of the first section of the table; the PTE of the comparison of TT 95 GHz constraints with the TT 95×150 GHz constraints is 0.93. The left panel displays the PTEs for the T&E and TT data. The right panel presents the PTEs for the TE and EE data. The comparison is performed on the 5 Λ CDM parameters excluding τ . We obtained the cosmological parameter covariance matrix according to Eq. (51).

presented in this work do not cover large-scale E-mode polarization anisotropies and hence cannot constrain with high precision the optical depth to reionization, τ_{reio} . We impose a Gaussian prior on $\tau_{\text{reio}} \sim \mathcal{N}(0.051, 0.006)$ based on [32] unless stated otherwise. We combine our primary CMB data with the SPT-3G gravitational lensing potential reconstruction presented in [16]. We expect a vanishing correlation between the lensing data and the lensed CMB band powers presented in this work [110] and we, therefore, combine the data sets

at the likelihood level after verifying their consistency (see §VII B 2).

We compare and combine our results with CMB data from *Planck* and ACT. For *Planck*, we choose to use the primary CMB data from the PR3 release [34] and the lensing data from PR4 [35], which is more constraining than the PR3 lensing data. This is the same combination of *Planck* data adopted in [16]. We take advantage of the *python* implementation of the *Planck* `clik` likelihood,

Name	Data Set
τ_{reio} prior	$\tau_{\text{reio}} \sim \mathcal{N}(0.051, 0.006)$ [32], unless specified otherwise
SPT-3G D1 T&E	This work, i.e. T&E band powers from SPT-3G D1 observations
MUSE $\phi\phi$	$\phi\phi$ band powers from SPT-3G D1 observations [16]
SPT-3G D1	SPT-3G D1 T&E + MUSE $\phi\phi$
<i>Planck</i> T&E	<i>Planck</i> 2018 PR3 Plik high- ℓ T&E + low- ℓ TT [33]
<i>Planck</i> $\phi\phi$	<i>Planck</i> NPIPE PR4 $\phi\phi$ [35]
<i>Planck</i>	<i>Planck</i> T&E + <i>Planck</i> $\phi\phi$
SPT+ <i>Planck</i>	SPT-3G D1 + <i>Planck</i>
ACT DR6 T&E	ACT DR6 T&E [5–7]
ACT DR6 $\phi\phi$	ACT DR6 $\phi\phi$ [36, 105]
ACT DR6	ACT DR6 T&E + ACT DR6 $\phi\phi$
P-ACT T&E	<i>Planck</i> + ACT DR6 combined T&E likelihood [6]
P-ACT $\phi\phi$	<i>Planck</i> + ACT DR6 combined $\phi\phi$ likelihood [35]
P-ACT	P-ACT T&E + P-ACT $\phi\phi$
SPT+ACT T&E	SPT-3G D1 T&E + ACT DR6 T&E
SPT+ACT $\phi\phi$	MUSE $\phi\phi$ + ACT DR6 $\phi\phi$
SPT+ACT	SPT-3G D1 + ACT DR6
CMB-SPA T&E	P-ACT T&E + SPT T&E
CMB-SPA $\phi\phi$	P-ACT $\phi\phi$ + MUSE $\phi\phi$ [39]
CMB-SPA	CMB-SPA T&E + CMB-SPA $\phi\phi$
DESI	DESI DR2 BAO data [3]
SDSS	SDSS BAO data [106–109]
SPT+DESI	SPT-3G D1 + DESI DR2 BAO data

TABLE III. Summary of data sets used in the analysis.

clipy²¹ [34]. For ACT, we use the DR6 lensing and primary CMB data [5–7, 36, 105]. When needed, we rerun the ACT analyses imposing our baseline τ_{reio} prior to allow for a consistent comparison. We use the ACT-lite likelihood for all of the results in this section, except when running TT data alone, for which we use the ACT DR6 multi-frequency likelihood. We combine the SPT and ACT DR6 data under the assumption that the two data sets are uncorrelated and call this combination SPT+ACT. We also combine *Planck*, SPT, and ACT DR6 and refer to this as CMB-SPA. When doing this, in order to minimize the correlations between the *Planck* and ACT DR6 data sets, we use the multipole cuts of the P-ACT combination of [6]. Note in particular that P-ACT uses only ACT DR6 TT data at $\ell > 1000$, excising the *Planck* TT high- ℓ data; this choice is relevant in §VII C and §VII D as the high- ℓ TT data of ACT DR6 and *Planck* prefer slightly different cosmologies. We assume no correlation between SPT and other data sets, justified by the small overlap in

the sky regions, see Fig. 3.²² We neglect correlations between lensing measured by ACT DR6 and SPT, as these (small) correlations were shown to have a negligible effect on cosmological parameters in [39].

We further contextualize our results using non-CMB data. Here, we draw on the latest BAO measurements from DESI DR2 data [3]. We also consider the SDSS BAO data [106–109] and SNe Ia data from the Pantheon+ data set [111].

²¹ <https://github.com/benabed/clipy>

²² As *Planck* data are based on full sky observations, the fraction of the SPT-3G Main footprint within the *Planck* mask is minimal. For ACT, the common area observed constitutes about 10% of the ACT DR6 mask.

Spectrum	All	TT	TE	EE
All	-	0.4σ	1.2σ	0.6σ
TT	0.67	-	1.0σ	0.3σ
TE	0.22	0.31	-	1.0σ
EE	0.57	0.78	0.33	-

TABLE IV. PTE table indicating levels of consistency between cosmological constraints from temperature and polarization and their combination. The lower triangle indicates the PTE values, while the upper triangle indicates the corresponding Gaussian equivalent σ values.

B. Constraints from CMB data

1. Consistency of the Λ CDM model with SPT data

We find that the Λ CDM model provides a good description of the SPT-3G D1 T&E spectra.²³ For our full data we report a best-fit χ^2 value of 1359, which corresponds to a PTE of 0.52.²⁴ We present the best-fit Λ CDM model for the band powers in Fig. 9 together with the measured band powers, and display the residuals between the data and the model. Λ CDM also fits well each spectrum individually; analyzing only TT, TE, or EE spectra we report χ^2 (PTE) values of 267 (0.84), 631 (0.52), and 429 (0.38) (for 291, 633, and 421 effective degrees of freedom), respectively. While this is a general affirmation that Λ CDM is a good description of the SPT data, we verify the consistency of model predictions between temperature and polarization as well as across angular scales in more detail.

First, we check the agreement of Λ CDM parameter constraints from temperature and polarization data. This not only allows us to assess the ability of Λ CDM to jointly describe temperature and polarization data, but it is also a powerful test to understand frequency-coherent biases in specific channels.²⁵ To evaluate consistency between spectra, we use the same framework as for inter-frequency parameter consistency in §VID. The calculated parameter correlation matrices are non-trivial, and we illustrate the expected correlation of constraints produced by data subsets for H_0 , $\Omega_c h^2$, n_s and $\Omega_b h^2$ in Fig. 17 as an example. We find consistency (at the $\sim 1\sigma$ level) between constraints from all individual and joint TT, TE, and EE likelihoods, which we show in Fig. 18 and Table IV. The precise measurements of the CMB polarization anisotropies drive our combined T&E result;

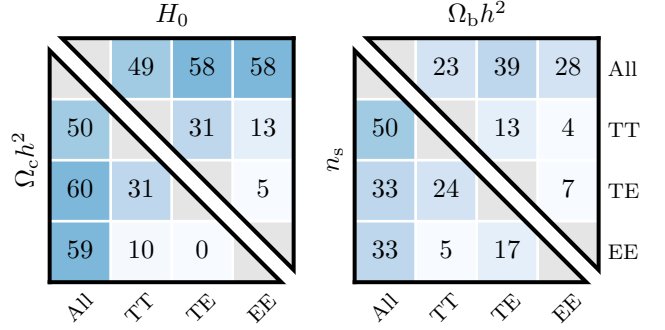


FIG. 17. Correlation matrix between constraints from the individual and joint TT, TE, and EE likelihoods. We show, as an example, the correlations for the parameter constraints on H_0 , $\Omega_c h^2$, n_s , and $\Omega_b h^2$. For instance, the determination of H_0 from TT is correlated at 31% with the measurement of H_0 from TE. The correlation matrices are non-trivial and their calculation allows for interesting data and model consistency tests.

excising the TT data only slightly loosens constraints, as shown in Fig. 18.

Next, we evaluate the consistency of Λ CDM parameter constraints obtained from different angular scales. We split our data into three multipole ranges, defined from the relative signal-to-noise of our data, see Fig. 11: low ($\ell < 1000$), intermediate ($\ell \in [1000, 2000]$), and high ($\ell > 2000$) and perform a cosmological analysis on each of these subsets. We then compare the results using the same framework as in §VID. Again, we report no discrepancies between the subsets and show our results in Fig. 19.

Together, these tests speak to the ability of the Λ CDM model to consistently and accurately predict (1) temperature and polarization data and (2) data across a wide range of angular scales. Moreover, they indicate that there are no substantial frequency-coherent biases in our data, assuming Λ CDM.

2. Λ CDM results

After verifying that the Λ CDM model provides a good description of the power spectrum measurements presented in this work, we establish that MUSE lensing and SPT-3G D1 T&E data are compatible, and discuss the combined SPT-3G D1 results. We then verify that SPT-3G D1 is consistent with other CMB data sets, before reporting joint constraints and situating these in the wider cosmological context.

From now on, we combine primary CMB data with CMB lensing reconstructions as per Table III, unless otherwise explicitly mentioned. We first assess the consistency between the SPT-3G D1 T&E measurements and the MUSE lensing data. The difference in the $S_8(0.25) \equiv \sigma_8(\Omega_m/0.3)^{0.25}$ measurements corresponds to a 0.5σ fluctuation, indicating agreement between these

²³ We note that all Λ CDM results presented in this section are based on the multi-frequency SPT-3G D1 T&E likelihood and are obtained using CAMB as the Boltzmann solver.

²⁴ Following the methodology of [112] (see Eq. 29 therein) we count 1362 effective degrees of freedom for the multi-frequency likelihood in Λ CDM.

²⁵ This test is model dependent, which is why we did not perform it prior to unblinding.

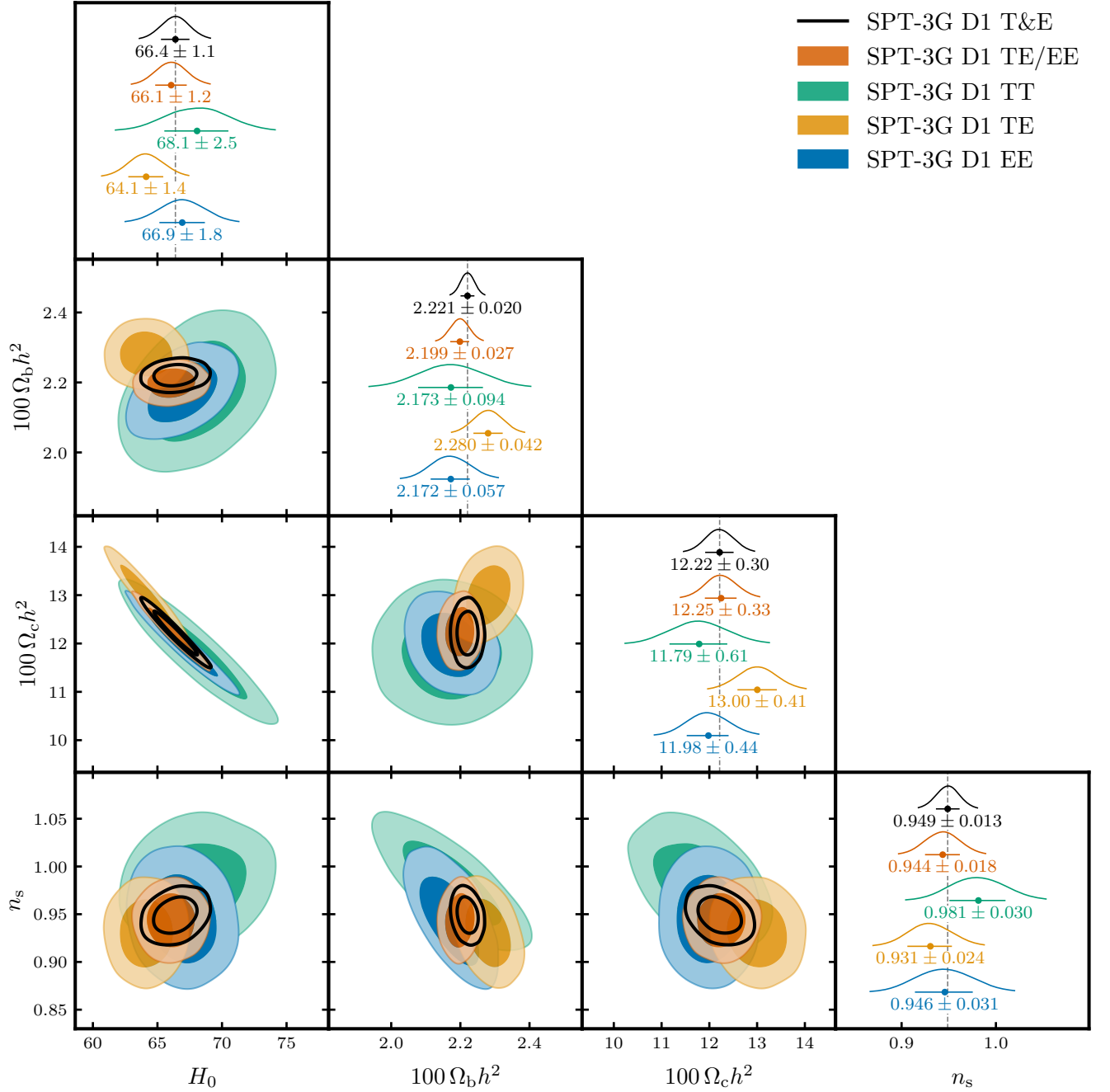


FIG. 18. Comparison of Λ CDM constraints from the SPT-3G D1 T&E data. The polarization data are exceptionally constraining and, in particular, the TE data holds most of the cosmological information. We report consistent results from temperature and polarization data in Λ CDM, with PTEs reported in Table IV.

two probes. We find $S_8(0.25) = 0.824 \pm 0.016$ for MUSE and $S_8(0.25) = 0.838 \pm 0.021$ for SPT-3G D1 T&E. Since the correlation between the lensing reconstruction and the lensed CMB data is expected to be negligible [110], and this has been confirmed by simulations, we combine the two data sets at the likelihood level. Comparing our measurement of the lensed CMB power spectrum to the MUSE reconstruction of the unlensed CMB on

the other hand is not straightforward as the correlation between the two data sets remains to be accurately quantified.²⁶ Having demonstrated the consistency of the SPT-3G D1 T&E and MUSE lensing data, we now

²⁶ The means of the marginalized one-dimensional posterior distributions of Λ CDM parameters derived from the combination of

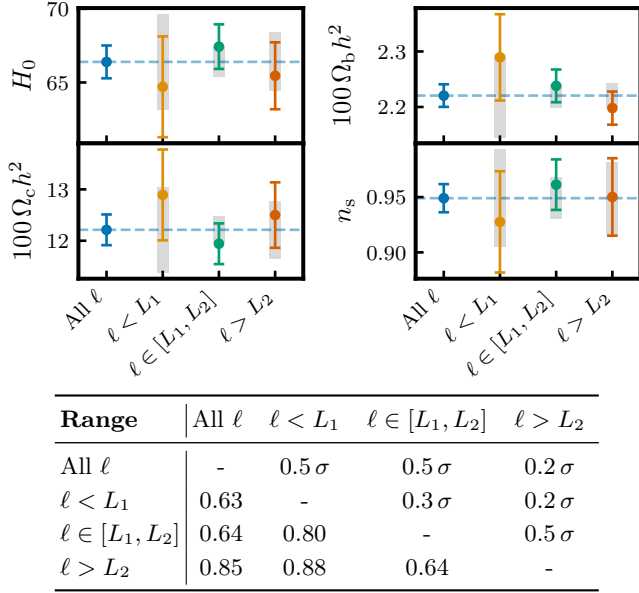


FIG. 19. *Top*: Comparison of Λ CDM parameter constraints from the low, intermediate, and high multipole ranges, with $L_1 = 1000$ and $L_2 = 2000$. The whiskers indicate the 68% confidence intervals of the marginalized posteriors, while the gray shaded regions indicate the expected variation of the parameters assuming the full ℓ -range following [113]. The different subsets are effectively independent of one another and their corresponding constraints are consistent. This is also true for the comparison with the full likelihood, in which case we do account for the expected correlation due to the shared data (gray boxes). We conclude that our data are consistent across angular scales given Λ CDM. *Bottom*: PTE table indicating level of consistency between cosmological constraints from different ℓ ranges. The lower triangle indicates the PTE values, while the upper triangle indicates the corresponding Gaussian equivalent σ values.

combine them and report constraints on cosmological parameters in Table I. We show marginalized posteriors for some parameters in Fig. 2.

We highlight the tight constraints that the SPT-3G D1 data place on cosmology. Notably, for H_0 , SPT data yield a $\sigma(H_0) = 0.60 \text{ km s}^{-1} \text{ Mpc}^{-1}$ constraint, compared to $0.49 \text{ km s}^{-1} \text{ Mpc}^{-1}$ for *Planck* and $0.64 \text{ km s}^{-1} \text{ Mpc}^{-1}$ for ACT DR6. Similarly, $100\Omega_c h^2$ is determined using SPT data with a precision of 1.3% to $\sigma(100\Omega_c h^2) = 0.16$; the *Planck* and ACT DR6 data achieve constraints of 0.11 (0.92% precision) and 0.17 (1.4% precision) on this parameter, respectively. For these two parameters, the SPT data benefit greatly from the degeneracy-breaking power of CMB lensing.

The constraints from the SPT data are in excellent agreement with other contemporary CMB experiments. The contours corresponding to *Planck* overlap visibly with those for SPT-3G D1 in Fig. 2. This agreement also holds up quantitatively; calculating the agreement over $(H_0, \Omega_b h^2, \Omega_c h^2, n_s, 10^9 A_s e^{-2\tau_{\text{reio}}})$ we obtain PTEs that correspond to one-dimensional Gaussian fluctuations of 0.4σ when comparing SPT-3G D1 to *Planck* and 1.1σ when comparing SPT-3G D1 to ACT DR6.²⁷ Overall, the agreement between CMB experiments at this level of precision is a remarkable achievement of the standard model of cosmology. The three data sets contain independent information and span a wide range of angular scales for both temperature and polarization data. This agreement motivates the combination of the data sets to further improve cosmological constraints.

While *Planck* remains the most constraining single CMB experiment, the SPT+ACT combination of SPT and ACT data achieves equally tight constraints on most Λ CDM parameters, as shown in Fig. 2 and quantified in Table I. This is a significant milestone for modern cosmology; for the first time, ground-based experiments reach *Planck*'s constraining power, most notably on the Hubble constant ($66.59 \pm 0.46 \text{ km s}^{-1} \text{ Mpc}^{-1}$ vs. $67.41 \pm 0.49 \text{ km s}^{-1} \text{ Mpc}^{-1}$ for *Planck*) and the amplitude of matter fluctuations parameterized by σ_8 (0.8169 ± 0.0042 vs. 0.8099 ± 0.0051 for *Planck*). At the same time, the SPT+ACT data set is consistent with the *Planck* results at 1.2σ (approximating them as completely independent).

We verified that the SPT+ACT constraints are still as good as *Planck* even if we replace the *Planck* 2018 high- ℓ T&E likelihood, *Plik*, which is the baseline for our *Planck* data set, with the *camspec* likelihood from [114].²⁸ We call the *Planck* data set using the *Camspec* likelihood *Planck-Camspec*. We find $H_0 = 67.22 \pm 0.44 \text{ km s}^{-1} \text{ Mpc}^{-1}$ and $\sigma_8 = 0.8076 \pm 0.0050$ from *Planck-Camspec*. The consistency between *Planck-Camspec* and SPT-3G D1 remains excellent, at the 0.2σ level. Similarly, we find consistency between *Planck-Camspec* and SPT+ACT at the 1.8σ level.²⁹ We conclude that the choice of *Planck* likelihood does not have a large impact on the findings reported above.

Crucially, the SPT and ACT data are highly complementary to *Planck*. The satellite data leverage full-sky access to deliver measurements of large angular

SPT-3G Main 19/20 EE and MUSE $\phi\phi$ are within 1.1σ of the MUSE EE + $\phi\phi$ results. Expensive joint simulations are needed to understand whether these parameter shifts are expected, given that the two pipelines use different methodologies, survey masks, and angular scales. This work is currently under way.

²⁷ Since we derive our absolute calibration from *Planck*, we verify that the parameter-level agreement persists when excluding the combined amplitude parameter, $10^9 A_s e^{-2\tau_{\text{reio}}}$, from the comparison. In this case, the agreement between SPT-3G D1 and *Planck* holds steady at 0.4σ , while the distance to ACT DR6 slightly increases, to 1.4σ .

²⁸ The *Camspec* likelihood from [114] is based on *Planck* PR4 maps, which contain about 10% more data than *Planck* PR3, and use a larger fraction of the sky with respect to the *Plik* PR3 likelihood.

²⁹ Note that the ACT DR6 results show a worse level of agreement with the *Camspec* PR4 likelihood compared to the *Plik* 2018 one, see [6].

scales that are difficult to constrain from the ground and—by virtue of measuring more modes—are the best measurements in the sample-variance-dominated regime. The SPT+ACT data, on the other hand, have been collected by instruments with higher resolution and significantly more detectors, thus providing exquisite measurements of the polarization of the CMB on scales where the *Planck* data is noise-dominated. Together, these data form an incredibly rich cosmological data set. Given the agreement demonstrated above, we combine SPT+ACT and *Planck* into CMB-SPA. We report constraints from this combination in Fig. 2 and Table I; these are the most precise determinations of Λ CDM parameters from CMB observations to date.

We use the square-root of the determinant of parameter covariance matrices as a measure of the allowed volume in higher dimensional spaces (i.e. $\sqrt{\det C}$, where C is the relevant matrix);³⁰ ratios of this metric allow us to quantify the volume reduction SPT data enable in joint constraints. In the five-dimensional space of $(H_0, \Omega_c h^2, \Omega_b h^2, n_s, 10^9 A_s e^{-2\tau_{\text{reio}}})$, adding SPT to *Planck* reduces the allowed parameter volume by a factor of 1.9. Similarly, when adding SPT to ACT data to form SPT+ACT, the allowed region shrinks by a factor of 2.8. These are significant improvements and speak to the constraining power of the SPT data. Compared to the previous SPT-3G T&E results presented in [14], the new SPT-3G D1 T&E data decrease the allowed parameter volume by a factor of 2.4.

Despite the small observed sky area this release is based on, SPT data are able to constrain the scalar spectral index to $n_s = 0.951 \pm 0.011$, which disfavors a scale-invariant spectrum of initial density perturbations at 4.1σ . While ACT and *Planck* data are more sensitive to this parameter, the SPT data make a non-negligible contribution to the combined constraints. ACT data alone rule out $n_s = 1$ at 4.6σ , which becomes 5.7σ in SPT+ACT when SPT data are added. Similarly, the preference for a scale-dependent spectrum increases from 7.8σ for P-ACT-L [6] to 10.5σ for CMB-SPA when SPT data are included. In addition to shrinking the width of the n_s posterior in joint constraints, SPT data also shift the central value low, which increases the evidence for $n_s < 1$.

We situate our results in the wider cosmological landscape, beginning with constraints on the expansion rate today:

$$H_0 = 66.66 \pm 0.60 \text{ km s}^{-1} \text{ Mpc}^{-1} \text{ for SPT-3G D1,} \quad (52)$$

$$H_0 = 66.59 \pm 0.46 \text{ km s}^{-1} \text{ Mpc}^{-1} \text{ for SPT+ACT,} \quad (53)$$

$$H_0 = 67.24 \pm 0.35 \text{ km s}^{-1} \text{ Mpc}^{-1} \text{ for CMB-SPA.} \quad (54)$$

These results are in stark contrast to the local universe determination of the expansion rate by SH0ES, i.e.

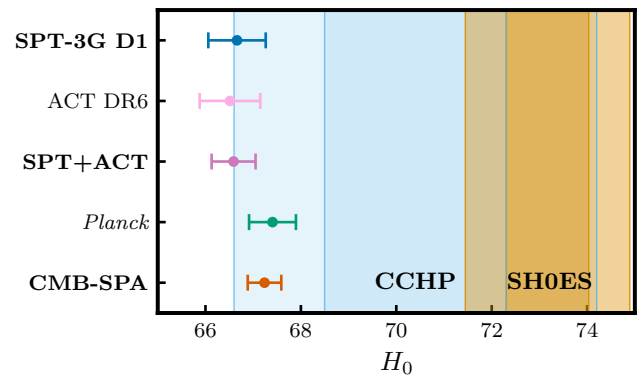


FIG. 20. Status of H_0 constraints. The data sets are described in Table III and the values are reported in Table I. In bold, we highlight the results of this work. We show the SH0ES and CCHP late-universe results in orange and blue, respectively.

$H_0^{\text{SH0ES}} = 73.17 \pm 0.86 \text{ km s}^{-1} \text{ Mpc}^{-1}$ [2, 103] as shown in Fig. 20. We report a 6.2σ tension between SH0ES and the result from SPT data alone, and a 6.4σ tension between SH0ES and CMB-SPA. CMB-SPA comprises three nearly independent CMB data sets that have different weighting across angular scales and relative contributions of the TT/TE/EE/ $\phi\phi$ spectra³¹ and were produced by different pipelines, all of which involve stringent consistency tests, yet they all agree and individually yield comparable offsets from SH0ES. Therefore, we are inclined to consider a systematic error in the CMB data an unlikely explanation for this discrepancy.

Local universe determinations of the Hubble constant have also undergone much scrutiny. While the analysis of SH0ES is not the only result, it is currently the most statistically precise one. We also highlight the work of the CCHP collaboration [115], who calibrate SNe Ia with Cepheids, the J region of the asymptotic giant branch, and the tip of the red giant branch. The distances they obtain with these techniques to nearby galaxies hosting SNe Ia are consistent with the same distances found by the SH0ES team. For the Hubble constant they report $H_0 = 70.4 \pm 1.9 \text{ km s}^{-1} \text{ Mpc}^{-1}$, which is 1.6σ lower than the SH0ES result and 2σ higher than the CMB-SPA constraint. While it is possible that the difference between CMB and local data sets is caused by a real failure of Λ CDM and not by systematics in any of the data sets, signatures of beyond- Λ CDM models that can accommodate a high H_0 have yet to be clearly discerned in CMB data [Section VII in this work, 7].

We now turn our attention to constraints on structure growth. We report constraints on the amplitude of matter fluctuations, σ_8 , and on the matter density, Ω_m ,

³⁰ The inverse of this is commonly referred to as the figure of merit.

³¹ From our TE/EE spectra alone we constrain $H_0 = 66.1 \pm 1.2$ (see Fig. 18), which is in 4.9σ tension with SH0ES.

of:

$$\left. \begin{aligned} \sigma_8 &= 0.8158 \pm 0.0058, \\ \Omega_m &= 0.3246 \pm 0.0091 \end{aligned} \right\} \text{ for SPT-3G D1,} \quad (55)$$

$$\left. \begin{aligned} \sigma_8 &= 0.8169 \pm 0.0042, \\ \Omega_m &= 0.3277 \pm 0.0072 \end{aligned} \right\} \text{ for SPT+ACT,} \quad (56)$$

$$\left. \begin{aligned} \sigma_8 &= 0.8137 \pm 0.0038, \\ \Omega_m &= 0.3166 \pm 0.0051 \end{aligned} \right\} \text{ for CMB-SPA.} \quad (57)$$

We highlight that the SPT results correspond to a 0.7% determination of σ_8 , which is competitive with both *Planck* and ACT.

Other cosmological probes are also able to constrain structure growth; specifically, individual and joint analyses of the following measurements allow for precise determinations: galaxy weak lensing [116–122], galaxy clustering and its combination with galaxy weak lensing (referred to as 3×2 point analyses) [38, 123, 124], and galaxy cluster statistics [40, 125]. These probes are less sensitive to either parameter individually, but constrain the combination $S_8(\alpha) = \sigma_8(\Omega_m/0.3)^\alpha$ tightly, where α is typically chosen to maximize the precision in $S_8(\alpha)$ for a given experiment. Until recently, there was moderate statistical evidence for a significantly lower $S_8(0.5)$ value from cosmic shear [118–120] and 3×2 point analyses [123, 124, 126] compared to primary CMB and CMB lensing constraints. However, many of the most recent analyses, with more data and improved modeling, as well as following comprehensive comparisons between experiments, are reporting $S_8(0.5)$ values that differ from the CMB Λ CDM constraints by less than 2σ [37, 127, 128]. We are now in a position where constraints on $S_8(\alpha)$ from vastly different cosmological probes agree. For example, the combined primary CMB constraint agrees at 1.8σ with the 3×2 point analysis of DES-Y3 [38] (using $\alpha = 0.5$), at 0.1σ with the CMB lensing analysis of [39] (using $\alpha = 0.25$), and at 1.1σ with the galaxy cluster statistics analysis of [40] (using $\alpha = 0.3$), as illustrated in Fig. 21. Moreover, the latest KiDS cosmic shear analysis agrees at 0.86σ with the combined primary CMB result [37]. This consistency is impressive, as the underlying experiments target different phenomenological signatures and span a wide range of redshifts and scales, which lead to different degeneracy directions in the Ω_m - σ_8 plane. Additionally, the instruments carrying out the surveys observe at a range of wavelengths and are affected by different systematics. In this light, the consistent determination of $S_8(\alpha)$ is remarkable.

Finally, we estimate τ_{reio} excising the *Planck* PR4 based prior used in our baseline. The amplitude of the CMB anisotropy power spectra is proportional to $A_s e^{-2\tau_{\text{reio}}}$. In the absence of large scale E-mode information, or a prior on τ_{reio} , the degeneracy between A_s and the amplitude suppression factor $e^{-2\tau_{\text{reio}}}$ can be broken via gravitational lensing, which is not affected by τ_{reio} . We remove the τ_{reio} prior from CMB-SPA

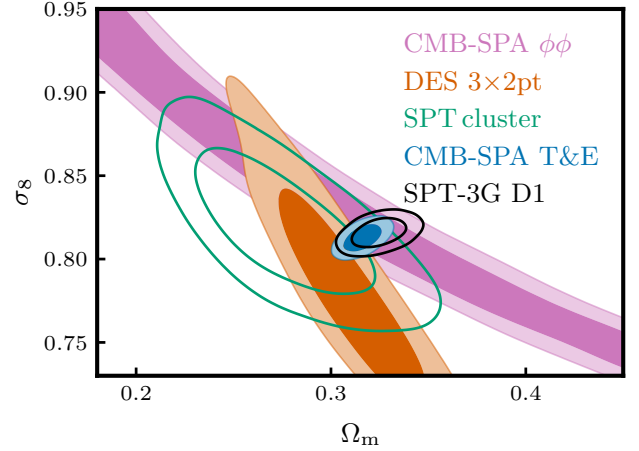


FIG. 21. Constraints on structure growth in the Ω_m - σ_8 plane using SPT-3G D1 (black, this work), primary CMB-SPA data (blue, this work), combined SPT, ACT, *Planck* CMB lensing data (pink, [39]), SPT galaxy cluster statistics (green, [125]), and DES-Y3 3×2pt analysis with fixed neutrinos (orange, [38]). The different data sets are sensitive to different parameterizations of $S_8(\alpha) = \sigma_8(\Omega_m/0.3)^\alpha$, where α corresponds to the slope in the Ω_m - σ_8 plane. Yet, the four cosmological probes yield consistent results.

and report $\tau_{\text{reio}} = 0.078 \pm 0.013$, which is within 2σ of our *Planck*-based prior of $\mathcal{N}(0.051, 0.006)$. We note that because CMB lensing is driving this low-E-free τ constraint, a slight preference for higher τ in CMB-SPA can effectively be restated as a slight CMB-SPA preference for an amplitude of lensing larger than the Λ CDM expectations, as shown in the next section. Moreover, this result is in excellent agreement with the corresponding *Planck* constraint on τ_{reio} free of large scale EE (“low-E”) data, $\tau_{\text{reio}} = 0.079 \pm 0.018$, but about 30% more precise.³² Note that while the CMB-SPA-based constraint is twice as wide as the *Planck* low-E result, it is as precise as the WMAP final mission result of $\tau_{\text{reio}} = 0.089 \pm 0.014$ [100]. Still, the τ_{reio} constraint above is too wide to serve as an accurate cross-check of the *Planck* E-mode measurement.

3. Lensing amplitude

We now study the consistency of the signature of gravitational lensing in primary CMB and CMB lensing reconstruction data. We stress that the models used in this section are non-physical: they are designed to facilitate further tests of the validity of the Λ CDM description of CMB data. As has been widely reported, analyses of *Planck* primary CMB anisotropy power

³² Compared to `base_plikHM_TTTEEE_lowl_lensing` from [1]

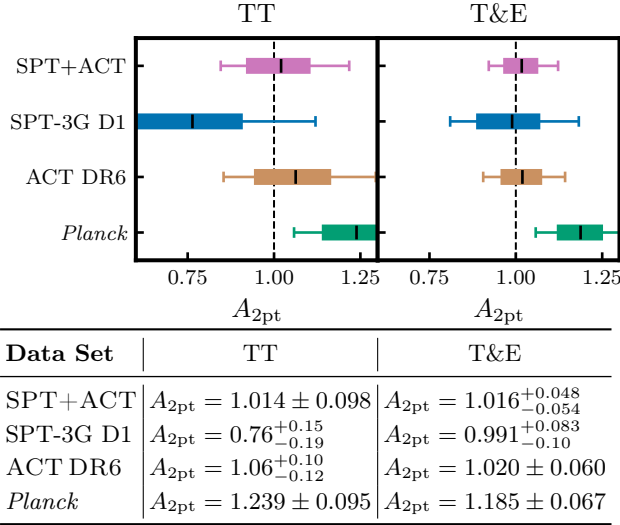


FIG. 22. Constraints on A_{2pt} , the amplitude of lensing inferred from effects on the T&E power spectra, from *Planck* (green), ACT DR6 (brown), SPT-3G D1 (blue), and SPT+ACT (purple) temperature data (left) and the combination of temperature and polarization data (right), both excluding lensing $\phi\phi$ data. We show the 68% and 95% confidence regions as the filled contours and whiskers, respectively. The black bar indicates the mean value of the posterior, while the dashed line indicates the Λ CDM expectation of $A_{2pt} = 1$. While the *Planck* data prefer a value of $A_{2pt} > 1$, driven by the temperature data, the ground-based data sets are consistent with the Λ CDM expectation. The combination of ACT and SPT data is in excellent agreement with the Λ CDM prediction. We report 68% confidence intervals in the table below the figure.

spectra show a preference for an excess in the effects of lensing on the primary CMB anisotropy power spectra compared to Λ CDM expectations at the $2\text{--}3\sigma$ level [1, 34]. This preference is largely driven by an apparent excess smoothing of the *Planck* high-multipole temperature data. We look for evidence of similar effects in the SPT data analyzed in this work, both alone and in combination with the latest ACT power spectrum measurements. Following [16], we introduce A_{2pt} , which quantifies the effect of gravitational lensing on CMB T&E power spectra. This parameter scales the amplitude of the lensing power spectrum used to predict the lensed CMB spectra from the unlensed expectation, but does not affect the prediction for the amplitude of the reconstructed CMB lensing potential ($\phi\phi$).

First, we consider constraints based exclusively on primary CMB temperature data. We show the marginalized posteriors of A_{2pt} for different combinations of CMB temperature data in Fig. 22. Constraints from SPT temperature data alone are relatively wide due to the limited survey area. They are in agreement with the standard model prediction of unity at 1.6σ . The combination of ground-based temperature power spectrum measurements, ACT+SPT TT (see Fig. 22) is

in excellent agreement with the Λ CDM expectation and the associated posterior mean is 2.3 standard deviations below the value preferred by the *Planck* temperature data.³³ While this may suggest that the features in the *Planck* spectrum that cause the observed preference for $A_{2pt} > 1$ may not be cosmological and rather statistical or systematic, we cannot make a definitive judgment at this point. We note that re-analyses of *Planck* data have generally led to a reduction of this feature [114, 129].

In the right panel of Fig. 22, we show the posteriors for the T&E case. For SPT data, the bulk of the constraining power lies in polarization data and we see a considerable tightening of the A_{2pt} posterior. The result is centered close to unity,³⁴ which is also true for the SPT+ACT combination. To further investigate the trends in the *Planck* data that project onto $A_{2pt} > 1$, it would be interesting to compare the *Planck* and ACT DR6 temperature data at $\ell > 1000$. However, this requires a quantification of the correlation of the two measurements, which is beyond the scope of this work.

We now also consider CMB lensing reconstruction data. We follow the prescription of [16] and add A_{recon} to our model. A_{recon} scales the prediction for the reconstructed CMB lensing power spectrum but has no effect on the prediction for the effects of lensing on the T and E power spectra. In other words, A_{2pt} and A_{recon} control independent subsets of the effects of the lensing power spectrum on CMB observables, which are often lumped together under one parameter A_L or A_{lens} (see below). Varying A_{2pt} and A_{recon} independently, we report

$$\left. \begin{aligned} A_{2pt} &= 0.986^{+0.078}_{-0.097} \\ A_{\text{recon}} &= 0.974^{+0.081}_{-0.11} \end{aligned} \right\} \text{ for SPT-3G D1,} \quad (58)$$

$$\left. \begin{aligned} A_{2pt} &= 1.026 \pm 0.048 \\ A_{\text{recon}} &= 0.990 \pm 0.050 \end{aligned} \right\} \text{ for SPT+ACT,} \quad (59)$$

$$\left. \begin{aligned} A_{2pt} &= 1.083 \pm 0.037 \\ A_{\text{recon}} &= 1.048 \pm 0.031 \end{aligned} \right\} \text{ for CMB-SPA.} \quad (60)$$

With these results, we see no evidence for inconsistent signatures of gravitational lensing between primary CMB and CMB lensing data for ground-based and large-scale satellite data. The CMB-SPA results are consistent with

³³ We do not quote a quantitative statistic for the consistency of the SPT+ACT and *Planck* posterior distributions as SPT+ACT contains ACT data that is significantly correlated with *Planck* data.

³⁴ Past SPT analyses have yielded $\lesssim 2\sigma$ fluctuations below one in A_{2pt} , however we stress that there is no one-to-one comparison. The closest case is [14] which is based on observations of the same part of the sky. However, the data presented here have a substantially lower noise level, which weights the constraint more towards polarization and shifts the posterior towards unity. Significant differences in the observation fields and in the spectra used compared to [9] and [64] prevent a meaningful comparison.

the Λ CDM expectation at the 1.7σ level in the two-dimensional space of $(A_{2\text{pt}}, A_{\text{recon}})$.

We now vary the amplitude of the lensing power spectrum and the signature of gravitational lensing in the primary CMB coherently, i.e. $A_{2\text{pt}} = A_{\text{recon}}$. We collapse the two parameters into one, A_{lens} , and report:

$$A_{\text{lens}} = 0.972^{+0.079}_{-0.089} \text{ for SPT-3G D1}, \quad (61)$$

$$A_{\text{lens}} = 1.011^{+0.045}_{-0.051} \text{ for SPT+ACT}, \quad (62)$$

$$A_{\text{lens}} = 1.055 \pm 0.030 \text{ for CMB-SPA}, \quad (63)$$

These results are all consistent within 1.9σ with the standard prediction of $A_{\text{lens}} = 1$. The fact that the amplitude of gravitational lensing measured in CMB data is consistent with the cosmology preferred by the lensing-marginalized CMB observables is yet another illustration of the internal consistency within and among the three CMB data sets treated here. We revisit this scenario adding BAO data in §VII D 1.

4. New light particles

We now search for the signature of new light particles in the early universe by using our data to estimate the effective number of relativistic species, N_{eff} . This parameter quantifies the density of relativistic particles other than photons. It is defined in terms of the density expected from an equivalent number of relativistic species under the assumption they have negligible chemical potential and none of the entropy of the electrons and positrons gets transferred to the neutrinos, so that:

$$\rho_{\text{rad}} = \rho_{\gamma} \left[1 + (7/8) \times (4/11)^{4/3} N_{\text{eff}} \right], \quad (64)$$

where ρ_{rad} is the radiation density and ρ_{γ} is the photon density. In the standard model, a small fraction of the entropy of the electrons and protons does get transferred to the neutrinos so that $N_{\text{eff}} = 3.044$ [133–136].

The impact of light relics on cosmological observables, and the CMB in particular, has been studied in many papers [137–144]. One of the more recent treatments [144] emphasized the connection between changes to various rate ratios and dimensionless observables (such as those of the CMB and BAO), as was done by [145] in the context of a time-varying gravitational constant G . The relevant rates here are free-fall rates for each component, $\sqrt{G\rho(z)}$, the Hubble rate $H(z)$, and the Thomson scattering rate $\sigma_T n_e(z)$. Extending Λ CDM to include additional light relics opens up a scaling transformation that preserves important rate ratios and, therefore, leads to an approximate degeneracy in the parameter space. BAO observables, $r_d H(z)$ and $r_d/D_M(z)$, where r_d is the comoving size of the sound horizon at the end of the baryon drag epoch, $H(z)$ the expansion rate, and $D_M(z)$ the transverse comoving distance, are particularly unaffected by this transformation so we refer to this as “BAO scaling.” Note that important

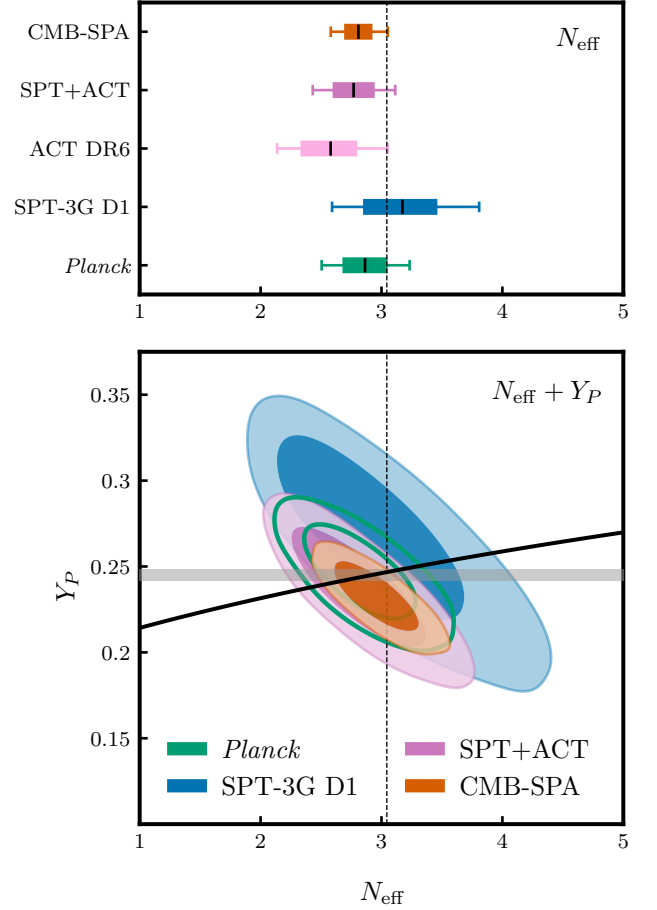


FIG. 23. Constraints on light relics from CMB data: SPT-3G D1 (blue), SPT+ACT (purple), *Planck* (green), and CMB-SPA (red). *Top panel:* Marginalized one-dimensional posterior on N_{eff} , when only varying this parameter. The SPT+ACT constraint is tighter than the *Planck* one, highlighting the exquisite constraining power of ground-based CMB data. All CMB constraints are within 2σ of the standard model prediction of $N_{\text{eff}} = 3.044$. *Bottom panel:* Constraints on N_{eff} and the primordial helium abundance Y_P , when the two parameters are varied simultaneously. The black line represents BBN consistency, while the gray band corresponds to the 1σ region of the Y_P constraint from observations of metal-poor galaxies [130]. We find no evidence for new light particles in the early universe from CMB data.

angular scales to which CMB power spectra are highly sensitive such as the angular scale of the sound horizon and the angular scale of the comoving Hubble length at matter-radiation equality, both projected from the last-scattering surface, are also approximately preserved under this transformation.³⁵

The degeneracy given CMB data is only approximate since other important rate ratios inevitably change. Chief

³⁵ These are exactly preserved if the redshift of last-scattering does not change.

	SPT-3G D1	SPT+ACT	CMB-SPA	SPT-3G D1 + DESI
N_{eff}	$3.17^{+0.29}_{-0.33}$	2.77 ± 0.17	2.81 ± 0.12	3.52 ± 0.23
Y_{P}	0.264 ± 0.022	$0.226^{+0.015}_{-0.013}$	0.2285 ± 0.0085	0.279 ± 0.022
N_{eff}	$2.97^{+0.40}_{-0.64}$	$2.85^{+0.32}_{-0.40}$	$2.99^{+0.22}_{-0.26}$	$3.63^{+0.39}_{-0.44}$
Y_{P}	$0.269^{+0.040}_{-0.030}$	$0.236^{+0.025}_{-0.021}$	0.231 ± 0.014	0.241 ± 0.034

TABLE V. Constraints on the effective number of relativistic species N_{eff} and the primordial helium abundance Y_{P} from CMB data alone and in combination with DESI. We report constraints on N_{eff} and Y_{P} when varying each parameter individually (first and second row) and when they are varied independently of each other at the same time (last two rows). The results are consistent with the standard model prediction of $N_{\text{eff}} = 3.044$, and the Y_{P} constraint is consistent with the Λ CDM BBN consistency prediction [131, 132] and the measurement of [130]. For the above models, we do not report constraints for SPT+ACT and CMB-SPA in combination with DESI as these combinations of CMB and BAO data do not meet our consistency requirements for joint analyses (see §VII D).

among these is the ratio of the Thomson scattering rate to the Hubble rate, $\sigma_{\text{T}} n_{\text{e}}(z)/H(z)$,³⁶ which impacts Silk damping and polarization generation [145]. Other important effects, recently reviewed in [144], further contribute to lift the degeneracies.

Due to these collective effects, the approximate degeneracy direction given CMB data is not along the BAO scaling direction, but instead along a parameter direction in which the matter density, ρ_{m} , scales up more slowly than the radiation density, ρ_{rad} , and the dark energy density, ρ_{Λ} . While BAO scaling leaves Ω_{m} unaltered, the CMB degeneracy direction has Ω_{m} decreasing with increasing N_{eff} , as we discuss in §VII D 2.

Since changes to N_{eff} lead to changes in diffusion damping, polarization generation, and acoustic peak locations, the weighting of SPT data toward small angular scales and polarization spectra make them particularly interesting for this search. Varying N_{eff} in our analysis, we report

$$N_{\text{eff}} = 3.17^{+0.29}_{-0.33} \text{ for SPT-3G D1,} \quad (65)$$

$$N_{\text{eff}} = 2.77 \pm 0.17 \text{ for SPT+ACT,} \quad (66)$$

$$N_{\text{eff}} = 2.81 \pm 0.12 \text{ for CMB-SPA.} \quad (67)$$

The SPT constraint is in excellent agreement with the standard model prediction of $N_{\text{eff}} = 3.044$ and the data allow for a $\sim 10\%$ determination of N_{eff} . This is weaker than the result from ACT ($N_{\text{eff}} = 2.58^{+0.22}_{-0.25}$). However, the SPT data carry considerable weight in the SPT+ACT combination, which yields a posterior of slightly tighter width to the *Planck* constraint, $N_{\text{eff}} = 2.86 \pm 0.19$ [34]. We report no evidence for a deviation from the standard model prediction of $N_{\text{eff}} = 3.044$ for the SPT+ACT and CMB-SPA combinations. Marginalized N_{eff} posteriors for these cases are shown in the top panel of Fig. 23.

The constraints above enforce consistency between N_{eff} and the helium abundance Y_{P} in the framework

of Big Bang Nucleosynthesis (BBN).³⁷ Y_{P} is relevant here, because the abundance of light relics affects the production of helium during BBN, and also because changes to Y_{P} lead to changes to the Thomson scattering rate through the electron density $n_{\text{e}}(z)$. For a fixed baryon density, increasing Y_{P} reduces $n_{\text{e}}(z)$. This is because more neutrons are captured in helium atoms instead of decaying to protons and electrons during BBN, and because helium recombines earlier than hydrogen, thus decreasing the electron density at the time of hydrogen recombination. Freely varying Y_{P} means that the ratio of the Thomson and Hubble rates can be preserved, even while $H(z)$ is scaling up with N_{eff} . This leads to a partial degeneracy between Y_{P} and N_{eff} for CMB data.

We now vary N_{eff} and Y_{P} simultaneously in our analysis and show the resulting constraints in Fig. 23. We report 68% confidence intervals for the two parameters:

$$\left. \begin{array}{l} N_{\text{eff}} = 2.97^{+0.40}_{-0.64} \\ Y_{\text{P}} = 0.269^{+0.040}_{-0.030} \end{array} \right\} \text{ for SPT-3G D1,} \quad (68)$$

$$\left. \begin{array}{l} N_{\text{eff}} = 2.85^{+0.32}_{-0.40} \\ Y_{\text{P}} = 0.236^{+0.025}_{-0.021} \end{array} \right\} \text{ for SPT+ACT,} \quad (69)$$

$$\left. \begin{array}{l} N_{\text{eff}} = 2.99^{+0.22}_{-0.26} \\ Y_{\text{P}} = 0.231 \pm 0.014 \end{array} \right\} \text{ for CMB-SPA.} \quad (70)$$

All of the N_{eff} results match the standard model prediction of 3.044. The helium fraction constraints are consistent with the predictions for each data sets under BBN (using the calculations of [131] or of [132]) as well as the $Y_{\text{P}} = 0.2453 \pm 0.0034$ determination of [130], which is based on observations of metal-poor galaxies. Lastly, varying only Y_{P} while keeping N_{eff} fixed to the Λ CDM value also shows no deviation from the standard model. Therefore, we conclude that we see no evidence for additional light relics in the early universe in CMB

³⁶ where σ_{T} is the Thomson scattering cross-section and $n_{\text{e}}(z)$ the electron density.

³⁷ We use the BBN tables based on [131], which are the default in the CLASS version used in our analyses. To be precise, we define Y_{P} as the helium *mass* fraction.

data. All results of this section are given in Table V. We revisit these models adding BAO data in §VII D 2.

5. Modified recombination

In the final extension to Λ CDM explored with CMB data alone, we reconstruct the free electron fraction, $X_e(z) = n_e(z)/(n_{\text{HI}}(z) + n_{\text{HII}}(z))$, during the epoch of cosmological recombination, approximately spanning redshifts $500 < z < 1600$. In modified recombination scenarios, a new degeneracy between H_0 and $X_e(z)$ emerges which preserves the angular scale of the sound horizon at last scattering, $\theta_s^* = r_s^*/D_A^*$. As such, modifications to the standard recombination scenario have enjoyed recent attention due to their potential role in resolving the Hubble tension [e.g. 146–150]. To explore these models in a non-parametric way, we use the *ModRec* model and emulator introduced in [149], which uses seven free parameters to define the deviation of $X_e(z)$ from the standard model prescription.³⁸ For MCMC analyses of this model, we use the same priors as [149]. We restrict ACT data to $\ell < 4000$, as the accuracy of the *ModRec* emulator has not been validated at higher multipoles. We do not include lensing data due to compatibility issues with the *ModRec* emulator; since this model does not substantially change lensing predictions, we expect this to be a small effect (see Fig. 12 in [41]). For this model, we consider chains to be sufficiently converged when the Gelman-Rubin statistic is $R - 1 < 0.03$.

Whereas we have previously compared the width of individual parameter posteriors across CMB data sets to assess constraining power, this is not informative in this model as we have seven considerably correlated extension parameters. Therefore, we use the volume reduction (VR) statistic to quantify the information data sets hold in this model space introduced in §VII B 2. Here, we report natural logarithms of the square-root of determinant ratios and report values relative to the *Planck* T&E case, i.e.

$$\log \text{VR} = \log(\det(\Sigma_{\text{Planck}})/\det(\Sigma)). \quad (71)$$

For Gaussian posteriors, this quantity gives the number of e-folds by which the posterior volume is reduced compared to the *Planck* posterior. Additionally, to assess the significance of the deviation from Λ CDM, we translate the goodness-of-fit improvement compared to the standard model to the equivalent significance for a one dimensional Gaussian distribution, accounting for the seven additional degrees of freedom introduced.

Results from CMB data are in excellent agreement with the standard model prediction: for SPT+*Planck*

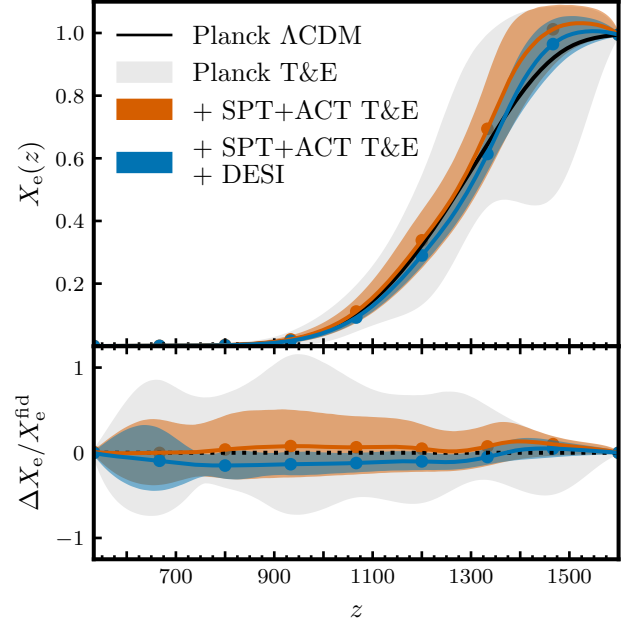


FIG. 24. *Top*: Reconstructions of the ionization fraction $X_e(z)$ during the epoch of recombination, using the *ModRec* model. Solid lines show the mean reconstructed ionization fraction, with points indicating the placement of the seven control points. Bands indicate the 95% confidence intervals. The black line is the fiducial ionization fraction for standard recombination with *Planck* cosmology [1]. *Bottom*: The fractional change relative to a fiducial ionization fraction calculated using the *Planck* T&E mean parameters [1]. Note that nodes can be significantly correlated and hence the significance of deviations from zero is difficult to judge by eye. The reconstructions based on CMB data are in excellent agreement with the standard model prediction; adding DESI BAO data leads to a mild 2.0σ preference for earlier recombination as discussed in §VII D 3.

data, the mean reconstructed ionization fraction matches the standard scenario within 0.18σ . The addition of SPT to *Planck* data significantly reduces uncertainties in the reconstruction of $X_e(z)$:

$$\log \text{VR} = 5.9 \text{ for } \textit{Planck} \text{ T\&E} + \text{SPT-3G D1 T\&E} \quad (72)$$

This reduction in posterior volume of almost 6 e-foldings is due to the inclusion of low-noise measurements of the EE and TE damping tail at $\ell > 1800$, which contain information on $X_e(z)$ at early times [149]. For comparison, we report the reduction of $\log \text{VR} = 7.1$ for the case when ACT data are added to *Planck* instead. Adding both SPT and ACT data to *Planck* reduces the reference posterior volume by about an additional e-fold:

$$\log \text{VR} = 8.3 \text{ for CMB-SPA T\&E.} \quad (73)$$

The significance of the deviation from the standard recombination scenario is 1.1σ in this case and we show this reconstruction in Fig. 24. Overall, we find no

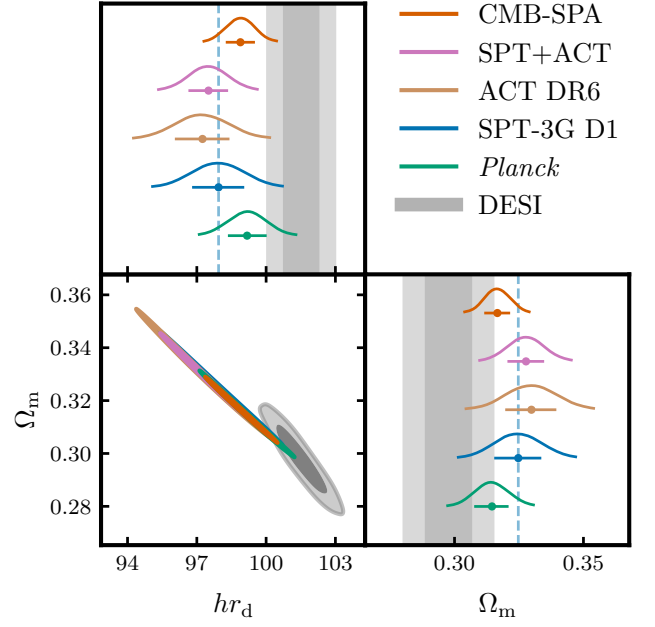
³⁸ Non-parametric here means that no particular functional form is imposed on $X_e(z)$. However, we are still using a parametric cosmological model, which the seven $X_e(z)$ parameters are now part of.

evidence for modified recombination from CMB data alone; given the volume reduction from *Planck*, this is a non-trivial validation of the standard recombination model on the new ground-based data. We add BAO data to this analysis in §VII D 3.

C. Evaluating the consistency of CMB and BAO data in Λ CDM

We now discuss the consistency of CMB and BAO measurements in the Λ CDM model. The state-of-the-art BAO data set is the DESI Data Release 2 (DR2) [3]. Under Λ CDM, BAO data constrain Ω_m and hr_d and the DESI data prefer a lower matter density and a higher hr_d than CMB data [151, 152].³⁹ In the discussion below, we translate the differences in the Ω_m and hr_d constraints from CMB and DESI data to equivalent statistical significances for a one-dimensional Gaussian distribution; we consider data sets consistent if they agree to better than 3σ according to this metric and allow ourselves to combine them.

We compare Λ CDM constraints from DESI and CMB data in Fig. 25. Using SPT data we find $\Omega_m = 0.3246 \pm 0.0090$ and $hr_d = 98.0 \pm 1.1$ Mpc. Accounting for the correlations between the two parameters, the difference with DESI translates to a one-dimensional Gaussian fluctuation of 2.5σ . This mild discrepancy becomes stronger when more CMB data are added. Among CMB data sets, ACT data prefer the highest Ω_m , and the tension between ACT and DESI is at the 3.1σ level.⁴⁰ We note that the ACT results used here include primary T&E CMB data as well as CMB lensing from ACT DR6 alone. This is different than what was done in [152], where the “ACT” data combination also contained *Planck* lensing, and also used a different prior on τ_{reio} , reducing the difference with DESI to the 2.7σ level.⁴¹ A combination of ground-based experiments into SPT+ACT yields tighter error bars. This increases the distance with DESI to 3.7σ . Similarly, the joint SPT+*Planck* constraints are different than DESI at 3.0σ due to the constraining power of the combined CMB data sets, even though the SPT-3G D1 and *Planck* constraints individually lie below the 3σ threshold. However, the *Planck* large angular scale data that are added going from SPT+ACT to CMB-SPA favor a lower Ω_m [1], and we report a difference with DESI of 2.8σ for CMB-SPA.



	$100 \Omega_m$	hr_d [Mpc]	Distance to DESI
CMB-SPA	31.66 ± 0.50	98.89 ± 0.63	2.8σ
SPT+ACT	32.77 ± 0.72	97.51 ± 0.87	3.7σ
SPT+ <i>Planck</i>	31.89 ± 0.54	98.63 ± 0.67	3.0σ
ACT DR6	33.0 ± 1.0	97.2 ± 1.2	3.1σ
SPT-3G D1	32.47 ± 0.91	97.9 ± 1.1	2.5σ
<i>Planck</i>	31.45 ± 0.67	99.18 ± 0.84	2.0σ
DESI	29.76 ± 0.87	101.52 ± 0.73	

FIG. 25. *Top*: Comparison of Λ CDM constraints from DESI and CMB data in the Ω_m - hr_d plane. *Bottom*: Mean and 68% confidence intervals, as well as significance of the discrepancy, ranging between 2.0 - 3.7σ for CMB data sets and particularly 2.8σ for the CMB-SPA combination.

We report joint constraints on Λ CDM parameters from SPT-3G D1 and CMB-SPA with DESI data in Table VI. The addition of the BAO data tightens constraints on some parameters, yielding notably $H_0 = 68.21 \pm 0.31 \text{ km s}^{-1} \text{ Mpc}^{-1}$ and $100 \Omega_c h^2 = 11.749 \pm 0.079$ for SPT+DESI. We forego reporting the combination of DESI with SPT+ACT as it does not meet our 3σ consistency requirement;⁴² if we were to do so, differences in the favored Ω_m and hr_d values would also lead to sizeable shifts in other cosmological parameters in the joint constraints compared to the CMB-preferred values. Due to the degeneracies of the model space,

³⁹ Tang et al. [153] show that if mock SDSS BAO data are generated assuming a dynamical dark energy model, a naive analysis in Λ CDM can lead to non-negligible biases in Ω_m . As DESI data greatly improve on SDSS data, more work is needed to understand to what extent this effect may apply to DESI data as well.

⁴⁰ We show the role that the τ prior plays in the consistency between ACT and DESI in Appendix I and Fig. 43.

⁴¹ We reproduce this result when using the same data combination, see Appendix I.

⁴² Though SPT+ACT is a subset of CMB-SPA, given the consistency of CMB data we have demonstrated in §VII B 2, we expect that the addition of large scale *Planck* data pushes the joint CMB constraints closer to the underlying mean. As the difference between CMB-SPA and DESI is below our 3σ threshold, we report the joint results.

Parameter	SPT-3G D1 + DESI	CMB-SPA + DESI
<i>Sampled</i>		
$10^4 \theta_s^*$	104.227 ± 0.056	104.180 ± 0.022
$100 \Omega_b h^2$	2.218 ± 0.022	2.2452 ± 0.0089
$100 \Omega_c h^2$	11.749 ± 0.079	11.813 ± 0.058
n_s	0.949 ± 0.012	0.9728 ± 0.0027
$\log(10^{10} A_s)$	3.066 ± 0.014	3.0574 ± 0.0094
τ_{reio}	0.0559 ± 0.0056	0.0625 ± 0.0050
<i>Derived</i>		
$H_0 [\text{km s}^{-1} \text{Mpc}^{-1}]$	68.21 ± 0.31	68.06 ± 0.24
Age [Gyr]	13.795 ± 0.025	13.783 ± 0.012
$10^9 A_s e^{-2\tau_{\text{reio}}}$	1.920 ± 0.021	1.8773 ± 0.0055
Ω_Λ	0.6983 ± 0.0039	0.6950 ± 0.0033
Ω_m	0.3017 ± 0.0039	0.3049 ± 0.0033
$r_d [\text{Mpc}]$	147.99 ± 0.33	147.51 ± 0.17
σ_8	0.8079 ± 0.0059	0.8120 ± 0.0038

TABLE VI. Joint Λ CDM parameters constraints from SPT-3G D1 and CMB-SPA with DESI. We report mean values and 68% confidence intervals.

H_0 , σ_8 , and n_s are particularly vulnerable to this effect. While special attention is often given to the first two parameters as they can be accurately determined by different cosmological probes (see §VII B 2), the precise value of n_s has consequences for the allowed model space of inflationary theories [154–156]; this compels caution in the interpretation of CMB+DESI results in Λ CDM.

As one would expect, the previous data release from DESI finds Ω_m and hr_d values similar to DESI DR2, though with larger error bars which reduces the significance of the discrepancy with CMB data [3, 157, 158]. DESI DR2 is also consistent with SDSS BAO data [3, 106–109]. While the SDSS data support values of Ω_m and hr_d that are more consistent with CMB data in Λ CDM, the data are also far less constraining than the DESI data, which prevents a meaningful comparison.

It is worth highlighting that the Λ CDM parameter constraints from CMB data presented in this work hinge on a determination of the optical depth to reionization τ_{reio} from *Planck* large scale polarization data [34]. Crucially τ_{reio} and Ω_m are anticorrelated for CMB constraints, and it has been noted that raising $\tau_{\text{reio}} \approx 0.09$ would bring Λ CDM predictions from CMB data and DESI into better agreement and regularize neutrino mass constraints [see §VIID 6, 159–162]. However, such a high value of τ_{reio} is not supported by the *Planck* E-mode data and there is no known significant systematic contamination in the measurement at this level. The prior-free estimation of τ_{reio} reported at the end of §VII B 2 is consistent both with our *Planck*-based prior, but also within 1σ of $\tau_{\text{reio}} = 0.09$. Together with the fact that the *Planck* large-scale polarization data are not sample variance limited, this motivates revisiting τ_{reio} using E-mode measurements from, e.g., the CLASS

telescope [163–165] or the LiteBIRD mission [166, 167].

We conclude that there is a growing discrepancy in the Λ CDM parameter values favored by CMB and BAO data, even though Λ CDM provides an excellent fit to CMB and BAO data separately. The magnitude of this effect is borderline statistically significant, which motivates the collection of more precise data as well as the search for a cosmological model that better fits both types of data simultaneously.

D. Constraints from CMB and BAO data on extended cosmological models

In this section, we combine DESI and CMB data to constrain different extended (beyond- Λ CDM) cosmological models. The combination of BAO and CMB data has the potential to constrain extensions beyond just the combined statistical power of the two probes. BAO data are sensitive to the expansion rate and angular-diameter distances at low redshifts, breaking geometric degeneracies that are otherwise present when analyzing CMB data alone. Additionally, we are motivated by the aforementioned possibility of identifying a model that improves on the joint description of CMB and BAO data provided by Λ CDM. To justify a joint analysis, we require that the individual constraints from CMB and DESI data are consistent in a given model space. We use the same metric and threshold we used to assess consistency in Λ CDM in §VII C; the differences in constraints in the Ω_m - hr_d plane between the data sets translate to a one-dimensional Gaussian fluctuation of less than 3σ . For models where DESI data are sensitive to the extension parameters, we also include these in the comparison.

We revisit constraints on the lensing amplitude (§VIID 1), light relics (§VIID 2), and modified recombination (§VIID 3). We also report constraints on spatial curvature, alone (§VIID 4) and in addition to varying the electron mass (§VIID 5), the sum of neutrino masses (§VIID 6), and time-evolving dark energy (§VIID 7). We find shifts away from the CMB-preferred values and at times moderate fluctuations away from the standard model. We report best-fit χ^2 values for CMB-SPA and DESI in Table VII and show constraints for select models in Fig. 26. We highlight similarities among the different constraints throughout the text.

1. Lensing amplitude revisited

We begin by revisiting constraints on the consistency of the signature of gravitational lensing across cosmological probes. We repeat the analysis of §VII B 3, but now add DESI data. We stress again that the models considered here are non-physical.

We first consider the case of A_{lens} , i.e. the coherent modification of the amplitudes of lensing in the CMB primary power spectrum and the CMB lensing recon-

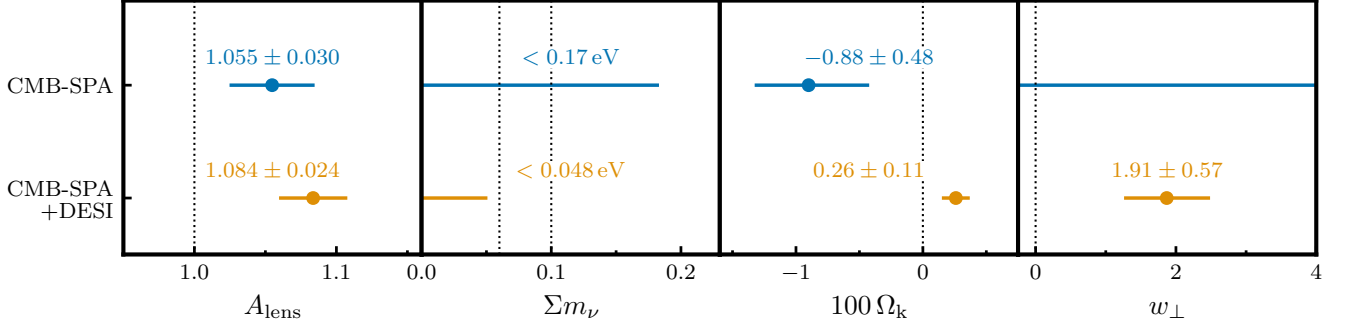


FIG. 26. Constraints on extension parameters from CMB-SPA (top, blue) and CMB-SPA+DESI (orange, bottom); from left to right: A_{lens} , Σm_ν , Ω_k , and w_\perp (which summarizes w_0 and w_a constraints as defined in Eq. (101)). We give the mean and 68% confidence interval above each constraint; for Σm_ν we indicate the 95% upper limit with arrows. We indicate the value of the extension parameter to which the model reverts in Λ CDM as the dotted vertical line; for Σm_ν we indicate the minimum values for the normal and inverted hierarchies (0.06 eV and 0.1 eV, respectively). We indicate the insensitivity of CMB data to the late-time evolution of dark energy by a horizontal blue line in the final panel. We find that differences between CMB and DESI data lead to sizeable shifts from the CMB-preferred values in joint constraints and moderate departures from Λ CDM. We see a similar behavior for modified recombination scenarios (§VII B 5, §VII D 3) (not shown in this figure).

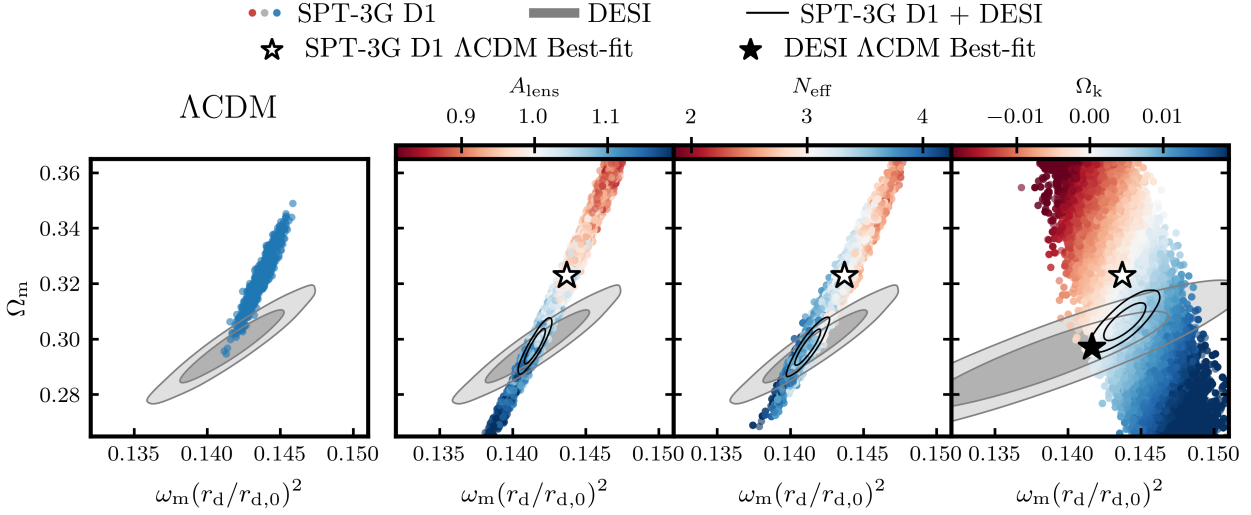


FIG. 27. Differences between CMB and DESI data project onto extension parameters in joint analyses. We show constraints in the $\Omega_m - \omega_m(r_d/r_{d,0})^2$ plane, where $r_{d,0} = 147$ Mpc, first in Λ CDM, then for extensions with free A_{lens} , N_{eff} , and Ω_k . In all panels, colored points represent samples from SPT-3G D1 chains, filled gray contours indicate DESI posteriors, and solid black line contours indicate the joint SPT-3G D1 + DESI posteriors. The black and white stars indicate the Λ CDM best-fit points of either probe individually, when these do not coincide with the centers of the contours or samples shown (SPT-3G D1 in white, DESI in black).

Far-left panel: SPT-3G D1 MCMC samples (blue dots) and DESI posteriors (gray contours) in Λ CDM. The SPT data favor higher values of Ω_m and $\omega_m r_d^2$ than the DESI data.

Center-left panel: same as far left but for a model with free CMB lensing amplitude (dots colored according to A_{lens}). In this model space, a degeneracy between Ω_m and $\omega_m r_d^2$ at constant θ_s^* extends the SPT posterior from Λ CDM into the parameter region supported by DESI data; with values A_{lens} varying along the band. Since the DESI constraints intersect the SPT band at comparatively low Ω_m and $\omega_m r_d^2$ values, the joint analysis favors $A_{\text{lens}} > 1$.

Center-right panel: same as far left but for a model with free effective number of relativistic species (dots colored according to N_{eff}). For CMB data, lower Ω_m and $\omega_m r_d^2$ can, to a certain degree, be accommodated by raising N_{eff} , leading to a moderate positive shift in the best-fit N_{eff} value in a joint analysis with DESI data.

Far-right panel: same as far left but for a model with free mean spatial curvature (points colored according to Ω_k). CMB data constrain a band in the $\Omega_m - \omega_m r_d^2$ plane; lines of constant Ω_k dissect the band diagonally, along the Λ CDM degeneracy direction. As in the other extension models, the DESI contour intersects the SPT band away from the Λ CDM value of the extension parameter, in this case leading to a mild preference for an open universe in the joint constraints.

Model	CMB-SPA	DESI	CMB-SPA+DESI		
	χ^2_{CMB}	χ^2_{DESI}	χ^2_{CMB}	χ^2_{DESI}	$\chi^2_{\text{CMB+DESI}}$
ΛCDM	1550.9	10.3	1556.0	14.8	1570.7
A_{lens}	1548.9 (2.0, 1.4 σ)	-	1550.4	10.9	1561.2 (9.5, 3.1 σ)
ModRec	(8.9, 1.1 σ)	-	(12.0)	(2.2)	(14.2, 2.0 σ)
Ω_k	1549.5 (1.4, 1.2 σ)	10.0 (0.3, 0.6 σ)	1553.5	10.9	1564.4 (6.3, 2.5 σ)
$\Omega_k + m_e$	-	-	1553.6	10.3	1563.9 (6.8, 2.1 σ)
Σm_ν	1551.0 (-0.1, 0.0 σ)	-	1551.2	11.8	1562.9 (7.8, 2.8 σ)
$w_0 w_a$	-	5.6 (4.7, 1.7 σ)	1550.0	7.3	1557.3 (13.5, 3.2 σ)

TABLE VII. χ^2 values of CMB-SPA and DESI for different cosmological models (first column) at their individual best-fit points (second and third columns, respectively) and for a joint analysis. For the joint fit, we list the total χ^2 value (last column) and the contributions from the CMB and DESI likelihoods (fourth and fifth columns, respectively). In parentheses, we report the χ^2 improvement with respect to the relevant ΛCDM reference case (first row) and translate this number to an equivalent frequentist significance for a one-dimensional Gaussian distribution. As we do not use lensing data when constraining modified recombination scenarios (see §VII B 5), the absolute χ^2 values are not directly comparable with the rest of the table; we instead only report differences with a corresponding ΛCDM analysis.

struction, as defined in §VII B 3. Adding DESI to CMB data changes the A_{lens} constraint from Eq. (61) to

$$A_{\text{lens}} = 1.084 \pm 0.035 \text{ for SPT-3G D1 + DESI,} \quad (74)$$

$$A_{\text{lens}} = 1.092 \pm 0.026 \text{ for SPT+ACT + DESI,} \quad (75)$$

$$A_{\text{lens}} = 1.084 \pm 0.024 \text{ for CMB-SPA + DESI.} \quad (76)$$

which are deviations from the standard model prediction of 2.4 σ , 3.5 σ , and 3.5 σ , respectively. We note that in moving from SPT+ACT to CMB-SPA we only include *Planck* data that by themselves do not prefer an anomalous A_{lens} (in particular by limiting *Planck* TT data to $\ell < 1000$). We find similar results when allowing $A_{2\text{pt}}$ and A_{recon} to vary independently from one another, e.g. for SPT+ACT+DESI we report a deviation from ΛCDM of 3.1 σ (see also [16]).

When considering only CMB data, allowing A_{lens} to vary does not improve the goodness-of-fit for ground based data, as explored in §VII B 3. The joint CMB-SPA+DESI fit improves on the ΛCDM minimum χ^2 value by 9.5 points; the fit to CMB data is improved by 5.6 points and the fit to BAO data by 3.9 points (see Table VII). This is a non-negligible improvement given the introduction of one additional parameter. Translating the quality-of-fit improvement to an equivalent one-dimensional Gaussian significance, this corresponds to a 3.1 σ event.

This result is a projection of differences in the CMB and DESI data that can be understood in the Ω_m - $\omega_m r_d^2$ plane, where $\omega_m \equiv \Omega_m h^2$ (for a more complete discussion, see [168]). The precise determination of θ_s^* by CMB data translates to a thin contour in this plane. When A_{lens} is allowed to vary, the contour is extended into a narrow band, along which the extension parameter varies. The extended contour meets the DESI constraint at comparatively low values of Ω_m and $\omega_m r_d^2$, which necessitates $A_{\text{lens}} > 1$ in order to not degrade the fit to the CMB data. We illustrate this effect in Fig. 27. Physically, the lower matter density preferred

by DESI would imply less gravitational lensing. This is at odds with the amplitude of the effect in CMB data (see §VII B 3) and hence A_{lens} is raised to compensate. As such, the deviation from $A_{\text{lens}} = 1$ in joint constraints with DESI is a rephrasing of the marginal agreement of the two probes in ΛCDM .

2. New light particles revisited

Next, we revisit constraints on light relics, now adding DESI to the CMB data. As discussed in §VII B 4, varying N_{eff} from its standard model prediction opens up the BAO scaling transform, under which BAO observables are effectively invariant [144].⁴³ In contrast, CMB observables change as mentioned in §VII B 4 and as discussed in [144], with the result that CMB data can be best fit if the matter density scales up more slowly than the radiation density. The different responses of the two probes to BAO scaling make it interesting to study N_{eff} using the combination of the two.

In this model space, the constraints from SPT+ACT and CMB-SPA on Ω_m and $h r_d$ are discrepant with those from DESI at more than 3 σ and hence do not meet our requirement for joint analyses (see the start of §VII D). This is also true for the cases of Y_p and $N_{\text{eff}} + Y_p$. Hence, we only report results for SPT-3G D1 in combination with DESI data.

We first allow for N_{eff} to vary in our analysis while maintaining BBN consistency. Adding DESI to

⁴³ Assuming a fixed redshift for the end of the baryon drag epoch, z_d , BAO observables remain exactly invariant. Even including the associated changes to z_d , the locations of the predicted BAO correlation peaks change by $< 0.2\%$ for the extreme case of $N_{\text{eff}} = 5$, which, as we saw in §VII B 4, is ruled out by CMB data.

SPT-3G D1, we report:

$$N_{\text{eff}} = 3.52 \pm 0.23 \text{ for SPT-3G D1 + DESI.} \quad (77)$$

The inclusion of DESI data shifts the central value of the posterior up by 1σ and we report a mild preference for $N_{\text{eff}} > 3.044$ at 2.1σ . The N_{eff} posterior tightens by 20% compared to the SPT CMB-only constraint.

Varying N_{eff} and Y_{P} simulatenously, we report:

$$\left. \begin{aligned} N_{\text{eff}} &= 3.64 \pm 0.42 \\ Y_{\text{P}} &= 0.241 \pm 0.034 \end{aligned} \right\} \text{ for SPT-3G D1 + DESI.} \quad (78)$$

Again, the inclusion of DESI data tightens the N_{eff} posterior and leads to mild $\sim 1\sigma$ shifts up in N_{eff} and down in Y_{P} from the CMB-preferred values as these two extension parameters are anticorrelated. The N_{eff} value is within 1.4σ of the Λ CDM value and the Y_{P} constraint matches the BBN predictions of [131] and [132] and the measurement of [130] at $\leq 0.2\sigma$. We find similar results when allowing only Y_{P} to vary while fixing $N_{\text{eff}} = 3.044$ (see Table V).

As we illustrate in Fig. 27 for the case of only varying N_{eff} , at the level of parameter degeneracies this result can be understood in a similar way to the A_{lens} case above. Again, the $\Omega_{\text{m}}\text{-}\omega_{\text{m}}r_{\text{d}}^2$ posterior of SPT-3G D1 is extended along its existing degeneracy direction compared to the Λ CDM case; N_{eff} varies along this direction, with values higher than the standard model prediction being supported at lower Ω_{m} and $\omega_{\text{m}}r_{\text{d}}^2$ values, as explained in §VII B 4. This is where the DESI posterior lies in the $\Omega_{\text{m}}\text{-}\omega_{\text{m}}r_{\text{d}}^2$ plane and hence the inclusion of the BAO data pulls the joint constraints towards higher N_{eff} compared to the values preferred by SPT-3G D1. The picture is the same when opening up the helium abundance; we find a shift down in Y_{P} as this parameter is anticorrelated with N_{eff} , but positively correlated with Ω_{m} and $\omega_{\text{m}}r_{\text{d}}^2$.

3. Modified recombination revisited

We examine the impact of DESI data on the previously presented reconstructions of the ionization fraction X_{e} (see §VII B 5). We find that the inclusion of DESI further significantly improves the reconstruction, with:

$$\log \text{VR} = 12.3 \text{ for CMB-SPA T\&E + DESI.} \quad (79)$$

This is a reduction by about four e-foldings compared to the CMB-SPA reconstruction and is due to the fact that the low-redshift measurements of the expansion history provided by BAO break the $H_0\text{-}X_{\text{e}}$ degeneracy that is present when only using CMB data.

Along with sharpening the posterior, we find that the inclusion of DESI data leads to a mild 2.0σ preference for earlier recombination as shown in Fig. 24. This goes

hand-in-hand with higher values of H_0 :

$$H_0 = 69.34 \pm 0.70 \text{ for Planck + SPT T\&E + DESI} \quad (80)$$

$$H_0 = 69.48 \pm 0.65 \text{ for CMB-SPA T\&E + DESI} \quad (81)$$

This result is in line with previous work that has found that DESI data lead to a preference for higher H_0 values in modified recombination scenarios when analyzed alongside CMB data [7, 150, 169–171]. However, differences between the above H_0 constraints and local measurements from SH0ES remain $> 3\sigma$. There is no statistically significant preference for this model over Λ CDM for CMB-SPA+DESI. Though the best-fit χ^2 value is reduced by 14.2 points, when accounting for the seven additional degrees of freedom this only translates to a mild 2σ preference for modified recombination (see Table VII). Still, a lower-dimensional model able to reproduce the essential features of ModRec could in principle lead to a comparable goodness-of-fit improvement and an increased significance.⁴⁴

Our findings above are again a result of features in the DESI data that project onto Ω_{m} and hr_{d} . Through earlier recombination, the ModRec model improves consistency between BAO and CMB data; this allows the model to adjust to higher H_0 values [see Fig. 2 in 170], while keeping the ratio $\theta_{\text{s}}^* = r_{\text{s}}^*/D_{\text{A}}^*$ consistent with CMB allowed values.

4. Spatial curvature

Inflation is expected to reduce primordial spatial curvature to levels well below current experimental sensitivity. A deviation from this prediction would constitute a major challenge for the standard model of cosmology. Though CMB data are sensitive to spatial curvature by themselves, they suffer from geometric degeneracies that make constraints rather weak.

For CMB alone, we find

$$100\Omega_{\text{k}} = 0.2^{+1.5}_{-1.2} \text{ for SPT-3G D1,} \quad (82)$$

$$100\Omega_{\text{k}} = -0.06^{+0.81}_{-0.70} \text{ for SPT+ACT,} \quad (83)$$

$$100\Omega_{\text{k}} = -0.88 \pm 0.48 \text{ for CMB-SPA.} \quad (84)$$

Though loose, these constraints are all compatible with a flat universe at $< 2\sigma$.

The addition of BAO data breaks the limiting degeneracies and tightens the posteriors by about a factor of four:

$$100\Omega_{\text{k}} = 0.40 \pm 0.18 \text{ for SPT-3G D1 + DESI,} \quad (86)$$

$$100\Omega_{\text{k}} = 0.51 \pm 0.17 \text{ for SPT+ACT + DESI,} \quad (87)$$

$$100\Omega_{\text{k}} = 0.26 \pm 0.11 \text{ for CMB-SPA + DESI.} \quad (88)$$

⁴⁴ Common parametrizations of primordial magnetic fields are of interest in this context [see 172–176].

The SPT-3G D1+DESI, SPT+ACT+DESI, and CMB-SPA+DESI constraints are 2.2σ , 3.0σ , and 2.4σ away from spatial flatness, respectively. The above deviations are reflected in moderately improved best-fit χ^2 values for joint CMB+DESI analyses as listed in Table VII. For CMB-SPA+DESI we report a reduction by 6.3 points compared to Λ CDM, which corresponds to a 2.5σ preference.

At the parameter level, this deviation from Λ CDM can be understood similarly to the A_{lens} and N_{eff} cases above, though when allowing for non-zero spatial curvature, the degeneracy direction between Ω_m and $\omega_m r_d^2$ flips compared to Λ CDM and the parameters are now anticorrelated. However, the data constrain a relatively thick band in this plane across which lines of constant Ω_k trace the original degeneracy direction of Ω_m and $\omega_m r_d^2$ present in Λ CDM (as well as in A_{lens} and N_{eff}). Hence, CMB data can support lower Ω_m without substantially modifying $\omega_m r_d^2$ if $\Omega_k > 0$. This is illustrated in the right-most panel of Fig. 27. Note that a positive Ω_k is at odds with the excess lensing preferred by the high- ℓ *Planck* TT spectrum, which tends to push Ω_k to negative values and Ω_m to higher values (the opposite of the low Ω_m preferred by DESI). Thus, it is possible that replacing the ACT DR6 high- ℓ TT used in the CMB-SPA+DESI combination with the one from *Planck* could weaken the preference for a $\Omega_k > 0$. For a more detailed discussion of CMB and BAO constraints in this model space, we point the reader to [177].

5. Spatial curvature and a varying electron mass

We now consider constraints when allowing for a varying electron mass in a non-flat geometry, $m_e + \Omega_k$. In the review of potential Hubble tension solutions by [178], this model was one of the most promising, and we re-evaluate it in the context of the most up-to-date and stringent CMB and BAO data here. We take the result of [2] from the SH0ES collaboration as the reference H_0 measurement for this section.

From a theoretical standpoint, scalar fields predicted by fundamental theories could couple to elementary particles, specifically electrons [179, 180], which could result in electrons having an effective mass at high redshift different than that measured in laboratory experiments today. Increasing the electron mass causes recombination to occur earlier, which in turn decreases the size of the sound horizon at that epoch. Moreover, allowing the spatial curvature to vary accommodates this change in a way that fits late-universe measurements, specifically BAO data. Together, these two effects allow for an increase in H_0 [178, 181, 182]. We model the transition of the electron mass as a step function at $z \sim 50$ [183, 184]; this is the simplest implementation currently available in CLASS and is sufficient for the precision of current cosmological data. We allow for the value at high redshift, m_e , to deviate from today's value,

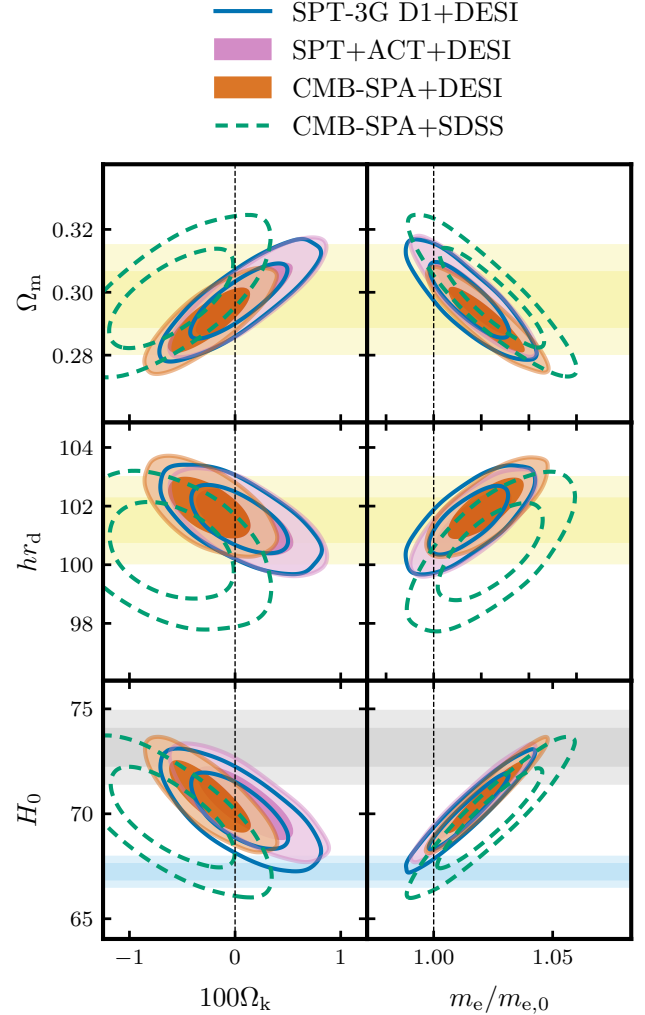


FIG. 28. Constraints on $m_e + \Omega_k$ from SPT-3G D1 + DESI (blue, empty), SPT+ACT + DESI (purple, filled), and CMB-SPA + DESI (orange, filled). In the two top rows, the yellow bands correspond to the 68% and 95% confidence levels of DESI assuming Λ CDM. In the bottom row, the light blue bands correspond to the 68% and 95% confidence levels for CMB-SPA in Λ CDM, while the gray bands indicate the SH0ES measurement [2]. This model is able to considerably reduce the Hubble tension (from $Q_{\text{MPCL}} = 5.9\sigma$ to 1.9σ , see text for details). We show the constraints when replacing DESI with SDSS BAO data as dashed green line contours.

$m_{e,0}$, such that $m_e/m_{e,0} \in (0.1, 2)$. Due to the heavy numerical cost of this model, we loosen our convergence requirement to $R - 1 < 0.04$.

In Fig. 28, we show joint constraints from CMB and BAO data on key parameters.⁴⁵ In this model space,

⁴⁵ We do not report CMB-only constraints as CMB data alone yield large parameter uncertainties in this model, rendering results inconclusive [178, 185].

Parameters	SPT-3G D1 +DESI	SPT+ACT +DESI	CMB-SPA +DESI
$H_0 [\text{km s}^{-1} \text{ Mpc}^{-1}]$	70.1 ± 1.2	70.5 ± 1.2	70.6 ± 1.1
$100 \Omega_m$	29.74 ± 0.79	29.83 ± 0.78	29.20 ± 0.68
$hr_d [\text{Mpc}]$	102.5 ± 0.8	101.4 ± 0.7	101.9 ± 0.7
$100 \left(\frac{m_e}{m_{e,0}} - 1 \right)$	1.5 ± 1.1	1.6 ± 1.1	1.97 ± 0.98
$100 \Omega_k$	0.04 ± 0.31	0.10 ± 0.30	-0.21 ± 0.25
$Q_{\text{MPCL}}^{\text{SH0ES}} [\sigma]$	2.3	2.0	1.9

TABLE VIII. Cosmological parameters for the $m_e + \Omega_k$ model for SPT-3G D1, SPT+ACT, and CMB-SPA data sets combined with DESI BAO data. We also report the Marginalized Posterior Compatibility Level, $Q_{\text{MPCL}}^{\text{SH0ES}}$, which quantifies the tension with SH0ES data [2].

current CMB data and DESI constrain H_0 to be:

$$H_0 = 70.1 \pm 1.2 \text{ km s}^{-1} \text{ Mpc}^{-1} \text{ for SPT-3G D1 + DESI,} \quad (89)$$

$$H_0 = 70.5 \pm 1.2 \text{ km s}^{-1} \text{ Mpc}^{-1} \text{ for SPT+ACT + DESI,} \quad (90)$$

$$H_0 = 70.6 \pm 1.1 \text{ km s}^{-1} \text{ Mpc}^{-1} \text{ for CMB-SPA + DESI.} \quad (91)$$

Due to the degeneracies that $(m_e/m_{e,0}, \Omega_k)$ exhibit with (Ω_m, hr_d) , raising H_0 brings Ω_m and hr_d in line with the DESI-preferred values in ΛCDM . Comparing the best-fit χ^2 value of CMB-SPA+DESI in this model to ΛCDM yields an improvement by 6.8 points, which given the introduction of two new free parameters, corresponds to a mild 2.1σ preference (see Table VII). Full parameter results can be found in Table VIII. We note that when replacing DESI data with SDSS BAO measurements, similarly high values of H_0 can be achieved, as shown in Fig. 28.

To assess to what degree this model reduces the tension between CMB+BAO and SH0ES data, we use two metrics. First, we calculate the Marginalized Posterior Compatibility Level, Q_{MPCL} ,⁴⁶ introduced by [178] [see also 186, 187]. We report a reduction from 5.9σ in ΛCDM to 1.9σ in this model for CMB-SPA+DESI compared to SH0ES. Second, we report the Difference of the Maximum a Posteriori (DMAP) criterion, defined as:

$$Q_{\text{DMAP}} = \sqrt{\chi_{\mathcal{D}+\text{SH0ES}}^2 - \chi_{\mathcal{D}}^2}, \quad (92)$$

where the best-fit χ^2 of the model is evaluated with a given data set \mathcal{D} with and without the SH0ES likelihood.

⁴⁶ This statistic quantifies the agreement between two data sets, generalized to the case where the posterior distributions are not necessarily Gaussian. It computes consistency directly from the MCMC chains, instead of using parameter covariance matrices (see Section 4.3 of [178] for more details).

By comparing the two χ^2 values, the statistic indicates whether the model fit to the data worsens when including SH0ES information. For CMB-SPA+DESI, we find a reduction of the tension with SH0ES data from $Q_{\text{DMAP}} = 5.6\sigma$ in ΛCDM to 2.6σ for $\Omega_k + m_e$. The model is able to reduce the tension to below the 3σ threshold, which was not the case in [178] using the previous SPT-3G data from [14], *Planck*, Pantheon SNIa, and part of the SDSS BAO measurements. We conclude that while at current sensitivity the CMB and BAO data show no statistically significant preference for this model over ΛCDM , the model passes key tests in the comparison with the SH0ES H_0 measurement and may therefore still be considered as a possible solution to the Hubble tension.

6. Neutrino cosmology

From neutrino oscillation experiments, we know that at least two neutrinos have a non-zero mass, which implies that the sum of neutrino masses is either larger than 0.06 eV or 0.1 eV assuming a normal or inverted mass hierarchy, respectively [188, 189]. As shown in many works, massive neutrinos hinder the formation of structure, leading to many observable effects, including a reduction in the amplitude of deflections due to gravitational lensing [190]. As [191] illustrate, CMB and BAO data have opposing degeneracies in a parameter space in which neutrino mass is a free parameter, which motivates a joint analysis. In this section, we fit jointly to CMB and DESI data allowing for arbitrary, positive Σm_ν .

We remind the reader that the CMB-SPA combination removes *Planck* data at $\ell > 1000$ to avoid correlations with ACT data. Because of the known preference of *Planck* TT data for a higher lensing amplitude (see §VII B 3), if we were to use the *Planck* data at those multipoles instead of ACT in that combination, we would likely find an even tighter upper limit on Σm_ν . Hence, for this model, our construction of the CMB-SPA data set is conservative. As a point of comparison, we first report 95% confidence upper limits for the CMB data alone of:

$$\Sigma m_\nu < 0.77 \text{ eV for SPT-3G D1,} \quad (93)$$

$$\Sigma m_\nu < 0.58 \text{ eV for SPT+ACT,} \quad (94)$$

$$\Sigma m_\nu < 0.17 \text{ eV for CMB-SPA.} \quad (95)$$

Though these upper limits remain away from the minimum Σm_ν values allowed based on neutrino oscillation experiments, we note that the CMB-SPA constraint improves on the *Planck* one of $< 0.25 \text{ eV}$ by about 30%.

We now add DESI data. In this model space, the constraints from SPT+ACT and DESI on Ω_m and hr_d are discrepant at more than 3σ , exceeding our requirement for joint analyses, and we do not report results for this combination. As in the ΛCDM case

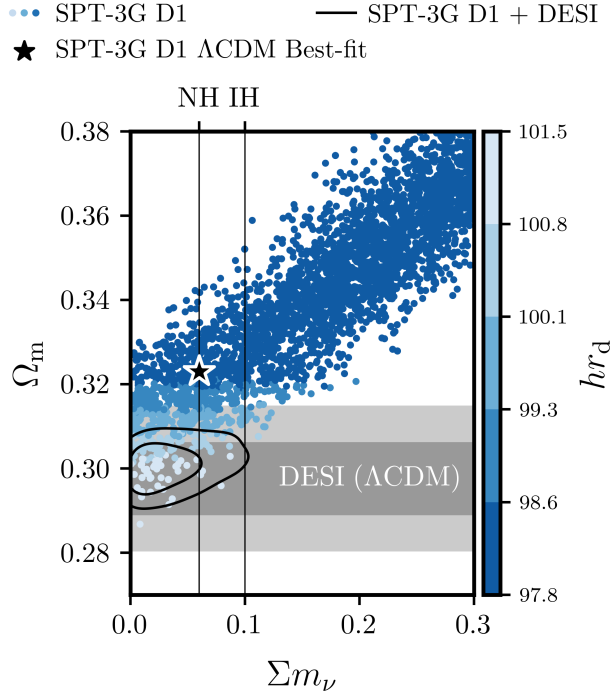


FIG. 29. Samples from an MCMC analysis of SPT-3G D1 constraining Λ CDM + Σm_ν (blue dots). The dot color corresponds to the hr_d value, where light blue is the mean for DESI data in Λ CDM ($hr_d = 101.5$ Mpc) and dark blue is $\geq 4\sigma$ low ($hr_d = 98.6$ Mpc). The horizontal gray bands indicate the Λ CDM DESI Ω_m constraint. As a reference, the black star shows the best-fit point for SPT in Λ CDM. We mark the lower limits for a normal (NH) and inverted (IH) neutrino hierarchy with black vertical lines. The differences between DESI and CMB data projected into Ω_m and hr_d lead to a preference for as small a value of Σm_ν as possible in joint constraints (black line contours).

(§VII C), the addition of *Planck* data to SPT+ACT regularizes the CMB constraint such that the differences between CMB-SPA and DESI are below our threshold. When adding DESI to SPT-3G D1 and CMB-SPA, we find at the 95% confidence level:

$$\Sigma m_\nu < 0.081 \text{ eV for SPT-3G D1 + DESI,} \quad (96)$$

$$\Sigma m_\nu < 0.048 \text{ eV for CMB-SPA + DESI.} \quad (97)$$

As expected, adding BAO data tightens the constraint substantially. While the upper limit derived from SPT data alone is consistent with neutrino oscillation data, with a posterior that peaks slightly away from zero, the CMB-SPA+DESI combination appears to rule out the normal and inverted hierarchies at 97.9% and 99.9% confidence, respectively. Moving Σm_ν close to zero reduces the best-fit χ^2 value by 7.8 points for joint CMB and BAO analyses compared to the minimal value for the normal hierarchy, which for one additional parameter corresponds to a 2.8σ significance (see Table VII).

The drive toward as low of a value for Σm_ν as allowed

in joint constraints (even negative Σm_ν values if the model is phenomenologically extended into this regime) is a known effect, and there exists a growing literature dissecting and contextualizing cosmological neutrino mass constraints [see e.g. 159, 160, 168, 191]. Increasing Σm_ν raises $H(z)$ during the matter-dominated epoch; to keep θ_s^* consistent with CMB data, the cosmological constant decreases and Ω_m increases. Though r_d does not change, lowering Λ decreases H_0 and hence hr_d and Σm_ν are anti-correlated [168, 192]. As DESI data prefer a low Ω_m and a high hr_d compared to CMB data, this forces the joint posterior against the $\Sigma m_\nu = 0$ boundary. This gives rise to the tight upper limits seen above and is illustrated in Fig. 29. These results are in growing discord with neutrino oscillation experiments; improved cosmological data sets will allow for more scrutiny.

7. Time-evolving dark energy

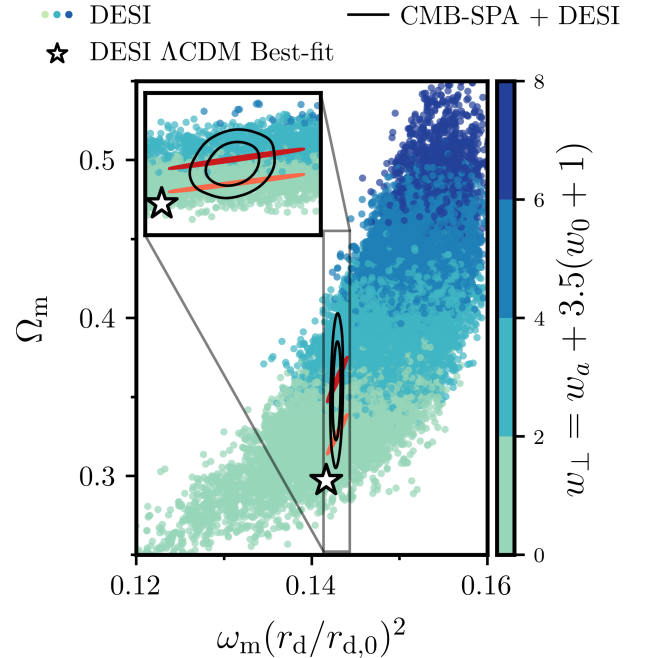


FIG. 30. DESI MCMC samples in the Ω_m - $\omega_m r_d^2$ plane for Λ CDM + $w_0 w_a$ (colored dots). The DESI constraints in this plane are substantially widened compared to the Λ CDM case. Samples are colored according to w_\perp (see Eq. (101)), which is highly correlated with Ω_m . The white star marks the DESI Λ CDM best-fit point for reference. We indicate the CMB information by showing the Λ CDM CMB-SPA 2σ contour (orange ellipse) and the CMB-SPA 2σ contour fixing w_\perp to the DESI mean value of $w_\perp = 3.1$ (red ellipse). By constraining $\omega_m r_d^2$, the CMB data break the degeneracies of the DESI data, which substantially tightens the joint w_\perp posterior (CMB-SPA+DESI shown as black line contours).

Lastly, we turn our attention to time-evolving dark

energy. Instead of assuming a cosmological constant model, we allow for the equation of state of dark energy to vary [193, 194] according to

$$w(z) = w_0 + w_a \frac{z}{1+z}. \quad (98)$$

This model has recently received attention as the combination of DESI data with CMB and uncalibrated SNe Ia data show a $\gtrsim 3\sigma$ preference for a deviation from Λ CDM [3, 157, 195].

Combining CMB-SPA and DESI data, we report

$$w_0 = -0.41 \pm 0.20, \quad (99)$$

$$w_a = -1.78 \pm 0.55. \quad (100)$$

In the w_0 - w_a plane, this is a 2.9σ deviation from $(w_0, w_a) = (-1, 0)$. Adding Pantheon+ uncalibrated SNe Ia data moves the constraint to $w_0 = -0.831 \pm 0.054$ and $w_a = -0.66 \pm 0.19$ and the w_0 - w_a central values remain 3σ from the Λ CDM prediction.⁴⁷

This deviation is reflected by improved best-fit χ^2 values (see Table VII). The goodness-of-fit improves by 13.5 points for CMB-SPA+DESI in this model compared to Λ CDM. This is close to the value reported by the DESI collaboration when using only *Planck* CMB data (12.5) [3] and equals to a 3.2σ event for a one-dimensional Gaussian distribution. The contributions from CMB and DESI data to the total χ^2 improve by 6.0 and 7.5 points compared to the standard model, respectively. We note that restricting the allowed parameter space or otherwise imposing priors in the w_0 - w_a plane to ensure the evolution of dark energy is physical tends to weaken the significance of departures from a cosmological constant [198, 199]. More generally, one can argue that our ignorance of the nature of dark energy makes it difficult to apply theoretically motivated priors on this model and hence particularly strong evidence from the data in favor of this model is needed [for a recent discussion in the context of DESI results, see 151]. Still, the sensitivity of BAO data to the late-time evolution of dark energy makes this model interesting and further data will help assess the robustness of the trends we are seeing.

To better understand the above result we introduce the summary parameter

$$w_\perp \equiv w_a + 3.5(w_0 + 1), \quad (101)$$

which maximally varies along the w_0 - w_a degeneracy direction of DESI data, such that $w_\perp = 0$ corresponds to $w_0, w_a = (-1, 0)$ and $w_\perp > 0$ to $w_0 > -1, w_a < 0$. As shown in Fig. 30, allowing for dynamical dark energy greatly relaxes the DESI constraints in the Ω_m - $\omega_m r_d^2$ plane (compare to the left-most panel of Fig. 27);

in this model, w_\perp is highly correlated with Ω_m . For DESI data, we report $w_\perp = 3.1 \pm 1.6$. CMB data, on the other hand, do not constrain w_\perp : they are not sensitive to the detailed time evolution of the dark energy equation of state at late times, but only to the integral effect on the angular diameter distance to the last scattering surface. However, CMB data still provide a tight constraint on $\omega_m r_d^2$ in this model and, therefore, break the degeneracies of the DESI data. The error bar on w_\perp shrinks substantially in joint constraints and we report $w_\perp = 1.91 \pm 0.57$ for CMB-SPA+DESI.

VIII. CONCLUSIONS

In this work, we have presented CMB temperature and E-mode polarization power spectrum measurements based on SPT-3G data and the constraints on cosmological parameters enabled by these measurements, both individually and in combination with other data sets. We used observations of the SPT-3G Main field at 95, 150, and 220 GHz collected in the austral winter seasons of 2019 and 2020.

The temperature and polarization maps made from these data, and the series of null tests used to demonstrate that these maps are free from significant systematic effects, are described in detail in Q25. These maps are the deepest high-resolution CMB data to be used for a measurement of the T&E power spectra, with a coadded white noise level of $3.3 \mu\text{K}$ -arcmin in temperature and $5.1 \mu\text{K}$ -arcmin in polarization.

This unprecedented sensitivity motivated a series of pipeline modernizations which are presented in this analysis. They represent a major step forward compared to previous SPT T&E analyses such as [14]. These improvements were summarized in the introduction Section I and described in detail in Sections III to V.

The power spectra estimated from these maps cover angular scales $400 \leq \ell \leq 3000$ in TT and $400 \leq \ell \leq 4000$ in TE and EE. An extensive campaign of consistency tests, at the power spectrum and cosmological parameter level, was performed while blind to final results to ensure robustness, as described in Section VI. While we respected all pre-established criteria to unblind the results, we still found that substantial changes to the pipeline were required after unblinding, in particular the modeling of higher-order temperature-to-polarization leakage and polarized beams, as described in detail in §IV B 2, §IV B 3, and Appendix B. The evidence for these effects is established in cosmology-independent tests. Constraining the polarized beams from T&E data alone degrades sensitivity to cosmological parameters, and a precise independent characterization of the polarized beams would increase the constraining power of the EE band powers in particular.

Overall, the model for astrophysical contaminants and systematic effects is accurate enough to ensure the consistency between power spectra across frequencies to

⁴⁷ Using Union3 [196] or DES-SN5YR [197] SNe Ia data instead of Pantheon+, the deviation is expected to be even larger [3].

better than 0.01% of sample variance. The resulting binned TT, TE, and EE power spectra, or band powers, constructed using data in pairs of SPT-3G frequency bands, are shown in Fig. 9. The minimum-variance combinations of all frequency pairs in TT, TE, and EE are shown in Fig. 1. The minimum-variance band powers in EE and TE are the most constraining to date at $\ell = 1800\text{--}4000$ and $\ell = 2200\text{--}4000$, respectively, as shown by the signal-to-noise ratio in Fig. 11.

We use these data to infer constraints on cosmological parameters which are summarized in Section II and reported in detail in Section VII. The main results are:

1. The Λ CDM standard model of cosmology provides an excellent description of our data, as shown in the residual plots in Fig. 9. We find excellent agreement with the results of *Planck* and ACT as shown in Fig. 1, Fig. 2, and Table I.
2. The SPT-3G data provide constraints on some cosmological parameters, such as the Hubble constant, which are comparable to those from *Planck*. In particular, we find $H_0 = 66.66 \pm 0.60 \text{ km s}^{-1} \text{ Mpc}^{-1}$, a 6.2σ tension between SPT-3G D1 data alone and the latest SH0ES results [2], as also shown in Fig. 20.
3. For the first time, a combination of data from ground-based experiments, namely SPT+ACT, reaches the constraining power of the *Planck* satellite data on some cosmological parameters, such as H_0 and σ_8 . This is a milestone for modern cosmology and the beginning of a new era for CMB experiments. The combination of these three CMB experiments in CMB-SPA provides the tightest CMB constraints to date. The Hubble tension with SH0ES grows to 6.4σ with the constraint of $H_0 = 67.24 \pm 0.35 \text{ km s}^{-1} \text{ Mpc}^{-1}$ derived from CMB-SPA. While the SPT-3G D1 data alone have large uncertainties on the spectral index n_s , we highlight that CMB-SPA provides $n_s = 0.9684 \pm 0.0030$ in Λ CDM, a 10.5σ difference from a scale invariant spectrum with $n_s = 1$.
4. The values of σ_8 and Ω_m from SPT-3G are in excellent agreement with the findings of other CMB experiments. We report $\sigma_8 = 0.8158 \pm 0.0058$ and $\Omega_m = 0.3246 \pm 0.0091$. We confirm that some recent results from large scale structure probes, such as the weak lensing measurements from DES-Y3 [38] and KiDS [37], align well with the CMB data, to better than 2σ , as shown in Fig. 21.
5. The SPT+ACT primary anisotropy data prefer a lensing amplitude A_{lens} in CMB spectra in agreement with the Λ CDM expectations, contrary to the slight excess found in *Planck* data, as illustrated in Fig. 22.
6. There is a growing discrepancy between the CMB and BAO measurements from DESI in Λ CDM

in the $\Omega_m\text{--}hr_d$ parameter space, as quantified in Fig. 25.

7. The above discrepancy is relaxed in extended models of cosmology. While the CMB alone does not prefer any deviations from Λ CDM at greater than 1.5σ , the combination with DESI shifts the extension parameters away from the CMB best-fit values. The most-preferred departures from the standard model are $A_{\text{lens}} > 1$, $\Sigma m_\nu < 0.06 \text{ eV}$, $\Omega_k > 0$, and $w_0 > -1, w_a < 0$; after accounting for the introduction of additional parameters, the corresponding models are preferred over Λ CDM at $2\text{--}3\sigma$ each, as summarized in Fig. 26 and Table VII.

At this point, we do not interpret the joint CMB plus DESI results as definitive evidence for a breakdown of the standard model; the goodness-of-fit improvements extended models offer over Λ CDM are moderate. The hints of new physics are driven by the combination of the two probes and are not detected by either of them independently. We conclude that while the differences between CMB and DESI data are an interesting avenue in the search for new physics, the possibility that these are due to statistical or systematic effects is not ruled out. Hopefully, improved data from these two probes and others will provide new insight.

This analysis is the third, after [16, 17], in a series of CMB power spectrum papers based on the SPT-3G observations carried out in 2019 and 2020. We are currently improving our measurement of the distortion of the CMB due to weak gravitational lensing by including temperature information, expanding on the results from polarization alone used in this paper and presented in [16].

One of the main limitations to the statistical power of this analysis is the sample variance resulting from the small sky fraction observed. This will be improved by the measurement of the primary CMB spectra on an additional 2600 deg^2 observed during the austral summer in 2019-2020 and 2020-2021 (the SPT-3G Summer fields, see Fig. 3). The analysis of these fields is ongoing and is in an advanced stage. Furthermore, the analysis of additional 6000 deg^2 observed in 2024 is also ongoing. Together with the SPT-3G Main and Summer fields, this will allow us to infer cosmological constraints from 10000 deg^2 , approximately 25% of the sky. Finally, the full SPT-3G survey will include a total of at least seven years of observations of the Main field, reaching a coadded sensitivity of no worse than $1.6 \mu\text{K}\text{--arcmin}$ in temperature. This will allow us to probe even smaller angular scales in polarization with high precision. We forecast that these upcoming data sets will provide substantial improvements compared to the results presented in this paper [41]. The analysis of this data will benefit from techniques introduced here which serve as a blueprint for future work. The full SPT-3G data set will enable precise new tests of the Λ CDM cosmological model and searches for physics beyond it.

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Appendix A: Filtering and transfer function

In §IV A 2, we detailed the computation of the filtering artifacts and the transfer function. In this appendix, we showcase a map-level example of filtering artifacts. We also detail how we propagate the error on the estimation of the transfer function to parameters, and we describe how we evaluate the consistency of the simulation pipelines.

1. Filtering artifacts

As discussed in §IV A 1, the filtering of timestream data, particularly the masking of point sources, introduces artifacts in the maps. These artifacts manifest as elongated stripes aligned with the scan direction in the T map, as shown on a simulation in Fig. 31. Following the notations of §IV A 1, we display a map computed as

$$T^{\text{artifacts}} \equiv T^{\text{f;MMmask}} - T^{\text{f;no MMmask}}. \quad (\text{A1})$$

The stripes follow lines of constant elevation, making them highly localized in Fourier space. This localization enables effective post-processing using an additional $a_{\ell m}$ weighting, as described in Eq. (5).

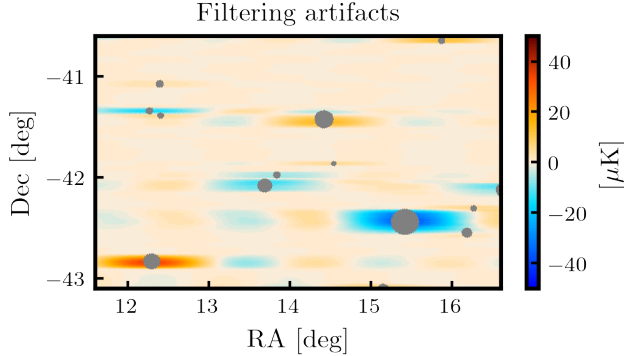


FIG. 31. Figure of filtering artifacts in the T map. The gray disks indicate the mask applied to remove point sources. Colored regions highlight filtering artifacts, which appear as elongated stripes aligned with the scan direction (horizontal axis). These artifacts result from the masking of point sources during timestream filtering in the map-making process.

2. Error on the estimation of the transfer function

When estimating the transfer function, we rely on $N_{\text{MC}} = 2000$ MC simulations of the data as described in §IV A 2. We expect a residual error on the estimation of the transfer function that we propagate to cosmological parameters. Since F_ℓ is the ratio of averages of correlated variables, the standard deviation of the error made on its

estimate ΔF_ℓ is related to the standard deviation of the individual power spectra $\Delta C_\ell \equiv \left\langle (\delta C_\ell)^2 \right\rangle_{\text{sims}}^{1/2}$ by

$$\frac{\Delta F_\ell}{F_\ell} = \left(\frac{1 + F_\ell^2/H_\ell - 2F_\ell E_\ell/H_\ell}{N_{\text{MC}} - 1} \right)^{1/2} \frac{\Delta C_\ell^{\text{f}}}{C_\ell^{\text{f}}}, \quad (\text{A2})$$

$$= \left(\frac{1 + H_\ell/F_\ell^2 - 2E_\ell/F_\ell}{N_{\text{MC}} - 1} \right)^{1/2} \frac{\Delta C_\ell^{\text{u}}}{C_\ell^{\text{u}}}, \quad (\text{A3})$$

where H_ℓ is defined in Eq. (11) and we introduce

$$E_\ell \equiv \frac{\left\langle \delta C_\ell^{\text{XY,f}} \delta C_\ell^{\text{X'Y',u}} \right\rangle_{\text{sims}}}{\left\langle \delta C_\ell^{\text{XY,u}} \delta C_\ell^{\text{X'Y',u}} \right\rangle_{\text{sims}}}. \quad (\text{A4})$$

Assuming $F_\ell \simeq H_\ell \simeq E_\ell$, as confirmed by simulations, we can simplify the above expressions to

$$\frac{\Delta F_\ell}{F_\ell} \simeq \left(\frac{1 - F_\ell}{N_{\text{MC}} - 1} \right)^{1/2} \frac{\Delta C_\ell^{\text{f}}}{C_\ell^{\text{f}}}, \quad (\text{A5})$$

$$\simeq \left(\frac{1/F_\ell - 1}{N_{\text{MC}} - 1} \right)^{1/2} \frac{\Delta C_\ell^{\text{u}}}{C_\ell^{\text{u}}}. \quad (\text{A6})$$

Setting aside the N_{MC} factor, it shows that, because the ratio of Eq. (9) cancels some of the cosmic variance, for all exploitable multipoles (where $0 \leq F_\ell \leq 1$), the relative error on F_ℓ is smaller than the one on C_ℓ^{f} , and also smaller than the one on C_ℓ^{u} for scales such that $F_\ell \geq 1/2$.

3. Consistency of the simulation pipelines

Quickmock and Fullmock simulations do not agree perfectly on the output $C(\ell)$ they produce for the same input maps. We found that the relative discrepancy between the respective transfer functions can be recast as

$$\frac{F_\ell^{\text{QM;XY};\mu\nu}}{F_\ell^{\text{FM;XY};\mu\nu}} - 1 = \rho_\ell + \varepsilon_\ell^{\text{XY};\mu\nu} \quad (\text{A7})$$

where

$$\rho_\ell \equiv (1.4 \times 10^{-3} + 2000/\ell^2) (1 - (350/\ell)^2) \quad (\text{A8})$$

is the same for all spectra and frequency combinations, having a peak value of $\rho = 4.8116 \times 10^{-3}$ at $\ell = 517.66$ while the residual error

$$\left| \varepsilon_\ell^{\text{TT};\mu\nu} \right| \lesssim 10^{-3} < \frac{1}{10} \frac{\Delta C_\ell^{\text{TT};\mu\nu}}{C_\ell^{\text{TT};\mu\nu}} \text{ for } \ell \leq 5000, \quad (\text{A9})$$

$$\left| \varepsilon_\ell^{\text{EE};\mu\nu} \right| \lesssim 2 \times 10^{-3} < \frac{1}{10} \frac{\Delta C_\ell^{\text{EE};\mu\nu}}{C_\ell^{\text{EE};\mu\nu}} \text{ for } \ell \leq 3000, \quad (\text{A10})$$

where ΔC_ℓ is the power spectrum standard deviation for a bin size of $\Delta\ell = 50$. The final transfer functions are then defined as

$$F^{\text{XY};\mu\nu} \equiv \frac{F^{\text{QM;XY};\mu\nu}}{1 + \rho_\ell}. \quad (\text{A11})$$

Appendix B: Post-unblinding changes

As discussed in the description of the blinding procedure in §VI A, after unblinding we discovered and corrected two previously untreated systematic effects: (1) quadrupolar T-to-P leakage and (2) partial depolarization of the beam sidelobes. The quadrupolar T-to-P leakage is seen most clearly in the coadded Q and U maps of bright point sources and was determined to be the cause of the failure of TE and EE band-power difference tests above $\ell = 3000$, which initially led to the exclusion of these data from the analysis. This effect had a substantial impact on the consistency between the TT and TE likelihoods. We addressed this by implementing models based on leakage beams measured from bright point sources and propagated the correction to the band powers. After applying this correction, all band-power difference tests passed within the baseline ℓ range, and improved agreement under the Λ CDM model was achieved between the TT and TE likelihoods.

Evidence for the depolarization of beam sidelobes—or, equivalently, for differences between the temperature and polarization beams—arises primarily from discrepancies in the EE power measured on the same sky at 95 and 150 GHz. While correcting for this effect clearly improves the pre-unblinding frequency consistency tests, we note that these tests formally passed prior to the correction. The observed sidelobe depolarization also led to differences in the cosmological parameter values favored by the TT and EE likelihoods, most notably in the n_s and $\Omega_b h^2$ plane. Although this initial discrepancy motivated further investigation of residual systematics in our data, we emphasize that we have compelling, cosmology-independent evidence supporting this model.

The models we use to correct for both of these effects were described briefly in §IV B and §V B; in this appendix we provide more detailed explanations of these effects and discuss their impact on inferred cosmological parameters.

1. Quadrupolar beam leakage

After unblinding, we found that we were not accounting for a substantial quadrupolar T-to-P leakage. This leakage was responsible for an initial failure of the band power-level consistency tests in TE and EE spectra for $\ell \in [3000, 4000]$, which led us to abandon these data before unblinding. Additionally, it resulted in a bias in the TE band powers and incompatible results between temperature and polarization data in Λ CDM. As illustrated in Fig. 7, this leakage exhibits coherence across frequencies at large angular scales and is largest, relative to the uncertainty on the power spectrum, at small scales. This explains why the leakage was not identified in the interfrequency consistency tests, except in the previously excluded $\ell \in [3000, 4000]$ range, and underscores a limitation of our blinding procedure, which proved insufficient for identifying and treating all

systematic effects present in the data. This leakage was not corrected for in the MUSE-based analysis of [16]. That work was based on EE-only measurements and did not detect significant quadrupolar leakage.

A quadrupolar T-to-P leakage arises because the linearly polarized beams from individual detectors are slightly elliptical along the axis of polarization. The difference in the response to an unpolarized source for two orthogonally polarized detectors, with this polarization-direction-dependent ellipticity, results in a leakage signal [49–51]. This leakage is a quadrupolar pattern in the Q and U maps, which is particularly visible around point sources, as shown in Fig. 32. In this figure, we show the residual point source T-to-P leakage in the Q and U maps after monopole subtraction, as well as the template for the quadrupolar signal obtained from the model fit described below.

Although the amplitude of this leakage is expected to be small, it can have a significant impact on power spectrum estimation if not accounted for. To properly capture this effect, we assume that the T-to-P leakage is a convolution of the underlying temperature map with leakage beams specific to the Q and U maps, which we label $B^{T \rightarrow Q}$ and $B^{T \rightarrow U}$, respectively. These functions correspond to the Q and U response from an unresolved pure temperature signal. We model the leakage beams using an expansion of orthogonal Hermite polynomials:

$$\begin{aligned} B^{T \rightarrow Q;\mu}(x, y) &= B_{\sigma_\mu}(x, y) \sum_{m+n} a_{m,n}^{T \rightarrow Q;\mu} H_{m,n} \left(\frac{x}{\sigma_\mu}, \frac{y}{\sigma_\mu} \right), \\ B^{T \rightarrow U;\mu}(x, y) &= B_{\sigma_\mu}(x, y) \sum_{m+n} a_{m,n}^{T \rightarrow U;\mu} H_{m,n} \left(\frac{x}{\sigma_\mu}, \frac{y}{\sigma_\mu} \right) \end{aligned} \quad (\text{B1})$$

with:

$$H_{m,n} \left(\frac{x}{\sigma_\mu}, \frac{y}{\sigma_\mu} \right) = H_m \left(\frac{x}{\sigma_\mu} \right) H_n \left(\frac{y}{\sigma_\mu} \right) \quad (\text{B2})$$

where, for each frequency μ , B_{σ_μ} is a Gaussian of width σ_μ , $a_{m,n}^{T \rightarrow Q;\mu}$ and $a_{m,n}^{T \rightarrow U;\mu}$ are the coefficients of the expansion, and $H_m(x/\sigma_\mu)$ are Hermite polynomials orthogonal with respect to the measure e^{-x^2/σ_μ^2} . Hermite polynomials depend on the Gaussian width σ_μ to guarantee their orthogonality.

We combine 418 thumbnail maps centered at the location of bright sources into 25 coadded maps of the leakage in Q and U. The coaddition allows us to reach a significant detection over the CMB signal and noise. Each of the coadded maps contains a variable number of sources in order to reach comparable signal to noise. The coaddition is straightforward thanks to the telescope's scanning strategy, which makes the detectors always oriented in the same direction with respect to the sky; the measured beams are thus independent of the sky location. The upper row of Fig. 32 displays the coadded maps from the location of all 418 bright sources used in this analysis. We find no significant detection of monopole leakage,

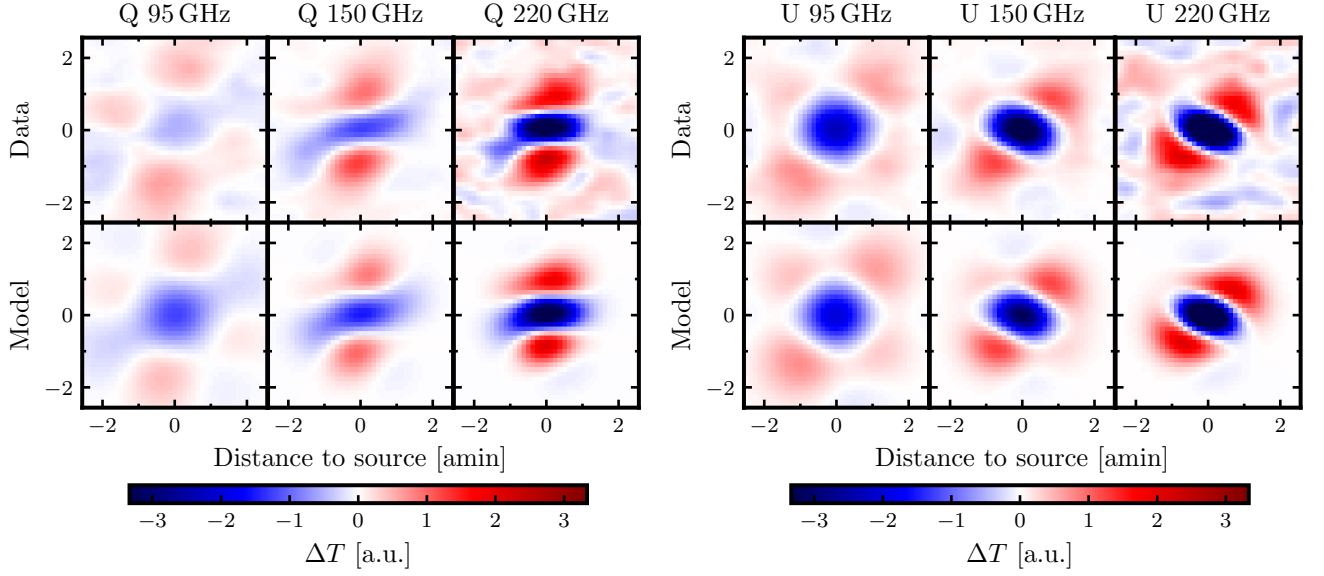


FIG. 32. Residual point source T-to-P leakage in the Q and U maps in arbitrary units. *Top*: Coaddition of 418 bright point-like sources in the Q and U maps after monopole subtraction. *Bottom*: Template obtained from the fit of the leakage beams, using the model of Eq. (B1).

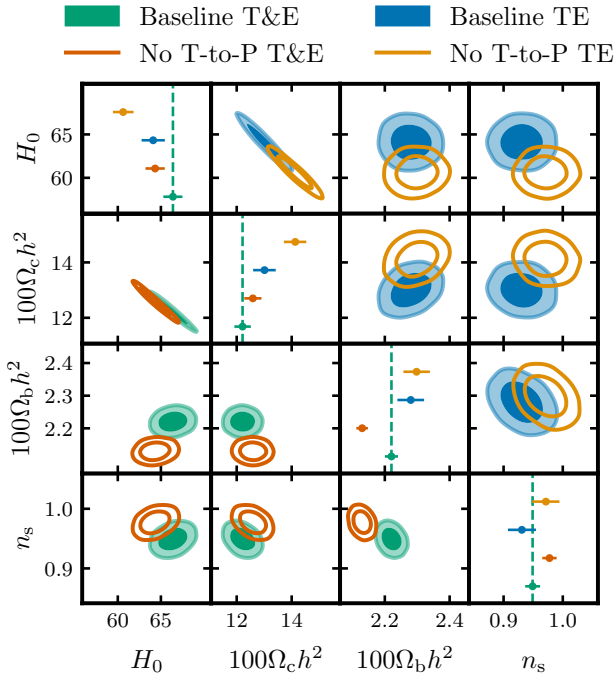


FIG. 33. Comparison between the baseline likelihood and a likelihood without the quadrupolar T-to-P leakage correction. Contours indicate the 68% and 95% confidence levels. The dashed vertical line marks the mean value of the results of the baseline likelihood. The upper right panel shows TE constraints, and the lower left shows T&E constraints.

as expected given the monopole deprojection described in §III D, but detect a significant quadrupolar and a mildly significant hexadecapolar leakage. We measure the mean and the standard deviation of the beam leakage parameters from fits to the 25 coadded maps. The lower row of Fig. 32 shows the mean template obtained from this procedure. We see a pattern of decreasing angular scale for the leakage with increasing frequency as expected from the temperature beam shapes. The pattern in U is similar to the one in Q but rotated by an angle of $\pi/4$.

We propagate the measured model coefficients to the band powers. This is done based on an analytical model of the leakage described in Eq. (14) and confirmed through simulations to be the convolution of the temperature map by the leakage beam. The amplitude of the quadrupolar leakage is quantified by

$$\epsilon_2^\mu = (a_{2,0}^{T \rightarrow Q;\mu} - a_{0,2}^{T \rightarrow Q;\mu} + a_{1,1}^{T \rightarrow U;\mu})/2. \quad (\text{B3})$$

As expected, the TE and EE contribution is only connected to the spin-2 component of the leakage at the map level, without the pure spin-0 radial contribution ($a_{2,0}^{T \rightarrow Q} + a_{0,2}^{T \rightarrow Q}$). We report these values in §IV B 2. We show the expected contribution to the band powers from this leakage in Fig. 7.

After accounting for this effect in the data model, all consistency tests between bands and spectra passed across the full angular multipole range. We tested variations of this model where the leakage was assumed to contain additional orders (dipole, octopole, hexadecapole). Although hexadecapolar leakage is detected with mild significance in the map-level analysis, we

verified that including it in the model does not impact cosmological parameter estimation. In this analysis, we incorporate the leakage through forward modeling in the theoretical prediction. As a cross-check, we also implemented an alternative approach in which the leakage is subtracted directly from the data band powers using the measured TT and TE band powers rather than the Λ CDM model band powers. Both methods yield consistent results. Given no clear preference for variations on the quadrupolar leakage model presented here, it was adopted in the baseline analysis described in §IV B 2.

In Fig. 33, we show the impact of the T-to-P leakage correction on cosmological parameters. We compare Λ CDM constraints obtained from the baseline T&E and TE likelihoods with those obtained without the T-to-P leakage correction. The impact on the TE constraints is significant, with close to 3σ shifts in H_0 , n_s , and $\Omega_c h^2$. We also show the effect of T-to-P leakage correction on the full T&E likelihood, but we caution that these results are more difficult to interpret, because without this correction the TE Λ CDM constraints are not sufficiently compatible with those from TT and EE to be properly combined. With this caveat, we note that the biggest shift for the T&E likelihood is on $\Omega_b h^2$ due to the degeneracy of this parameter with both H_0 and $\Omega_c h^2$. Including the T-to-P leakage correction in the T&E likelihood yields a $\Delta\chi^2 = 94$ improvement in the fit to the data. This improvement is highly significant, considering that the model introduces only three additional degrees of freedom.

2. Polarized beams

Unlike for the temperature beam, there are no sufficiently bright and polarized detected sources that can be used to directly map the polarized beam with the required signal-to-noise. As has been done for previous CMB analyses, prior to unblinding we assumed that the polarized beam shape was identical to that of the measured temperature beam. The measured temperature beam has significant sidelobe power, arising from sources such as diffraction from primary mirror panel gaps and multiple reflections from optical elements, that is not captured in simulations and for which we have no knowledge of the optical path. In retrospect, it is unrealistic to assume that this sidelobe power is polarized identically to the main beam.

Similarly to the MUSE analysis [16], after unblinding it became apparent that our assumptions about the polarized beam were unjustified. As stated in §IV B and §V B, prior to unblinding we assumed beam sidelobes polarized equally to the main beam, i.e. $\beta_{\text{pol}} = 1$. In the baseline model adapted for this work and for the MUSE [16] analysis of the same data, we allow the degree of sidelobe polarization for each of the bands relative to the main beam, β_{pol}^ν , to vary.

There is clear evidence from the data supporting this model. First, consistency of the polarization data in the three observation bands (particularly for $\ell < 1200$) requires fractional depolarization of the beam sidelobes. Under the original, rigid, assumption of identical temperature and polarized beams, the PTE of the 150×150 GHz vs 95×150 GHz difference test was 0.004. Though borderline passing our blinding threshold of 0.28%, this is low and an inconsistency was visible in the difference spectrum (see Fig. 41), particularly in all the comparisons between 95 and 150 GHz channels. When allowing for a varying sidelobe polarization fraction, the PTE of the difference test is 51%, a substantial improvement. When considering only the agreement at multipoles below $\ell = 1000$, the PTE for the $(95 \times 95 \text{ GHz}, 95 \times 150 \text{ GHz})$, $(95 \times 95 \text{ GHz}, 150 \times 150 \text{ GHz})$, and $(95 \times 150 \text{ GHz}, 150 \times 150 \text{ GHz})$ EE difference tests improve from 4.4%, 1.1%, and 0.27% to 60%, 43%, and 24%, respectively.

Second, a cosmological-model-free reconstruction of the binned CMB power spectra strictly prefers depolarization of the beam sidelobes. We perform this reconstruction without the need to assume a cosmological dependence using the *lite* framework introduced in §V D. In order to marginalize over nuisance and foreground contributions, the framework first estimates the best fit parameters of the likelihood, replacing in Eq. (40) the cosmological set of parameters ψ by binned CMB spectra [31]. Within the *lite* framework, we call this procedure the estimation of the reconstruction likelihood. This relies on the assumption that up to nuisance and foregrounds parameters, the power of the CMB signal is the same in bins covering the same multipole range of the same spectrum.⁴⁸ In other words, this allows us to estimate the best foreground and nuisance model parameters that minimize the discrepancy between cross-frequency spectra, without assuming any cosmological model. Note that minimizing the foreground and nuisance-induced discrepancy between cross-spectra is not immune from any coherent effect across frequency channels. With this limitation in mind, the *lite* framework is an efficient tool to test for different models of the systematics, in a cosmology-independent way and compare the best-fit χ^2 values.

We use this approach to investigate the preference of the data for different models of the instrumental beam. We test two cases: (1) the baseline (post-unblinding) model that includes a fractional depolarization of the beam sidelobes and (2) the pre-unblinding model with identical beam shapes in temperature and polarization, i.e. fixing $\beta_{\text{pol}}^\nu = 1$ in Eq. (40). We minimize the reconstruction likelihoods and calculate the associated χ^2

⁴⁸ E.g. the CMB power is the same in the EE bin covering $1000 < \ell < 1050$ at 95×150 GHz and 220×220 GHz. This assumption does not strictly hold as the window functions are not identical across frequencies, though deviations are less than 0.2% and hence negligible for this purpose.

values. We report χ^2 (PTE) values of 1198.64 (31.05%) for model (1) and 1229.46 (14.57%) for model (2), respectively. While both models provide an acceptable fit to the data, the difference in χ^2 is 30.8. The χ^2 improvement is concentrated on large angular scales with 22.7 points below $\ell = 1000$ and 26.9 points below $\ell = 1500$. This is significant and, when adjusting for the additional degrees of freedom between the models, the χ^2 improvement on the full ℓ range translates to a difference in the Akaike Information Criterion of 24.8. The data strongly prefer the beam sidelobe depolarization model, when assuming no cosmological model for the CMB signal in the data, and reject the alternative hypothesis that the temperature and polarized beam shapes are identical.

Third, we note that, similar to the **SPT-lite** reconstruction discussed above, the MUSE analysis [16] also found strong evidence in favor of the polarized beam model, independent of any cosmological model assumptions. In that pipeline, which is close to simulation-based inference, the binned EE and $\phi\phi$ spectra are first reconstructed—together with systematics—without imposing a cosmological model. At this stage, the baseline polarized beam model presented here was preferred over the rigid one.

Finally, we can evaluate this preference when assuming a Λ CDM model. In this case, the polarized beam model is preferred over the rigid one at $\Delta\chi^2 = 30$. This number is consistent with the $\Delta\chi^2$ obtained from the *lite* reconstruction likelihood, highlighting that the preference for the polarized beam model does not require assuming any particular cosmological model, but is rather required by the differences between frequency bands. One expects the largest signature of polarized beams in the EE data, which in fact does contribute $\Delta\chi^2 = 24$.

In the left panel of Fig. 34, we show the impact of the polarized beam correction on the cosmological parameters. We compare the Λ CDM constraints obtained from the baseline T&E and EE likelihoods with those obtained without the polarized beam correction. The polarized beam model mostly affects the n_s and $\Omega_b h^2$ plane, with a 2σ shift in n_s for EE alone. In the right panel, we show that fitting the beam parameters within the Λ CDM model yields consistent values for β_{pol} when using either the EE-only or the full T&E data set. Both likelihoods yield a 5σ detection of $(\beta_{\text{pol}}^{95}, \beta_{\text{pol}}^{150}, \beta_{\text{pol}}^{220}) \neq (1, 1, 1)$ in the associated 3d parameter space. The small statistical shift observed between the two cases arises from common modes in the polarized beam corrections that, within the uncertainties, are degenerate with cosmological parameters. This degeneracy is reduced when all spectra are included, leading to tighter constraints. The cosmological parameters most sensitive to the beam depolarization effect shift by at most 1σ . As demonstrated in §VID and Fig. 16, accounting for the correlation between the T&E and EE-only data sets, both likelihoods yield fully consistent results.

It is important to highlight that this model is not specifically designed to treat small angular scales, where beam differences propagate to relative slopes (see [6]), but has the largest impact on large angular scales where the data are poorly described without it. Excising the large scale data naturally weakens the detection of $(\beta_{\text{pol}}^{95}, \beta_{\text{pol}}^{150}, \beta_{\text{pol}}^{220}) \neq (1, 1, 1)$, though it does not shift cosmological constraints by more than the expected amount, as demonstrated in §VIIB1. In particular, constraints on n_s and $\Omega_b h^2$ are stable when removing $\ell < 1200$, $\ell < 1500$, or $\ell < 2000$ data. In fact, forcing $(\beta_{\text{pol}}^{95}, \beta_{\text{pol}}^{150}, \beta_{\text{pol}}^{220}) = (1, 1, 1)$ does not impact constraints on Λ CDM parameters when fitting data from $\ell > 2000$ alone.

In addition, we performed several tests of the underlying assumptions of the polarized beam model. First, we investigated the possibility of a systematic effect mimicking reduced sidelobe efficiency in the TT spectra by applying the same polarized beam model to the temperature data. This serves as a check of the analysis pipeline, as there is no physical motivation for such an effect. We find no evidence for reduced sidelobe efficiency in the temperature data; the measured sidelobe efficiencies are consistent with unity within 1.5σ . The absence of evidence for reduced sidelobe efficiency is further supported by the TT difference tests presented in §VIB1. Second, we tested a variation of the beam model which allowed for a sidelobe polarization fraction that varies as a function of scale. Introducing this freedom did not further improve the consistency between the band powers from different frequencies using the *lite* framework or the fits to Λ CDM cosmology. Third, we tested another variation where the shape of the main beam, which is calculated analytically and informs the shape of the sidelobes, is replaced by its best-fit Gaussian approximation for each frequency band, instead of using the calculated beams. Despite this extreme change, the preference for sidelobe depolarization and the resulting cosmological parameters were nearly identical. We conclude that the polarized beam model is insensitive to reasonable variations of the main beam shape and the assumption of a uniform depolarization of the beam sidelobes. The data are well described by the simple model we present, with a highly polarized main beam and a sidelobe that is uniformly fractionally polarized compared to the main beam and described by a single polarization parameter for each band.

Altogether, there is strong evidence in favor of the polarized beam model, which we adopt as the baseline. We stress that this is a conservative choice and that there is no evidence in our data to support the rigid assumption of identical temperature and polarized beams that has been commonly adopted for CMB analyses. We choose to parametrize and marginalize over the uncertainty in the sidelobe polarization fraction with the goal of eliminating bias at the cost of some constraining power on cosmological parameters. Freeing the sidelobe polarization fractions from their best fit posterior values

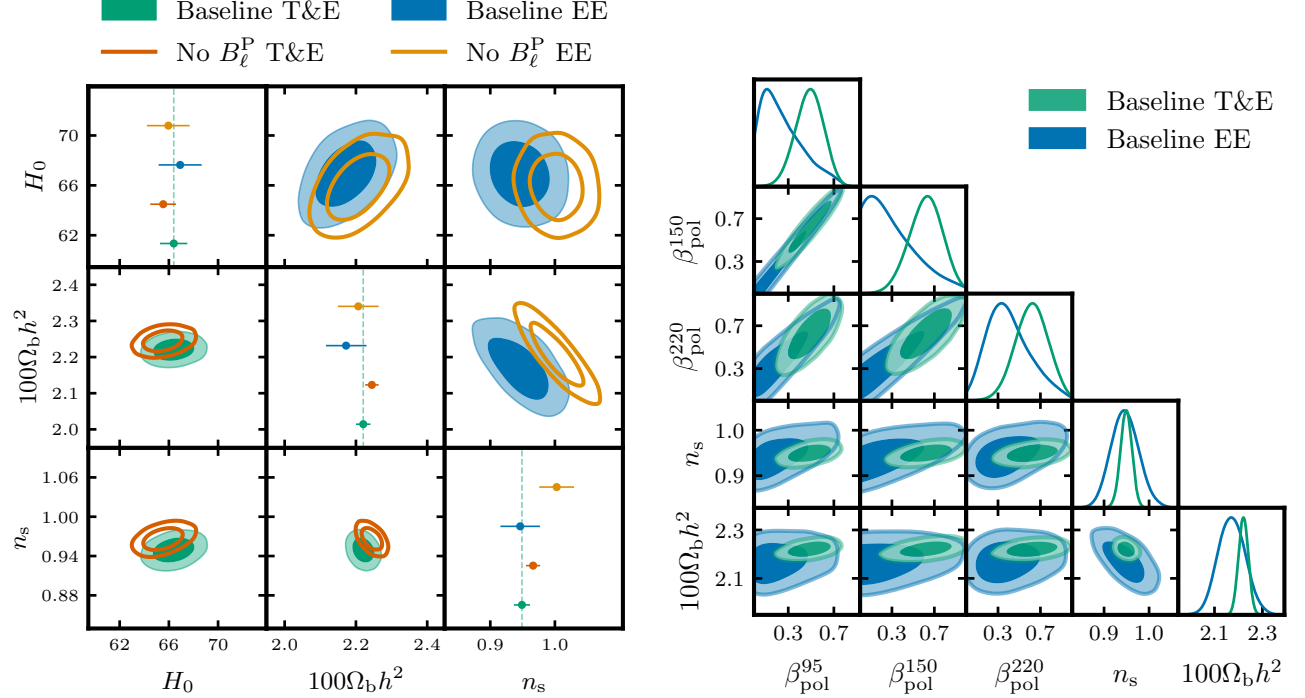


FIG. 34. *Left*: Comparison between the baseline likelihoods with and without the polarized beam correction for T&E and EE alone. The upper right panel shows EE constraints, and the lower left shows T&E constraints. *Right*: Posterior distributions of the polarized beam parameters from the baseline T&E and EE likelihoods, shown alongside parameter constraints in the n_s - $\Omega_b h^2$ plane. The distributions are non-Gaussian and unbounded, so we do not show whiskers in the diagonal plots. The figure highlights the correlation between beam parameters. The T&E and EE contours show good agreement. This pronounced preference for sidelobe depolarization is driven by the mismatch between the 95 and 150 GHz data at $\ell < 1200$, resulting in a clear exclusion of the point $(\beta_{\text{pol}}^{95}, \beta_{\text{pol}}^{150}) = (1, 1)$ in the upper left two-dimensional panel.

degrades the constraint on n_s by 20%, as can be seen in Fig. 34. Direct measurements of the polarized beams have the potential to recover this lost constraining power. However, there are no sufficiently bright and highly polarized detected sources in the southern sky that would allow us to measure the polarized beam with the required precision. Because of the large primary mirror of the SPT, it is prohibitive to place a ground-based polarized source both in the far-field and at an elevation sufficiently high to prevent detector saturation from atmospheric loading. Satellite based polarized sources have the potential to enable a direct high signal-to-noise measurement of polarized beams for a large aperture telescope and reduce this source of uncertainty [203, 204].

3. Miscellaneous post-unblinding changes

In this subsection, we present various minor updates that have been made to the pipeline after unblinding. These updates did not have a significant impact on the results; they are listed for completeness.

First, we realized that the covariance used for unblinding was erroneously missing the lensing contribution

described in §IV F, as well as the regularization factor (Eq. (38)) even though the decision had been made to include both of them already. Adding the lensing contribution increases the error bars on cosmological parameters by 10%. The covariance conditioning is required to correctly interpret the χ^2 of the data, but has a negligible impact on cosmological parameters. We also updated CMB and foreground template spectra used to compute the covariance for the best-fit model obtained from the T&E analysis to ensure consistency between the model and the data.

We corrected a mistake in the implementation of the tSZ-CIB cross-correlation model in the likelihood software. We also fixed the amplitude of EE Poisson power to zero based on expectations from the source masking threshold and the low polarization of radio sources, see [64].

Finally, minor improvements were introduced in the beam pipeline, resulting in changes to the beam B_ℓ smaller than 0.2%. We also removed unnecessary priors on polarization calibration for the T&E likelihood, letting the data drive the calibration. We further removed priors on the relative inter-frequency temperature calibration of the 95 and 220 GHz channels. However, we restore

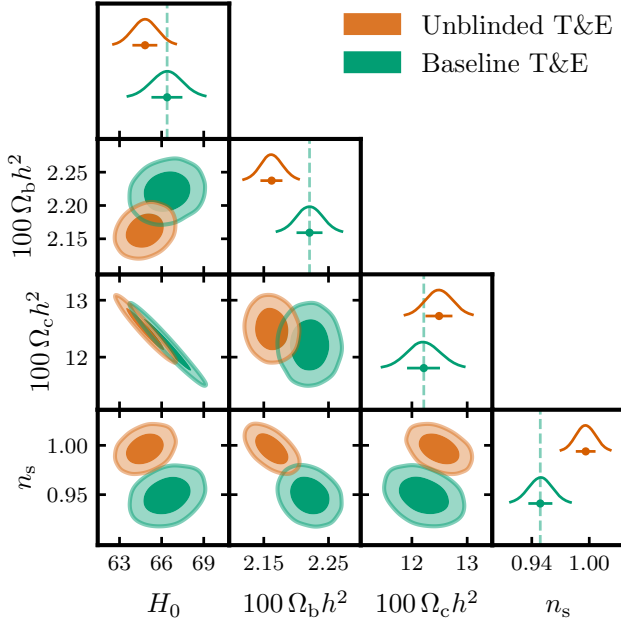


FIG. 35. Initial cosmological results from the T&E likelihood at unblinding, compared to final results. Contours indicate the 68% and 95% confidence levels. The unblinded results revealed strong inconsistencies among the individual TT, TE, and EE likelihoods, motivating the investigations detailed in Appendices B 1 and B 2. The unblinded contours plotted here rely on the combination of inconsistent subsets and are not formally meaningful.

priors on temperature and polarization calibration when analyzing subsets of the data (e.g. EE or TE alone) to break degeneracies. We also increased the uncertainty on the global temperature calibration prior calculated from the comparison with *Planck* data to account for systematic changes depending on the choice of the *Planck* map. The temperature calibration prior changed from $\sigma(A_{\text{ext}}^{\text{cal}}) = 0.0019$ to $\sigma(A_{\text{ext}}^{\text{cal}}) = 0.0036$. This conservative change has an overall negligible impact on the cosmological parameters, though it widens the $\log(10^{10} A_s)$ posterior by 4%.

4. Initial cosmological results

In accordance with our blinding procedure, we did not examine the cosmological results prior to unblinding. The initial results obtained from the T&E likelihood at the time of unblinding are shown in Fig. 35. The observed shifts in the Λ CDM parameters are substantial and can be attributed to specific systematic effects. First, the primary source of the shift in $\Omega_b h^2$ is the quadrupolar T-to-P leakage, which biases the TE band powers, as illustrated in Figs. 7 and 33. The TE-derived parameters were highly inconsistent with those from TT and EE, with none of the PTE tests passing. Enforcing agreement

among the channels within the Λ CDM model resulted in the significant discrepancies shown in the figure. Then, the shift in n_s is driven by both the T-to-P leakage and the introduction of the polarized beam sidelobe model.

Our experience highlights that the blinding procedure implemented in this analysis was insufficient to identify all systematic effects present in the data, as evidenced by the substantial shifts in cosmological parameters following unblinding. The procedure was primarily designed to flag systematics that manifest as incoherent features across frequencies. Consequently, the quadrupolar T-to-P leakage went undetected at $\ell < 3000$, and the depolarized sidelobes were similarly missed due to their near coherence across frequencies. Notably, a low PTE in the EE band-power difference test at $\ell < 1200$ did provide an early indication of the latter effect.

Appendix C: Covariance matrix

The covariance matrix computation is detailed in §IV F. In this appendix, we justify the additional noise term in Eq. (34) and we showcase the mixing matrix for deeper understanding of the data set.

1. Cross-bundle covariance matrix

In this work, we estimate the power spectrum by taking cross-products of maps from different bundles to avoid a noise bias. Another approach is to estimate the band powers directly from all the data compiled together in a single coadd. The covariance matrix of the mean cross-bundle power spectrum is different than the covariance matrix of the coadd power spectrum. Assuming that the noise N_ℓ is Gaussian and uncorrelated between n_b bundles and that maps cover the full sky, the covariance matrix of the mean cross-bundle power spectrum is

$$\Sigma_{\ell\ell}^{\text{bundles}} \propto \left(2C_\ell^2 + 2C_\ell N_\ell + \frac{2n_b}{n_b - 1} N_\ell^2 \right), \quad (\text{C1})$$

where C_ℓ is the fiducial signal. This equation is derived from the framework introduced in [44]. The noise variance is boosted by a factor of $n_b/(n_b - 1)$ compared to the coadd covariance. This is due to the fact that we avoid auto-bundle spectra in our framework, again to avoid a noise bias. For our case of 30 bundles, the noise variance is increased by 3%. In the infinite bundle limit, the covariance matrix of the mean cross-bundle power spectrum is the same as the covariance matrix of the coadd power spectrum.

2. Mixing matrix

From the covariance computed in §IV F, we can compute the mixing matrix, which describes the contribution

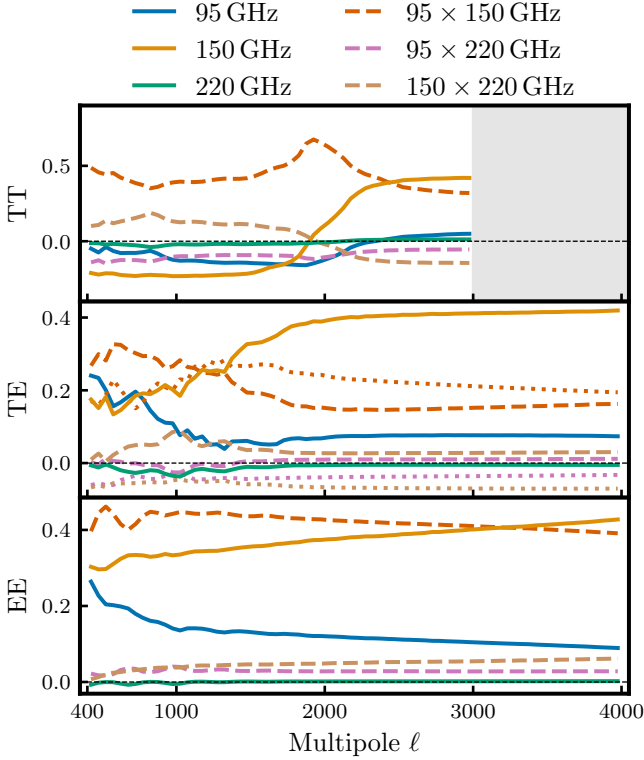


FIG. 36. Diagonal elements of the mixing matrix M^{mix} for the TT (top), TE (middle), and EE (bottom) spectra. The dashed lines indicate the contributions from the cross-frequency spectra, 95×150 GHz, 95×220 GHz, and 150×220 GHz. The dotted lines in the TE plot indicate the contributions from the inverted cross frequency spectra, 150×95 GHz, 220×95 GHz, and 220×150 GHz. The mixing matrix is defined in Eq. (C2). The diagonal elements of the mixing matrix are the weights of each cross-frequency spectrum in the minimum variance estimator of the band powers. We normalize the mixing matrix here such that the absolute value of all elements at a given ℓ sum to one for TT, TE, and EE separately.

of each of the cross-frequencies to the minimum variance estimator of the band powers. The mixing matrix is defined as

$$M^{\text{mix}} = (X^{\top} \Sigma^{-1} X)^{-1} X^{\top} \Sigma^{-1}, \quad (\text{C2})$$

where X is the design matrix that dictates which cross-frequency spectra are used to estimate the band powers and Σ is the covariance matrix. It makes it possible to compute the minimum variance estimator of the band powers as used in Figs. 1 and 11 (for details on the procedure see [31, 86, 205]).

The mixing matrix is straightforward for EE, highlighting that the deepest channels, 95 and 150 GHz, are contributing the most to the measurement of the band powers. This also explains why we are particularly sensitive to discrepancies between the two channels. In TE, the mixing matrix is still dominated by the 95 and 150 GHz channels, although the contributions from

95×150 GHz is different than 150×95 GHz, due to the noise structure. Negative terms in the mix matrix can arise when channels are correlated. Finally, the TT mixing matrix displays two regimes. At large scales, the signal and the noise are correlated; while the 95×150 GHz spectrum contributes the most weight here, the information is more equally distributed between the cross-frequency spectra compared to the TE and EE cases. For $\ell \gtrsim 2000$ the noise correlation falls (see Fig. 10) and the mixing matrix is then dominated by the 150×150 GHz and 95×150 GHz spectra.

Appendix D: Likelihood nuisance parameters

A set of nuisance parameters is included in the likelihood to model uncertainties in calibration, beam characterization, and residual systematics, as described in §V B. Table IX summarizes these parameters, providing their descriptions, priors, and default values used in simulation fits. For calibration and polarization efficiency parameters, we also specify the priors adopted when analyzing data subsets; these are derived from comparisons with *Planck* data over the same sky region (see §III D). Beam eigenmode priors are standard normal distributions, reflecting the amplitude uncertainties of the associated modes (see Eq. (43)). Priors on the quadrupolar beam leakage amplitudes are determined from map-level analyses (see Appendix B 1). Uniform priors are assigned to the sidelobe polarization fractions (see Appendix B 2). For each parameter, we report the best-fit value and the mean with 68% confidence intervals from the Λ CDM analysis.

We find that the posterior values of the subset calibration parameters are consistent with the priors derived from the calibration pipeline, which are not applied in the baseline analysis.

To evaluate the robustness of the systematic model, we first broaden the priors on the beam eigenmodes by a factor of 10. This increases the flexibility of the beam model. This leads to shifts in cosmological parameters of less than 0.2σ , except for n_s , which shifts by 0.4σ and shows a 30% increase in its uncertainty. Since n_s characterizes the tilt of the primordial power spectrum, it is naturally correlated with the beam parameters. The observed shift is not statistically significant, supporting the robustness of the beam modeling. Next, we allowed the quadrupolar beam leakage amplitudes to vary freely and found less than a 3σ Gaussian distance between the likelihood and the prior set by the map-based analysis. Thus, the data are consistent with the quadrupolar beam leakage model.

Appendix E: Foreground model

In this section, we present the explicit functional forms of the foreground model components used in the

Parameter	Description	Prior	Subset Prior	Λ CDM Best-fit	68% CL limit	Section
$A_{\text{cal}}^{\text{ext}}$	External calibration	$\mathcal{N}(1.0, 0.0036)[1.0]$		1.00002	1.0002 ± 0.0036	§III D
$A_{\text{cal}}^{\text{rel};95}$	Relative calibration factors	$\mathcal{U}(0.8, 1.2)[1.0]$	$\mathcal{N}(1.0, 0.0024)$	1.000196	1.00020 ± 0.00039	§III D
$A_{\text{cal}}^{\text{rel};220}$			$\mathcal{N}(1.0, 0.010)$	1.00853	1.0087 ± 0.0012	
$E_{\text{cal}}^{\text{ext}}$	Polarization efficiency	$\mathcal{U}(0.8, 1.2)[1.0]$	$\mathcal{N}(1.0, 0.0095)$	1.0085	1.0095 ± 0.0051	§III D
$E_{\text{cal}}^{\text{rel};95}$	Relative polarization efficiencies	$\mathcal{U}(0.8, 1.2)[1.0]$	$\mathcal{N}(1.0, 0.0022)$	0.99869	0.9986 ± 0.0011	§III D
$E_{\text{cal}}^{\text{rel};220}$			$\mathcal{N}(1.0, 0.0067)$	0.99579	0.9957 ± 0.0030	
β_1	Beam eigenmodes	$\mathcal{N}(0.0, 1.0)[1.0]$		-0.47	-0.45 ± 0.95	§IV B §V B 1
β_2				-0.62	-0.61 ± 0.88	
β_3				0.47	0.51 ± 0.92	
β_4				-0.76	-0.87 ± 0.70	
β_5				-0.06	-0.03 ± 0.93	
β_6				-1.11	-1.10 ± 0.90	
β_7				0.15	0.13 ± 0.72	
β_8				-0.36	-0.36 ± 0.99	
β_9				0.20	0.20 ± 0.98	
ϵ_2^{95}	2nd order	$\mathcal{N}(-0.0065, 0.0011)[0.0]$		-0.00687	-0.00690 ± 0.00078	§IV B 2
ϵ_2^{150}	T→P leakage	$\mathcal{N}(-0.012, 0.0021)[0.0]$		-0.01458	-0.0146 ± 0.0015	§V B 1
ϵ_2^{220}	amplitudes	$\mathcal{N}(-0.023, 0.0066)[0.0]$		-0.02844	-0.0285 ± 0.0036	
β_{pol}^{95}	Sidelobe	$\mathcal{U}(0.0, 1.0)[1.0]$		0.555	$0.48^{+0.13}_{-0.12}$	§IV B 3
β_{pol}^{150}	polarization			0.709	$0.62^{+0.17}_{-0.15}$	§V B 1
β_{pol}^{220}	fractions			0.687	0.62 ± 0.15	

TABLE IX. Summary of systematic nuisance parameters in the SPT T&E likelihood. The third column lists the priors used in the Λ CDM analysis; we fix the nuisance parameters to the values given in brackets when fitting the simulations because they are not included in the simulation generation. Calibration and polarization efficiency priors are derived from comparison with *Planck* maps; only the external temperature calibration prior is used for the full likelihood, while “subset priors” are applied when analyzing data subsets. For each nuisance parameter, we report the best-fit value from the Λ CDM analysis, along with the mean and 68% confidence interval. The final column references the section where each parameter is discussed.

analysis (see §V B 2 for additional discussion). The model includes the following terms.

First, unresolved sources are modeled as a Poisson component with a power-law dependence on multipole ℓ :

$$D_{\ell,\mu\nu}^{\text{Poisson}} = A_{\mu\nu}^{\text{Poisson}} \left(\frac{\ell}{3000} \right)^2, \quad (\text{E1})$$

where $A_{\mu\nu}^{\text{Poisson}}$ is the amplitude at $\ell = 3000$ for each cross-frequency pair $\nu \times \mu$. Based on expectations from our source masking threshold and the results of [64] we set the Poisson amplitude to zero for all TE and EE cross-frequency pairs.

Second, the clustered CIB component is modeled as a power law in ℓ :

$$D_{\ell,\mu\nu}^{\text{CIB}} = A_{\mu\nu}^{\text{CIB}} \left(\frac{\ell}{3000} \right)^{\alpha^{\text{CIB}}}, \quad (\text{E2})$$

where $A_{\mu\nu}^{\text{CIB}}$ is the amplitude at $\ell = 3000$ for each cross-frequency and α^{CIB} is the spectral index. The CIB clustering term is fit only for the 150×150 GHz, 150×220 GHz, and 220×220 GHz spectra; for other cross-frequencies, it is marginalized over in the covariance, as it is not significantly detected.

Third, the tSZ contribution is modeled using a fixed template with a frequency-dependent scaling:

$$D_{\ell,\mu\nu}^{\text{tSZ}} = A^{\text{tSZ}} f_{\nu_0}^{\text{tSZ}}(\nu, \mu) D_{\ell}^{\text{tSZ, template}}, \quad (\text{E3})$$

where A^{tSZ} is the amplitude at $\ell = 3000$, $D_{\ell}^{\text{tSZ, template}}$ is the template power spectrum, and $f_{\nu_0}^{\text{tSZ}}(\nu, \mu)$ encodes the SED of the standard tSZ frequency dependence relative to primary CMB fluctuations [78]. The template power spectrum is obtained from AGORA simulations [46] and is fixed in the analysis.

Fourth, the kSZ contribution is constant in CMB units and is modeled with a fixed template:

$$D_{\ell,\mu\nu}^{\text{kSZ}} = A^{\text{kSZ}} D_{\ell}^{\text{kSZ, template}}, \quad (\text{E4})$$

where A^{kSZ} is the amplitude at $\ell = 3000$ and $D_{\ell}^{\text{kSZ, template}}$ is the template power spectrum. The kSZ template is obtained from AGORA simulations [46] and is fixed in the analysis.

Finally, the Galactic dust contribution is modeled as a modified black body with a power-law dependence on ℓ :

$$D_{\ell,\mu\nu}^{\text{dust}} = A^{\text{dust}} g_{\nu_0}^{\text{dust}}(\mu, \nu, \beta^{\text{dust}}) \left(\frac{\ell}{80} \right)^{\alpha^{\text{dust}} + 2}, \quad (\text{E5})$$

where A_{80}^{dust} is the amplitude at $\ell = 80$, $g_{\nu_0}^{\text{dust}}(\mu, \nu, \beta^{\text{dust}})$ describes the frequency dependence of the dust emission and α^{dust} is the power-law index. The frequency dependence is given by:

$$g_{\nu_0}(\mu, \nu, \beta) = \frac{g(\mu)g(\nu)}{g(\nu_0)^2} \left(\frac{\mu\nu}{(\nu_0)^2} \right)^\beta, \quad (\text{E6})$$

where $g(\nu)$ is the modified black body function. We use this model in TT, TE, and EE spectra, with independent amplitudes and priors.

In an extended version of the likelihood, we included a model for the tSZ–CIB correlation:

$$D_{\ell, \mu\nu}^{\text{tSZ} \times \text{CIB}} = -\xi_{\text{tSZ} \times \text{CIB}} \left(\sqrt{D_{\ell, \mu\mu}^{\text{tSZ}} D_{\ell, \nu\nu}^{\text{CIB}} + \mu \leftrightarrow \nu} \right), \quad (\text{E7})$$

where the tSZ–CIB cross-correlation coefficient, $\xi_{\text{tSZ} \times \text{CIB}}$, was treated as a free parameter with a Gaussian prior $\mathcal{N}(0.18, 0.33)$. We found, however, that the data do not strongly constrain this parameter, and allowing it to vary substantially increased the computational cost of the likelihood evaluation. Therefore, in the baseline analysis, we fix $\xi_{\text{tSZ} \times \text{CIB}}$ to its best-fit value of 0.26 and marginalize over its uncertainty in the covariance matrix computation.

In Table X, we summarize the priors on the foreground parameters used in the analysis. The priors originate from our previous T&E analysis [14], except on the Poisson and CIB amplitudes, which are set to be uniform in the range $[0, 200] \mu\text{K}^2$. We also updated the prior on the CIB clustering power law index to $\mathcal{N}(0.53, 0.1)$, to reflect the results of [76].

In Fig. 37, we present the best-fit total foreground contributions as determined from the data. At small angular scales, the foregrounds are dominated by Poisson sources, while clustered CIB becomes increasingly important at higher frequencies. The feature near $\ell \sim 500$ in the 95×220 GHz spectrum arises from the tSZ–CIB cross-correlation, which is negative; this effect is only mildly significant given the uncertainties. We note that the foreground model provides a good fit to the data. In particular, the successful passing of the TT power spectrum difference and conditional tests (see §VIB) demonstrates that the foreground model is sufficiently flexible to account for the observed differences in the band powers across frequencies.

Appendix F: Construction and performance of the CMB-only likelihood

We provide additional information on the construction and performance of the *lite* likelihood, SPT-*lite*. The framework for the construction of the CMB-only likelihood introduced by [31] and used here cannot translate the information of parameter boundaries for nuisance parameters to the covariance of the CMB-only band powers [112] (unless one resorts to MCMC sampling the

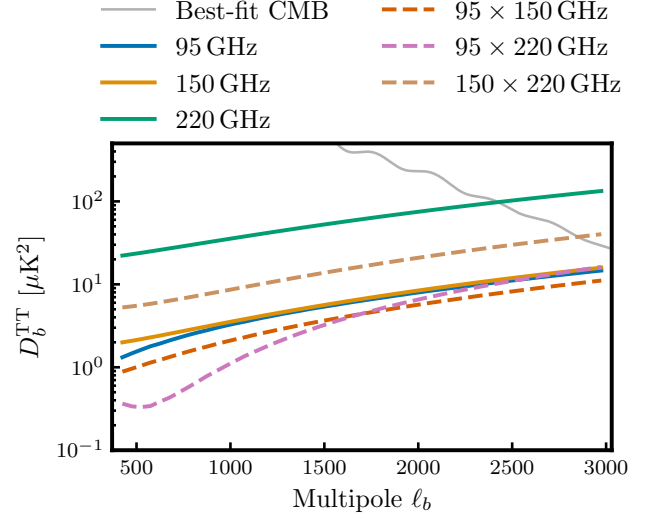


FIG. 37. Total foreground model contributions to the TT band powers at each frequency, assuming Λ CDM cosmology. We show the auto-frequency channels as solid lines and the cross-frequency channels as dashed lines. The total foreground model is the sum of the Poisson, CIB clustering, tSZ, kSZ, and Galactic dust contributions. The best-fit Λ CDM CMB prediction for the SPT-3G D1 T&E is shown as a gray line.

reconstruction likelihood [75], which we would like to avoid due to the numerical cost). However, the multi-frequency likelihood and the chosen nuisance parameter priors are fairly constraining so this is typically not an issue. The only exceptions are the TT Poisson power parameters and the beam sidelobe depolarization parameters, as the *lite* framework is by design only sensitive to frequency-differences. In response, we condition SPT-*lite* on the results of the multi-frequency likelihood in Λ CDM by (1) imposing regularization priors on the TT Poisson parameters centered on the best-fit values and 10 times wider than their corresponding constraints and (2) setting the beam sidelobe depolarization parameters to their best-fit values. This breaks the otherwise complete degeneracy of the Poisson parameters and ensures the bulk of the posterior mass is in the physical range for all parameters; this allows the reconstruction procedure to better capture the uncertainty due to the nuisance parameters in the covariance of the CMB-only band powers. Though this procedure is no longer strictly independent of cosmology, we prefer to condition the *lite* likelihood this way to improve its performance: indeed, as we show below this leads to a good match between the parameter constraints inferred from the multi-frequency likelihood and SPT-*lite* in Λ CDM and beyond. We have verified that changing the width of the Poisson priors or offsetting the β_{pol} parameters in a frequency-coherent way has a small impact on cosmological parameters.

We compare the one-dimensional marginalized posterior distributions assuming Λ CDM for H_0 , $\Omega_b h^2$, $\Omega_c h^2$,

Parameter	Description	Prior	Λ CDM Best-fit	Posterior
$A^{\text{tSZ}} [\mu\text{K}^2]$	Thermal SZ amplitude	$\mathcal{N}(3.23, 2.4) \ \& \ \mathcal{U}(0, \infty)$	1.03	$0.93^{+0.41}_{-0.60}$
$A^{\text{kSZ}} [\mu\text{K}^2]$	Kinetic SZ amplitude	$\mathcal{N}(3.7, 4.6) \ \& \ \mathcal{U}(0, \infty)$	0.09	< 2.90
$A_{150 \times 150}^{\text{CIB}} [\mu\text{K}^2]$	CIB clustering amplitude	$\mathcal{U}(0, 200)$	2.17	1.88 ± 0.80
$A_{150 \times 220}^{\text{CIB}} [\mu\text{K}^2]$		$\mathcal{U}(0, 200)$	7.24	7.5 ± 1.7
$A_{220 \times 220}^{\text{CIB}} [\mu\text{K}^2]$		$\mathcal{U}(0, 200)$	34.0	35 ± 5
α_{CIB}	CIB clustering power law index	$\mathcal{N}(0.53, 0.1)$	0.464	0.513 ± 0.092
$A_{95 \times 95}^{\text{Poisson}} [\mu\text{K}^2]$	Poisson amplitude	$\mathcal{U}(0, 200)$	12.44	$10.7^{+2.2}_{-1.7}$
$A_{95 \times 150}^{\text{Poisson}} [\mu\text{K}^2]$		$\mathcal{U}(0, 200)$	10.22	$8.6^{+2.1}_{-1.6}$
$A_{95 \times 220}^{\text{Poisson}} [\mu\text{K}^2]$		$\mathcal{U}(0, 200)$	18.47	$16.7^{+2.5}_{-2.3}$
$A_{150 \times 150}^{\text{Poisson}} [\mu\text{K}^2]$		$\mathcal{U}(0, 200)$	13.42	$11.9^{+2.1}_{-1.7}$
$A_{150 \times 220}^{\text{Poisson}} [\mu\text{K}^2]$		$\mathcal{U}(0, 200)$	34.19	32.1 ± 2.8
$A_{220 \times 220}^{\text{Poisson}} [\mu\text{K}^2]$		$\mathcal{U}(0, 200)$	98.9	$95.1^{+6.5}_{-5.5}$
κ	Super-sample lensing	$\mathcal{N}(0., 0.00045)$	$0.3 \cdot 10^{-5}$	0.00000 ± 0.00045
$A^{\text{Dust}} [\mu\text{K}^2]$	Galactic dust amplitude	$\mathcal{N}(1.88, 0.96)$	1.80	2.00 ± 0.82
α^{Dust}	Galactic dust power law index	$\mathcal{N}(-2.53, 0.05)$	-2.5312	-2.530 ± 0.049
β^{Dust}	Galactic dust spectral index	$\mathcal{N}(1.48, 0.02)$	1.4801	1.480 ± 0.020
$A_{\text{TE}}^{\text{PolGalDust}} [\mu\text{K}^2]$	TE polarized dust amplitude	$\mathcal{N}(0.12, 0.051)$	0.1040	0.104 ± 0.034
$\alpha_{\text{TE}}^{\text{PolGalDust}}$	TE polarized dust power law index	$\mathcal{N}(-2.42, 0.04)$	-2.4307	-2.434 ± 0.039
$\beta_{\text{TE}}^{\text{PolGalDust}}$	TE polarized dust spectral index	$\mathcal{N}(1.51, 0.04)$	1.5139	1.511 ± 0.040
$A_{\text{EE}}^{\text{PolGalDust}} [\mu\text{K}^2]$	EE polarized dust amplitude	$\mathcal{N}(0.05, 0.022)$	0.0566	0.057 ± 0.013
$\alpha_{\text{EE}}^{\text{PolGalDust}}$	EE polarized dust power law index	$\mathcal{N}(-2.42, 0.04)$	-2.4140	-2.417 ± 0.039
$\beta_{\text{EE}}^{\text{PolGalDust}}$	EE polarized dust spectral index	$\mathcal{N}(1.51, 0.04)$	1.5089	1.508 ± 0.040
Marginalized in baseline likelihood				
$A_{95 \times 95}^{\text{CIB-cl.}} [\mu\text{K}^2]$	CIB clustering amplitude	$\mathcal{N}(0.26, 0.15)$		
$A_{95 \times 150}^{\text{CIB-cl.}} [\mu\text{K}^2]$		$\mathcal{N}(0.04, 0.53)$		
$A_{95 \times 220}^{\text{CIB-cl.}} [\mu\text{K}^2]$		$\mathcal{U}(0, 200)$		
$\xi_{\text{tSZ-CIB}}$	tSZ-CIB correlation	$\mathcal{N}(0.18, 0.33) \ \& \ \mathcal{U}(-1, 1)$		

TABLE X. Overview of the foreground model parameters for the SPT T&E likelihood. The first part of the table lists the parameters that are varied in the baseline analysis, while the second part lists the parameters that are marginalized over by adding a suitable constant contribution to the band power covariance matrix. The best-fit values and 68% confidence limits are shown for the Λ CDM model. The priors are either Gaussian $\mathcal{N}(\mu, \sigma)$ or uniform $\mathcal{U}(a, b)$ distributions.

n_s , and $\log(10^{10} A_s)$ obtained from the multi-frequency likelihood and SPT-*lite* in Fig. 38. The means of the marginalized posterior distributions shift by $\lesssim 0.1\sigma$, where σ is the width of the *lite* posteriors. Error bars match to $\leq 6\%$. We calculate the size of the mean shift Δp in the full N -dimensional parameter space respecting the correlation of the parameters as given by the SPT-*lite* parameter covariance C via: $\sqrt{(\Delta p^T C^{-1} \Delta p)/N}$. This metric calculates the Euclidean distance in the parameter space transformed according to C and adjusts it for the dimensionality; in the one-parameter case, it reduces to the familiar $\Delta p/\sigma$. Calculating this metric yields an average offset of 0.09 per parameter, which is negligible; we conclude that the SPT-*lite* likelihood performs well in Λ CDM.

The good performance also holds up in extended model spaces. We further explore the comparison of SPT-*lite* and the multi-frequency likelihood by extending Λ CDM by A_{lens} , N_{eff} , Ω_k , and Σm_ν separately. For the latter two cases we also add DESI data, as primary CMB

data by themselves suffer from strong degeneracies in these models. We compare the same parameters as for Λ CDM plus the relevant extension parameters. This yields a total of 24 parameters to compare; the means differ by $< 0.3\sigma$ and error bars match to $< 10\%$ in all cases. Though we record slightly larger fluctuations than in Λ CDM, given the increased size of tests this is not surprising. As before, we calculate the size of the mean shift in the full parameter space respecting the correlation of parameters, finding all offsets to be < 0.2 . For Σm_ν , the 95% confidence limit inferred from the *lite* likelihood is 15% higher. In general, the *lite* likelihood leads to a minor widening of posteriors due to the treatment of the beams detailed above; this makes it a conservative choice. We conclude that the *lite* likelihood performs well for all model spaces considered and that the Λ CDM conditioning performed has a negligible impact on other cosmological models.

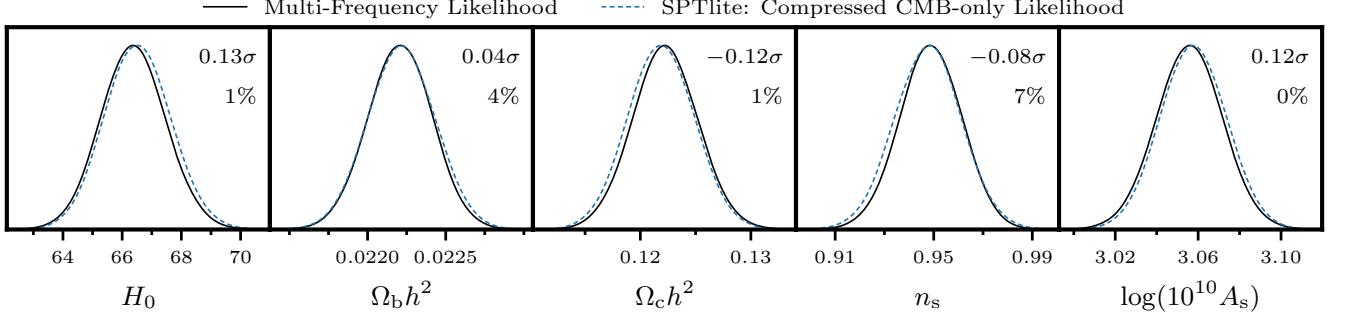


FIG. 38. Comparison of Λ CDM parameter constraints from the multi-frequency likelihood (black lines) and the compressed CMB-only SPT-lite likelihood (blue dashed lines). The means of the one-dimensional marginalized posterior distributions of the *lite* likelihood shift by $\lesssim 0.1\sigma$ compared to the multi-frequency likelihood, where σ is the width of the posterior when using SPT-lite. Similarly, the width of the posteriors match to $\leq 7\%$. Details on the construction of SPT-lite can be found in §VD.

Appendix G: Power spectrum consistency tests between frequencies

During the validation of the analysis pipeline detailed in Section VI, we perform a series of tests to check the consistency of the data across different frequency combinations. We only show an extract of those in Fig. 12. In this section, we show the complete results of these tests.

1. Probability to exceed

We outline the procedure for computing the PTE for each test presented in this work. First, we construct the vector Δ of observed values and the corresponding predicted covariance matrix Σ^Δ . The associated χ^2 statistic is then calculated as:

$$\chi^2 = \Delta^T \Sigma^\Delta \Delta. \quad (\text{G1})$$

We then compute the PTE as the probability of observing a χ'^2 value greater than or equal to the observed χ^2 value, given the degrees of freedom N_{dof} of the test. The PTE is defined as:

$$p = 1 - P(\chi'^2 \leq \chi^2, N_{\text{dof}}) \quad (\text{G2})$$

$$= 1 - \int_0^{\chi^2} d\chi'^2 P(\chi'^2, N_{\text{dof}}) \quad (\text{G3})$$

where $P(\chi'^2 \leq \chi^2, \text{dof})$ is the cumulative distribution function of the χ^2 probability function $P(\chi'^2, N_{\text{dof}})$ for N_{dof} degrees of freedom.

2. Difference tests

To validate the pipeline, we conduct a series of frequency-difference tests, comparing spectra derived from different frequency combinations. The objective is

to assess the consistency among the various frequency channels and to identify any potential issues in the data or analysis pipeline. The methodology is described in §VIB1 and a subset of the results is presented in Fig. 12. Here, we provide the full results of these tests. The figures in this section display the frequency-difference tests for TT, TE, and EE, along with the probability to exceed (PTE) for each spectral combination.

3. Conditional tests

In addition to the frequency-difference tests, we conduct conditional tests to evaluate the consistency of the data across different frequency combinations. Unlike the frequency-difference tests, which compare pairs of spectra, the conditional tests assess a given spectrum against its prediction based on all other spectra. This approach yields more stringent constraints on data consistency, as it leverages the information from the full set of remaining spectra to predict the spectrum in question. The methodology is outlined in §VIB2, and selected results are shown in Fig. 12. Here, we present the complete results of these tests. In this section, Figs. 39 to 41 display the conditional tests for TT, TE, and EE, along with the PTE for each spectral combination.

Appendix H: Comparison with data recorded in 2018

We compare these results against those derived from data recorded during the 2018 observing season reported in [14]. While these data are also subject to quadrupolar leakage and the depolarization of beam sidelobes, these effects went unmodeled as they were unknown at the time. Though this may bias the results of [14], the uncertainties of the 2018 data set were much larger; they therefore did not necessitate the sophisticated methods employed here to be modeled accurately and any relative

bias is expected to be smaller than what it would be for the new data. We stress that the 2018 data and results are superseded by this work.

In principle, a comparison at the band power level would necessitate a detailed understanding of the correlation between the two data sets, which is difficult to model for various reasons.⁴⁹ Instead, we perform a parameter-level comparison assuming the Λ CDM model. Since this model provides a good description of both data sets it allows for a qualitative check. We restrict the two data sets to the common multipole moment range and perform MCMC analyses jointly of the full set of T&E spectra, as well as of each spectrum type individually. The parameter set we choose to compare constraints across includes the parameters θ_s^* , $\Omega_b h^2$, $\Omega_c h^2$, n_s , and $10^9 A_s e^{-2\tau_{\text{reio}}}$. We quantify parameter differences in two ways, either by assuming the data sets are independent or by assuming the 2018 data is a subset of the new data, which mathematically corresponds to either adding or subtracting the parameter covariance matrices, respectively. Strictly speaking both of these tests are inappropriate; while in the first case, we ignore the shared sample variance fluctuations, in the second case we assume shared sample variance but also some common noise fluctuations. Still, by having an overly conservative and an overly optimistic test, we can gain a qualitative understanding of the consistency.

We first carry out the optimistic test, comparing constraints assuming the data sets are independent. We assume a total of three independent tests for each case and regard PTEs above $2.5\%/3 = 0.83\%$ to be passing. For the full set of T&E spectra we obtain a PTE of 76.96%, signaling good agreement. For TT, TE, and EE fits individually, we obtain PTEs of 85.98%, 71.54%, and 1.07%. All of these lie above the PTE threshold. Though differences in EE constraints may be related to the updated beam and leakage modeling, the associated PTE is statistically normal. Therefore, we conclude that the results presented here and the ones of [14] are broadly consistent.

The conservative test exhibits numerical difficulties as the resulting parameter difference covariance matrices are generally not positive definite. This is particularly the case for the combined T&E constraints, where changes in the degeneracy directions between parameters from [14] compared to this work lead to instabilities. Still, using the diagonal of the covariance calculated this way, all individual parameter shifts are $< 3\sigma$

for all cases. For T&E specifically, four out of the five parameters are offset by $< 1.5\sigma$. For individual spectrum fits, we are able to numerically stabilize the test by restricting ourselves to the two parameters that are the most discrepant in the one-dimensional marginalized posteriors. For TT-only fits, these are n_s and $10^9 A_s e^{-2\tau_{\text{reio}}}$, for which we obtain a PTE of 1.5%. Since the spectral tilt profits from the lower-noise measurement of the CMB damping tail that the new data set offers, it is not surprising to find a fluctuation in this parameter plane. We note that debiasing the new temperature band powers using the 2018 beam leads to a negligible shift in the TT-derived parameter constraints. Moreover, we have verified that the shift in the $n_s - 10^9 A_s e^{-2\tau_{\text{reio}}}$ plane produced by multiplying or dividing the TT band powers by up to eight powers of the pixel window function leads to a comparatively small shift that does not fully align with the direction of differences between the old and new data. Given that the other three parameters agree to $\leq 1.5\sigma$ we do not regard this PTE as problematic. For TE, we perform the test over (n_s, θ_s^*) and report a PTE of 4.83%, whereas for EE we restrict ourselves to the $(n_s, \Omega_c h^2)$ plane and report a PTE of 2.25%. We conclude that even under overly conservative assumptions, the results of this analysis agree with the predecessor work of [14].

⁴⁹ Different analysis choices were made regarding, for example, the filtering strategy, the source flux cut threshold, whether to analyze maps in a flat- or curved-sky framework, whether or not to inpaint sources, or how to bin the power spectrum measurement into band powers. Moreover, the focal plane was replaced after the 2018 observing season; this leads to small beam and bandpass changes. Together, these aspects lead to differences in the covariance matrices and band power window functions that are non-trivial to model.

Appendix I: Consistency of ACT DR6 and DESI data

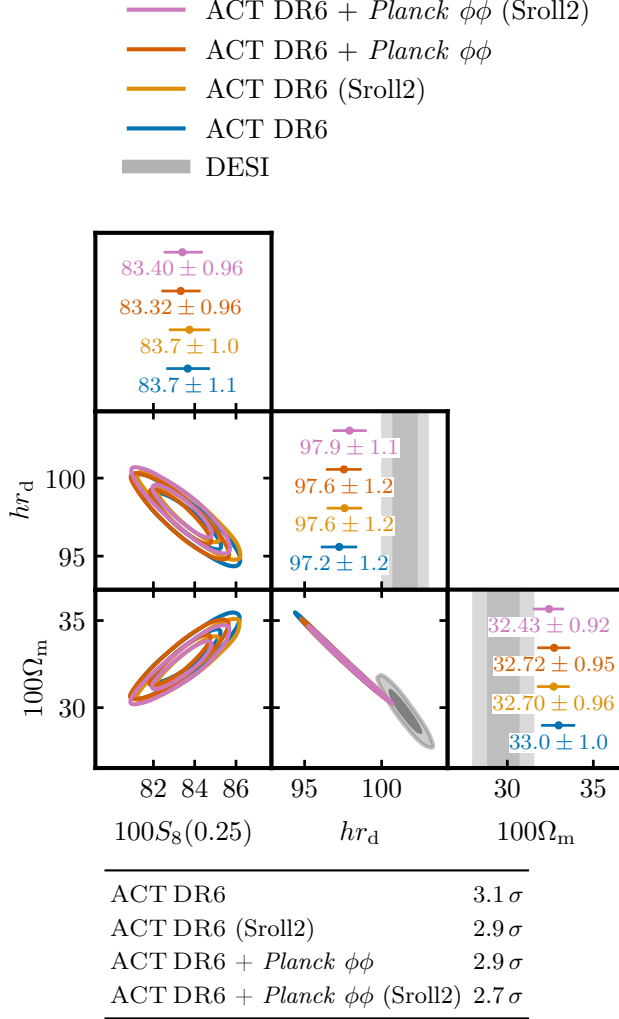


FIG. 43. Comparison of the Ω_m - hr_d parameters between the ACT DR6 and DESI data sets. We varied the τ prior, using Sroll2 likelihood [206] as used in [6]. The lensing reconstruction combined with the T&E data is either the $\phi\phi$ data from ACT DR6 alone [36] or incorporates the *Planck* lensing data [35]. Reported values in the table indicate the Gaussian distance between the data sets and the DESI constraints in the Ω_m - hr_d parameter plane.

The ACT DR6 Λ CDM cosmological parameters show a mild tension with the DESI data set, corresponding to a 3.1σ separation in the Ω_m - hr_d plane. When substituting our τ prior with the Sroll2 likelihood [206], which is the baseline choice in [6], we find a 2.9σ distance. We note that we are using the ACT-lite likelihood, whereas the baseline ACT DR6 results are based on multi-frequency likelihoods, and this may lead to a small difference in the reported distance. In addition, a recent study by the DESI collaboration [152] found that the

ACT DR6+*Planck* $\phi\phi$ Λ CDM cosmological parameters are 2.7σ distant from the DESI data set. The results reproduce those of [152]. In Fig. 25, we also examine how this result changes when replacing our τ prior with the Sroll2 likelihood. We conclude that the significance of the ACT DR6–DESI discrepancy is sensitive to analysis choices. Throughout this work, we reported values with consistent priors, based on the NPIPE τ prior, see Table III.

Appendix J: Glossary of cosmological parameters

We provide an overview of cosmological parameters used in this manuscript in Table XI.

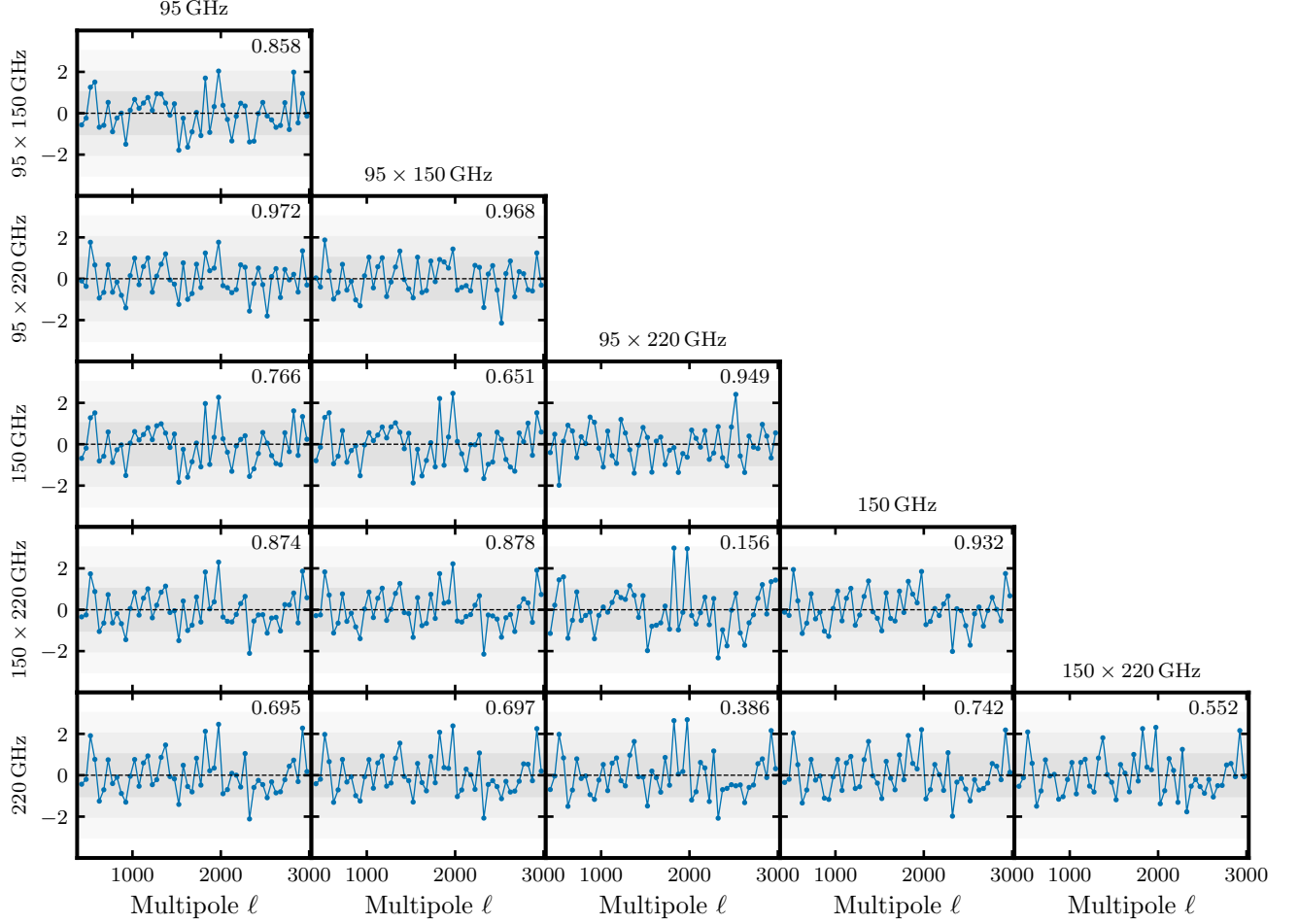


FIG. 39. Difference tests for TT spectra. Each block shows the difference between the two cross-frequency spectra $\mu \times \nu$ (row labels) and $\alpha \times \beta$ (column labels), e.g. the upper left panel is the difference spectrum $C_{\ell}^{95 \times 150} - C_{\ell}^{95 \times 95}$. We show quantities in units of the expected error bars, as calculated from the covariance matrix in Eqs. (46) and (47). The PTE values quoted on the upper right corners are calculated from Eq. (G2), with associated number of degrees of freedom per tests being $N_{\text{dof}} = 52$. Combining TT, TE, and EE, there are 18 independent difference tests, thus the PTE threshold is $0.05/18 = 0.0028$. The gray shaded regions indicate the 1, 2, and 3 σ regions of the expected distribution.

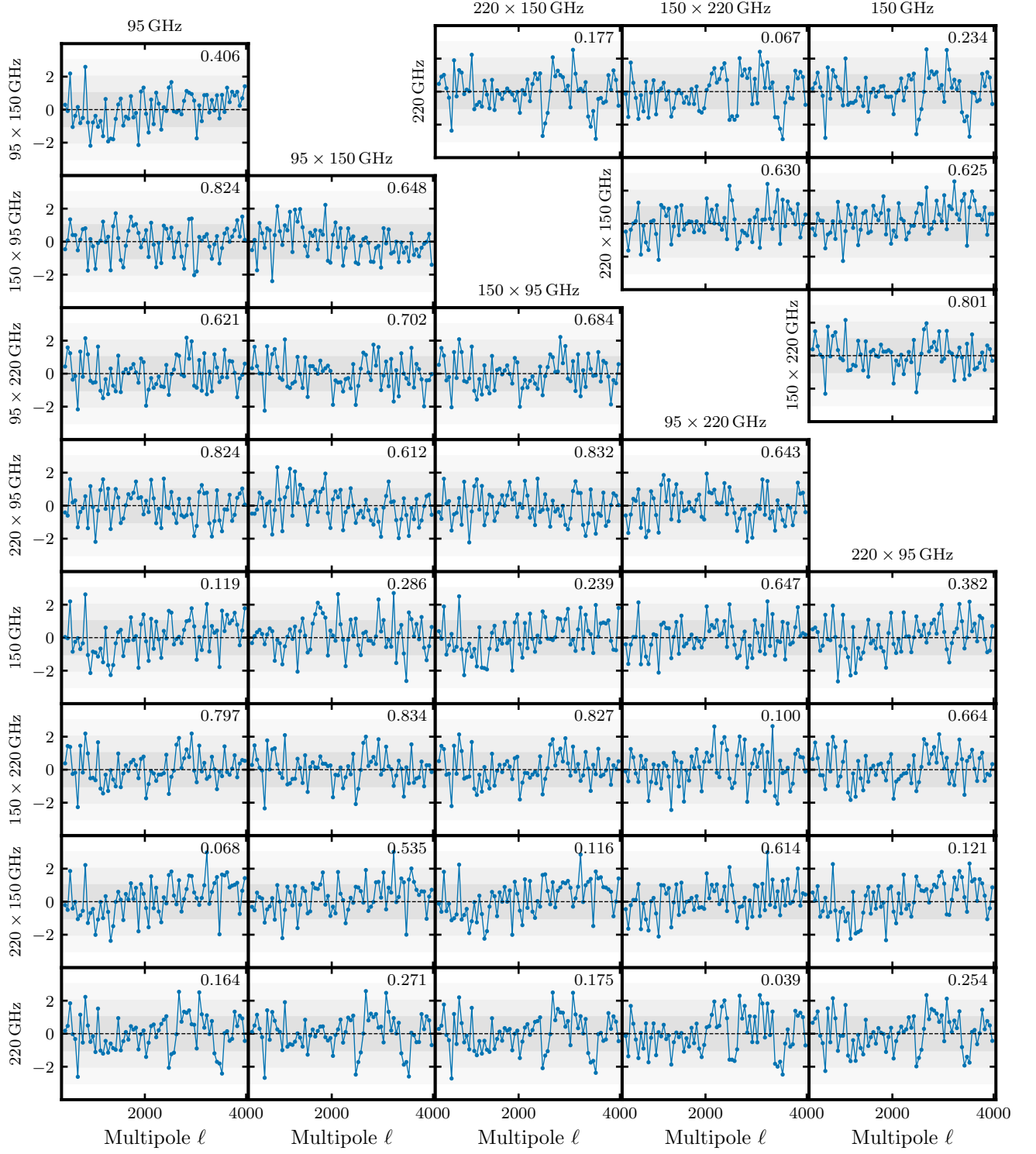


FIG. 40. Difference tests for TE spectra. Each block shows the difference between the two cross-frequency spectra $\mu \times \nu$ (row labels) and $\alpha \times \beta$ (column labels), e.g. the upper left panel is the difference spectrum $C_{\ell}^{95 \times 150} - C_{\ell}^{95 \times 95}$. We show quantities in units of the expected error bars, as calculated from the covariance matrix in Eqs. (46) and (47). The PTE values quoted on the upper right corners are calculated from Eq. (G2), with associated number of degrees of freedom per tests being $N_{\text{dof}} = 72$. Combining TT, TE, and EE, there are 18 independent difference tests, thus the PTE threshold is $0.05/18 = 0.0028$. The gray shaded regions indicate the 1, 2, and 3 σ regions of the expected distribution.

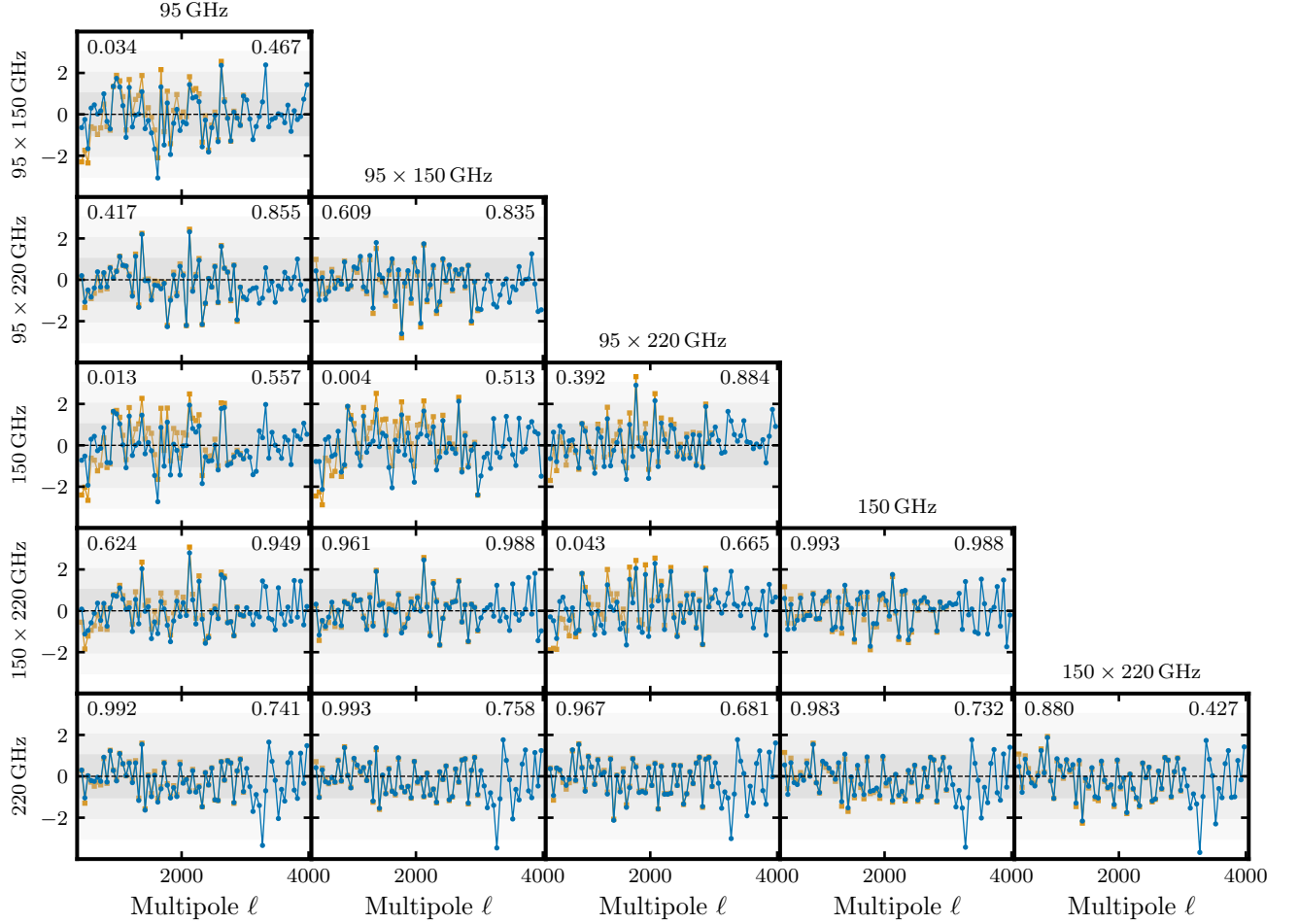


FIG. 41. Difference tests for EE spectra. Each block shows the difference between the two cross-frequency spectra $\mu \times \nu$ (row labels) and $\alpha \times \beta$ (column labels), e.g. the upper left panel is the difference spectrum $C_{\ell}^{95 \times 150} - C_{\ell}^{95 \times 95}$. We show quantities in units of the expected error bars, as calculated from the covariance matrix in Eqs. (46) and (47). The PTE values quoted on the upper right corners are calculated from Eq. (G2), with associated number of degrees of freedom per tests being $N_{\text{dof}} = 72$. The orange line and the PTEs quoted on the upper left are those obtained prior to unblinding, exhibiting a close-to-threshold value in 95-150 GHz, which is improved by the polarized beam modeling. Combining TT, TE, and EE, there are 18 independent difference tests, thus the PTE threshold is $0.05/18 = 0.0028$. The gray shaded regions indicate the 1, 2, and 3σ regions of the expected distribution.

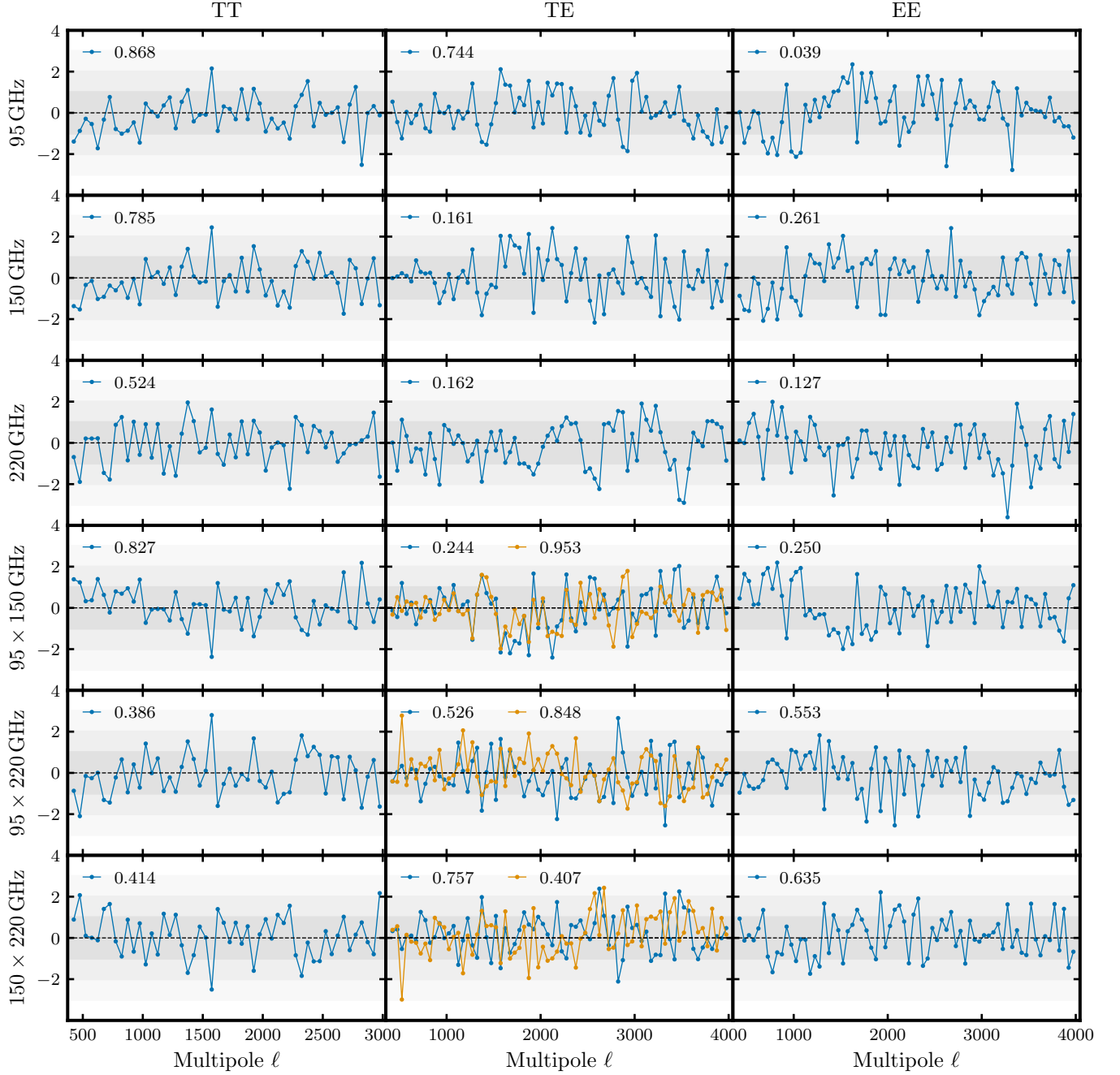


FIG. 42. Conditional tests. For each TT, TE, and EE cross frequency spectrum, we show the difference between the data and the prediction conditioned on the other cross-frequency spectra of the same channel. We show the difference in units of the expected error bars, as calculated from the covariance matrix in Eqs. (46) and (47). The PTE values quoted on the upper right corners are calculated from Eq. (G2), with associated number of degrees of freedom per tests being $N_{\text{dof}} = 52$ for TT and 72 for the others. There are 18 independent tests, thus the PTE threshold is $0.05/18 = 0.0028$. The gray shaded regions indicate the 1, 2, and 3 σ regions of the expected distribution. For the TE asymmetrical cross-frequencies case, we show on the same plot the $\mu \times \nu$ and $\nu \times \mu$ combinations in blue and orange, respectively.

Parameter	Definition
θ_s^*	Angle of the sound horizon at recombination
θ_{MC}	Approximate angle of the sound horizon at recombination
H_0	Hubble constant, expansion rate today in $\text{km s}^{-1} \text{Mpc}^{-1}$
h	$H_0/100 \text{ km s}^{-1} \text{Mpc}^{-1}$
$\Omega_b h^2$	Physical baryon density
$\Omega_c h^2$	Physical cold dark matter density
Ω_m	Fractional matter density
A_s	Amplitude of the power spectrum of initial scalar fluctuations
n_s	Tilt of the power spectrum of initial scalar fluctuations
τ_{reio}	Optical depth to reionization
$w_0 w_a$	Dark energy equation of state parameters (see Eq. (98))
w_\perp	Deviation from a cosmological constant along the BAO degeneracy direction, $w_a + 3.5(w_0 + 1)$
Ω_k	Mean spatial curvature
σ_8	Root mean square of matter fluctuations today in linear theory in a sphere of comoving radius of $8 h^{-1} \text{Mpc}$
N_{eff}	Effective number of neutrino species
Y_P	Primordial helium abundance (mass fraction)
Σm_ν	Sum of neutrino masses
r_d	Comoving size of the sound horizon at the end of the baryon drag epoch
$A_{2\text{pt}}$	Amplitude of the signature of gravitational lensing in the primary CMB power spectra
A_{recon}	Amplitude of the CMB gravitational lensing power spectrum
A_{lens}	Coherent variation of the amplitude of gravitational lensing, $A_{2\text{pt}} = A_{\text{recon}}$
m_e	Electron mass
$m_{e,0}$	Electron mass today
X_e	Ionization fraction of the universe at recombination

TABLE XI. Definition of all cosmological parameters used.

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